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Lecture 1

Fundamentals of Monte Carlo Particle Transport



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Fundamentals of Monte Carlo Particle Transport

Solving particle transport problems with the Monte Carlo method is simple just simulate the particle behavior. The devil is in the details, however. This course provides a balanced approach to the theory and practice of Monte Carlo simulation codes, with lectures on transport, random number generation, random sampling, computational geometry, collision physics, tallies, statistics, eigenvalue calculations, variance reduction, and parallel algorithms. This is not a course in how to use MCNP or any other code, but rather provides in-depth coverage of the fundamental methods used in all modern Monte Carlo particle transport codes. The course content is suitable for beginners and code users, and includes much advanced material of interest to code developers. (10 lectures, 2 hrs each)

The instructor is Forrest B. Brown from the X-5 Monte Carlo team. He has 25 years experience in developing production Monte Carlo codes at DOE laboratories and over 200 technical publications on Monte Carlo methods and high-performance computing. He is the author of the RACER code used by the DOE Naval Reactors labs for reactor design, developed a modern parallel version of VIM at ANL, and is a lead developer for MCNP5, MCNP6, and other Monte Carlo codes at LANL.

Topics



1. Introduction

- Monte Carlo & the Transport Equation
- Monte Carlo & Simulation
- 2. Random Number Generation
- 3. Random Sampling
- 4. Computational Geometry
- 5. Collision Physics
- 6. Tallies & Statistics
- 7. Eigenvalue Calculations Part I
- 8. Eigenvalue Calculations Part II
- 9. Variance Reduction
- **10. Parallel Monte Carlo**

11. References



- Von Neumann invented scientific computing in the 1940s
 - Stored programs, "software"
 - Algorithms & flowcharts
 - Assisted with hardware design as well
 - "Ordinary" computers today are called "Von Neumann machines"
- Von Neumann invented Monte Carlo methods for particle transport in the 1940s (with Ulam, Fermi, Metropolis, & others at LANL)
 - Highly accurate no essential approximations
 - Expensive typically the "method of last resort"
 - Monte Carlo codes for particle transport have been proven to work effectively on all types of computer architectures:

SIMD, MIMD, vector, parallel, supercomputers, workstations, PCs, Linux clusters, clusters of anything,...

Introduction



- Two basic ways to approach the use of Monte Carlo methods for solving the transport equation:
 - Mathematical technique for numerical integration
 - Computer simulation of a physical process

⇒ Each is "correct"

– Mathematical approach is useful for:

Importance sampling, convergence, variance reduction, random sampling techniques, eigenvalue calculation schemes,

- Simulation approach is useful for: collision physics, tracking, tallying,
- Monte Carlo methods solve integral problems, so consider the <u>integral</u> form of the Boltzmann equation
- Most theory on Monte Carlo deals with fixed-source problems. Eigenvalue problems are needed for criticality and reactor physics calculations.

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Simple Monte Carlo Example

Evaluate
$$G = \int_{0}^{1} g(x) dx$$
, with $g(x) = \sqrt{1 - x^2}$ g(x)

Mathematical approach:

For k = 1, ..., N: choose \hat{x}_k randomly in (0,1)

$$G = (1-0) \cdot [\text{average value of } g(x)] \approx \frac{1}{N} \cdot \sum_{k=1}^{N} g(\hat{x}_k) = \frac{1}{N} \cdot \sum_{k=1}^{N} \sqrt{1-x_k^2}$$

• Simulation approach:

"darts game" For k = 1, ..., N: choose \hat{x}_k , \hat{y}_k randomly in (0,1), if $\hat{x}_k^2 + \hat{y}_k^2 \le 1$, tally a "hit" G = [area under curve] $\approx (1 \cdot 1) \cdot \frac{\text{number of hits}}{N}$



▲



Monte Carlo is often the method-of-choice for applications with integration over many dimensions

Examples: high-energy physics, particle transport, financial analysis, risk analysis, process engineering,

Evaluate

$$G = \int_{a_1 a_2}^{b_1 b_2} \int_{a_1}^{b_M} g(r_1, r_2, ..., r_M) dr_1 dr_2 ... dr_M$$

where r1, r2, ..., rM are all independent variables

h

For k = 1, ..., N:

For m = 1, ..., M: choose $R_m^{(k)}$ randomly in (a_m, b_m)

$$G \sim (b_1 - a_1) \cdot ... \cdot (b_M - a_M) \cdot \frac{1}{N} \sum_{k=1}^{N} g(R_1^{(k)}, R_2^{(k)}, ..., R_M^{(k)})$$

Introduction – Probability Density Functions

Continuous Probability Density
 f(x) = probability density function (PDF)
 f(x) ≥ 0

Probability $\{a \le x \le b\} = \int_{a}^{b} f(x) dx$ Normalization: $\int_{a}^{\infty} f(x) dx = 1$



Discrete Probability Density

$$\{ f_k \}, k = 1,...,N, \text{ where } f_k = f(x_k)$$

$$f_k \ge 0$$

$$\text{Probability}\{ x = x_k \} = f_k$$

$$\text{Normalization: } \sum_{k=1}^N f_k = 1$$







Mean, Average, Expected Value

$$x = \mu = \langle x \rangle = E[x]$$

$$\mu = \int_{-\infty}^{+\infty} xf(x) dx \quad \text{[continuous]} \qquad \mu = \sum_{k=1}^{N} x_k f_k \quad \text{[discrete]}$$

Variance

$$var(x) = \overline{(x-\mu)^{2}} = \sigma^{2} = \langle (x-\mu)^{2} \rangle = E[(x-\mu)^{2}]$$

$$\sigma^{2} = \int_{-\infty}^{+\infty} (x-\mu)^{2} f(x) dx \qquad \sigma^{2} = \sum_{k=1}^{N} (x_{k}-\mu)^{2} f_{k}$$

Standard Deviation

$$\sigma = \sqrt{\sigma^2}$$

Functions of a Random Variable

Consider g(x), where x is a random variable with density f(x) $E[g(x)] = \int_{-\infty}^{+\infty} g(x)f(x) dx \qquad E[g(x)] = \sum_{k=1}^{N} g_k f_k$

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The key to Monte Carlo methods is the notion of *random sampling*.

• The problem can be stated this way:

Given a probability density, f(x), produce a sequence of X's. The \hat{X} 's should be distributed in the same manner as f(x).



- The use of random sampling distinguishes Monte Carlo from other methods
- When Monte Carlo is used to solve the integral Boltzmann transport equation:
 - Random sampling models the outcome of physical events (e.g., neutron collisions, fission process, sources,)
 - Computational geometry models the arrangement of materials



Monte Carlo & Transport Equation



Boltzmann transport equation — time-independent, linear $\Psi(\mathbf{r}, \mathbf{v}) = \int \left[\int \Psi(\mathbf{r}', \mathbf{v}') C(\mathbf{v}' \to \mathbf{v}, \mathbf{r}') d\mathbf{v}' + Q(\mathbf{r}', \mathbf{v}) \right] T(\mathbf{r}' \to \mathbf{r}, \mathbf{v}) d\mathbf{r}'$

where

- $\Psi(\mathbf{r}, \mathbf{v})$ = particle collision density
- $Q(\mathbf{r}', \mathbf{v}) = \text{source term}$
- $C(\mathbf{v}' \rightarrow \mathbf{v}, \mathbf{r}') =$ collision kernel, change **velocity** at fixed position
- $T(\mathbf{r}' \rightarrow \mathbf{r}, \mathbf{v})$ = transport kernel, change **position** at fixed velocity

• Angular Flux
$$\psi(\mathbf{r}, \mathbf{v}) = \frac{\Psi(\mathbf{r}, \mathbf{v})}{\Sigma(\mathbf{r}, |\mathbf{v}|)}$$

• Scalar Flux $\Phi(\mathbf{r}, |\mathbf{v}|) = \int_{\hat{\Omega}} \frac{\Psi(\mathbf{r}, \mathbf{v})}{\Sigma(\mathbf{r}, |\mathbf{v}|)} d\hat{\Omega}, \quad \mathbf{v} = |\mathbf{v}|\hat{\Omega}$



Source term for the Boltzmann equation:

$$Q(\mathbf{r}, \mathbf{v}) = \begin{pmatrix} S(\mathbf{r}, \mathbf{v}) & \leftarrow \text{Fixed Source} \\ S(\mathbf{r}, \mathbf{v}) + \int \Psi(\mathbf{r}, \mathbf{v}') F(\mathbf{v}' \to \mathbf{v}, \mathbf{r}) d\mathbf{v}' & \leftarrow \text{Fixed Source} + \text{Fission} \\ \frac{1}{K} \int \Psi(\mathbf{r}, \mathbf{v}') F(\mathbf{v}' \to \mathbf{v}, \mathbf{r}) d\mathbf{v}' & \leftarrow \text{Eigenvalue} \end{cases}$$

where

• S (r , v)	=	fixed source
• $F(\mathbf{v}' \rightarrow \mathbf{v}, \mathbf{r})$	_ =	creation operator (due to fission), particle at (r, v) creates particle at (r, v
• K	=	eigenvalue



$$\Psi(r,v) = \int \left[\int \Psi(r',v') \cdot C(v' \to v,r') dv' + Q(r',v) \right] \cdot T(r' \to r,v) dr'$$

Assumptions

- Static, homogeneous medium
- Time-independent
- Markovian next event depends only on current (r,v,E), not on previous events
- Particles do not interact with each other
- Neglect relativistic effects
- No long-range forces (particles fly in straight lines between events)
- Material properties are not affected by particle reactions
- Etc., etc.

\Rightarrow Can use the superposition principle



Basis for the Monte Carlo Solution Method

Let
$$p = (\vec{r}, \vec{v})$$
 and $R(p' \rightarrow p) = C(v' \rightarrow v, r') \cdot T(r' \rightarrow r, v)$

Expand Ψ into components having 0,1,2,...,k collisions $\Psi(p) = \sum_{k=0}^{\infty} \Psi_k(p), \quad \text{with} \quad \Psi_0(p) = \int Q(r',v)T(r' \to r,v)dr'$

By definition,

$$\Psi_{k}(p) = \int \Psi_{k-1}(p') \cdot \mathsf{R}(p' \to p) dp'$$

Note that collision k depends only on the results of collision k-1, and not on any prior collisions k-2, k-3, ...



Histories

• After repeated substitution for Ψ_{k}

$$\Psi_{k}(p) = \int \Psi_{k-1}(p') \cdot R(p' \to p) dp'$$

= $\int \dots \int \Psi_{0}(p_{0}) \cdot R(p_{0} \to p_{1}) \cdot R(p_{1} \to p_{2}) \dots R(p_{k-1} \to p) dp_{0} \dots dp_{k-1}$

• A "history" is a sequence of states (p₀, p₁, p₂, p₃,)



 For estimates in a given region, <u>tally</u> the occurrences for <u>each collision of each "history"</u> within a region



$$\Psi_{k}(p) = \int \dots \int \Psi_{0}(p_{0}) \cdot \mathsf{R}(p_{0} \rightarrow p_{1}) \cdot \mathsf{R}(p_{1} \rightarrow p_{2}) \dots \mathsf{R}(p_{k-1} \rightarrow p) dp_{0} \dots dp_{k-1}$$

Monte Carlo approach:

- Generate a sequence of states (p₀, p₁, p₂, p₃,) [i.e., a history] by:
 - Randomly sample from PDF for source: $\Psi_0(p_0)$
 - Randomly sample from PDF for k^{th} transition: $R(p_{k-1} \rightarrow p_k)$
- Generate estimates of results by averaging over states for M histories:

$$A = \int A(p) \cdot \Psi(p) dp \approx \frac{1}{M} \cdot \sum_{m=1}^{M} \left(\sum_{k=1}^{\infty} A(p_{k,m}) \right)$$



Monte Carlo & Simulation

"Simulation is better than reality" Richard W. Hamming, 1991



Simulation approach to particle transport:

Faithfully simulate the history of a single particle from birth to death.

- Random-walk for a single particle
 - Model collisions using physics equations & cross-section data
 - Model free-flight between collisions using computational geometry
 - Tally the occurrences of events in each region
 - Save any secondary particles, analyze them later





A "history" is the simulation of the original particle & all of its progeny



- Repeat for many histories, accumulating tallies
- Fundamental rule: Think like a particle !



Source

- Random sampling
 - E, Ω analytic, discrete, or piecewise-tabulated PDF's
- Computational geometry
 - m sample from region in 3-D space, or from discrete PDF

Tracking

- Random sampling
 - d_{collide} distance to collision, from mfp & exponential PDF
- Computational geometry

dgeom — distance-to-boundary, ray-tracing, next-region,

Collisions

Random sampling

E', Ω' - analytic, discrete, or piecewise-tabulated PDF's

Physics

 Σ , f(μ) — cross-section data, angular PDF's, kinematics,

Tallies

Statistics

Variance Reduction

Random sampling

Monte Carlo & Simulation



Single particle

- random-walk for particle history
- simulate events, from birth to death
- · tally events of interest



Batch of histories ("generation")

- random-walk for many particle histories
- tally the aggregate behavior

Overall

- timesteps
 - geometry changes
 - material changes
 - fuel depletion
 - burnable absorbers
 - control rods

*Lo	* Loop over timesteps						
ŀ	* Loop over batches						
	 * Loop over <i>histories</i> 						
I.	 + random walk 						
I۰							
ŀ	 						
I٠ -	•••						
•	→ compute statistics						
•	→ update number densities,						
	-						







Lecture 2

Random Number Generation

"Randomness is a negative property; it is the absence of any pattern." Richard W. Hamming, 1991

"...random numbers should not be generated by a method chosen at random." Donald Knuth, 1981

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Random Number Generators (RNGs)

- <u>Numbers</u> are not random; a <u>sequence of numbers</u> can be.
- Truly random sequences are generally <u>not</u> desired on a computer.
- Pseudo-random sequences:
 - Repeatable (deterministic)
 - Pass statistical tests for randomness
- RNG
 - Function which generates a sequence of numbers which appear to have been randomly sampled from a uniform distribution on (0,1)
 - Probability density function for f(x)



- Typical usage in codes: r = ranf()
- Also called "pseudo-random number generators"
- All other random sampling is performed using this basic RNG



Most production-level Monte Carlo codes for particle transport use <u>linear congruential random number</u> <u>generators</u>:

$$\mathbf{S}_{i+1} = \mathbf{S}_i \cdot \mathbf{g} + \mathbf{c} \mod 2^m$$

 $S_i = seed, g = multiplier, c = adder, 2^m = modulus$

- Robust, over 40 years of heavy-duty use
- Simple, fast
- Theory is well-understood (e.g., DE Knuth, Vol. 2, 177 pages)
- Not the "best" generators, but good enough RN's are used in unpredictable ways during particle simulation
- To achieve reproducibility of Monte Carlo calculations, despite vectorization or varying numbers of parallel processors, there must be a fast, direct method for skipping ahead (or back) in the random sequence



due to Lehmer, 1949
most common method, excellent (when not abused)
S₀

initial value

Method:

r_k

s_{k+1}

[g • s_k + c] mod p

where

 s_k , g, c, p = integers, r_k = real

s_k = seed

- g = generator, or multiplier
- c = increment
- p = modulus
- r_k = psuedo-random number, $0 \le r_k \le 1$
- "mod p" ⇒ "remainder after division by p", absolutely no roundoff is permitted
- Multiplicative: c = 0
- Mixed: c > 0

Simple RNG – Example #1



Example #1:	S _{k+1}	\leftarrow	[g·	s _k + c] mo	d p			
		with	g = 47,	c = 1	, s ₀ =	=1, p	= 100)	
s(0) =	1								
s(1) =	(47x1 + 1) mod	100	=	48	mod	100	=	48
s(2) = 1	(47x48 + 1)) mod	100	=	2257	mod	100		57
s(3) =	(47x57 + 1)) mod	100	=	2680	mod	100	=	80
s(4) =	(47x80 + 1)) mod	100	=	3761	mod	100	=	61
s(5) =	$(47 \times 61 + 1)$) mod	100	=	2868	mod	100	=	68
s(6) =	(47x68 + 1)) mod	100	=	3197	mod	100	=	97
s(7) =	(47x97 + 1)) mod	100	=	4560	mod	100	=	60
s(8) =	(47x60 + 1)) mod	100	=	2821	mod	100	=	21
s(9) =	(47x21 + 1)) mod	100	=	988	mod	100	=	88
s(10) =	(47x88 + 1)) mod	100	,= •	4137	mod	100	=	37
s(11) =	(47x37 + 1)) mod	100	=	1740	mod	100	=	40
s(12) =	(47x40 + 1)) mod	100	=	1881	mod	100	=	81
s(13) =	(47x81 + 1)) mod	100	=	3808	mod	100	= ,	8
s(14) =	(47x8 + 1)) mod	100	=	377	mod	100	=	77
s(15) =	(47x77 + 1)) mod	100	=	3620	mod	100	=	20
s(16) =	(47x20 + 1)) mod	100	= 1	941	mod	100		41
s(17) =	(47x41 + 1)) mod	100	=	1928	mod	100	=	28
s(18) =	(47x28 + 1)) mod	100	=	1317	mod	100	=	17
s(19) =	(47x17 + 1)) mod	100	=	800	mod	100	=	0
s(20) =	(47x0 + 1)) mod	100	=	1	mod	100	=	1
s(21) =	(47x1 + 1)	L) mod	100	=	48	mod	100	=	48
s(22) =	(47x48 + 1)	L) mod	100	=	2257	mod	100	=	57
etc.		-							



Example #2:

a(0) =

1

 $s_{k+1} \leftarrow [g \cdot s_k + c] \mod p$

with g = 5, c = 1, $s_0 = 1$, p = 100

5	0,	_	–										
s(1)	-	(5x1	+	1)	mod	100	=	6	mod	100	=	6
s(2)	=	(5x6	+	1)	mod	100	=	31	mod	100	=	31
s(3)	=	(5x31	+	1)	mod	100	=	156	mod	100	=	56
s(4 j	=	(5x56	+	1)	mod	100	=	281	mod	100	=	81
s(5 j	=	(5x81	+	1)	mod	100	=	406	mod	100	=	6
s(6)	=	(5x6	+	1)	mod	100		31	mod	100	=	31
et	.c.		-		-								

Example #3:

 $s_{k+1} \leftarrow [g \cdot s_k + c] \mod p$

with g = 5, c = 0, $s_0 = 1$, p = 100

Selecting Parameters for Linear Congruential RNGs



$\mathbf{s_{k+1}} \leftarrow \texttt{[g \cdot s_k + c] mod p}$

- Modulus (p):
 - choose $p = 2^N$
 - simplifies "mod p"- discard all but the N least significant bits
 - simplifies division by p— shift the "point" left by N bits
 - N should be as large as possible, N > 35 is best.
 - Usually, choose N to be number of bits in largest positive integer.
- Generator (g), Initial Seed (s₀), & Increment (c) :
 - choose g & c to maximize the period
 - large g is best to reduce serial correlation
 - obviously, g=1 or g=0 are bad

- For c = 0 (multiplicative PRNG):

choosing (1) $g \mod 8 = 3 \text{ or } 5$ (2) $s_0 = \text{ odd}$

results in: period = 2^{N-2} , the maximum possible period.

— For c > 0 (mixed PRNG):

choosing	(1)	c relatively prime to p
Ū	(2)	(g-1) to be a multiple of every prime factor of p
	Ì3Ì	(g-1) to be a multiple of 4 if p is a multiple of 4

results in:

period = 2^N, the maximum possible period.



Multiplicative congruential method — Lehmer

S _{i+1}	=	g•S _i +c	mod 2 ^m ,	$0 < S_i < 2^m$
ξį	=	S _i / 2 ^m ,		0<ξ<1

Typical parameters

		<u>2</u> m	<u>period</u>	g	C
RACER	(KAPL)	2 ⁴⁷	2 ⁴⁵	84,000,335,758,	957 0
RCP	(BAPL)	2 ⁴⁸	2 ⁴⁸	2 ⁹ +1 59	,482,192,516,946
MORSE	(ORNL)	2 ⁴⁷	2 ⁴⁵	5 ¹⁵	0
MCNP	(LANL)	2 ⁴⁸	2 ⁴⁶	5 ¹⁹	0
VIM	(ANL)	2 ⁴⁸	2 ⁴⁶	5 ¹⁹	0
RANF	(CRAY)	2 ⁴⁸	2 ⁴⁶	44,485,709,377,	909 0
— (G. I	Marsaglia)	2 ³²	2 ³²	69069	1
MCNP5	(LANL)	2 ⁶³	2 ⁶³	(varies)	1

Linear Congruential RNGs



Aside ...

For the multiplicative congruential method, why is the period limited to a maximum of 2^{N-2} ??

 $\mathbf{s}_{\mathbf{k}+1} \leftarrow \mathbf{g} \cdot \mathbf{s}_{\mathbf{k}} \mod \mathbf{p},$

s_0 odd, g mod 8 = 3 or 5

All sk's are odd, g is odd

 \Rightarrow g·s_k will always be odd,

reduces period by a factor of two.

```
    For g mod 8 = 3, trailing bits of g are (...011)
```

 $g \cdot s_k = (...011) \cdot (....11) = (...11)$

 $g \cdot s_k = (...011) \cdot (....01) = (...01)$

or

 \Rightarrow next-to-last bit of s_k will not change, reduces period by a factor of two.

 For g mod 8 = 5, trailing bits of g are (...101) $g \cdot s_k = (...101) \cdot (...1x1) = (...1x1)$ or $g \cdot s_k = (...101) \cdot (...0x1) = (...0x1)$

 \Rightarrow third-to-last bit of s_k will not change, reduces period by a factor of two.



Example -- CYBER-205 RANF

$s_{k+1} \leftarrow [g \cdot s_k + c] \mod p$							
FORTRAN		META					
common /q8ranfc/ seed	LOD	s_descr, s	*load the seed				
r = ranf()	EX	g, 84000335758957	*generator				
	EX	e, 65489	*exponent, 2**-47				
	MPYL	g, s, s	*mult, keep last 47 bits				
	STO	s_descr, s	*store new seed				
	PACK	e, s, r	*insert exponent				
	ADDN	r, , r	<pre>*normalized result</pre>				
 (1) 0							

Note:

(1) 0 < r < 1
(2) scalar timing ~320 ns / prn
(3) to vectorize — "unroll" or "replicate", vector timing ~30 ns / rn

How long will the PRNs last ?

time to generate ALL 245 RNs

	1 M yr
scalar	4 mós
vector	15 days
2-pipe vector	12 days
4-pipe vector	6 days
vector x 4	3 days
vector x 4	30 hr
vector x 8	13 hr
vector x 16	4 hr
	scalar vector 2-pipe vector 4-pipe vector vector x 4 vector x 4 vector x 8 vector x 16

2 - 10

MCNP5 Random Number Generator


MCNP5 Random Number Generator – Usage



Program mcnp5

```
! Initialize RN parameters for problem
call RN_init_problem( new_seed= ProblemSeed )
.....
```

```
do nps = 1, number_of_histories
```

.

```
! Analyze one particle history
call RN_init_particle( nps )
....
if( rang() > xs ) ...
! Terminate history
call RN_update_stats
```

Random Number Generators



Other PRNGs

- Middle-square method:
- · Quadratic-congruential:
- Modified Middle-square:
- Additive:
- · Additive (or Shift):
- Generalized Additive (or Shift):
- · Quasi-random sequences
- etc., etc.,

Testing PRNGs

• See Knuth, Vol. 2, pp. 38-113

Spectral test

Empirical Tests:

Chi-square test Serial pair, triplet,, distributions Coupon Collector test Collision test etc.	Gap test Run test Serial Correlation coefficients	Poker tests Maximum-of-t test
--	---	----------------------------------

 $s_{k+1} = middle digits of s_k^2$

 $s_{k+1} = [s_k(s_k+1)] \mod p$

 $s_{k+1} = [s_k + s_{k-i}] \mod p$

 $s_k = [s_{k-j} + s_{k-i}] \mod p$

 $s_{k+1} = [a \cdot s_k^2 + b s_k + c] \mod p$

 $s_k = [a_1s_{k-1} + a_2s_{k-2} \dots a_is_{k-i}] \mod p$

Theoretical Tests

Serial Correlation (g	lobal)
-----------------------	--------

etc.



Reproducibility of a Particle History

- use separate, distinct random sequence for each particle
- starting seeds for separate particles are separated by "stride"



- stride should be large enough to prevent overlap (for most histories)
 - 1000 is common for reactor analysis problems
 - splitting & variance reduction not needed for in-core physics
 - réduces total random number usage
 - 4,297 is the "old" default for MCNP & VIM
 - 152,917 is the default for MCNP & VIM
 - prepared for lots of splitting & variance reduction
 - potential for lots of secondary particles

Random Number Generators – Reproducibility



particle stride

Parallel processing

 take "super-stride" in random sequence for particles on each processor



Eigenvalue Problems





To skip ahead k steps in the random sequence, [initial seed] \rightarrow [kth seed] $S_k = g \cdot S_{k-1} + c \mod 2^m$ $= g \cdot (g \cdot S_{k-2} + c) + c \mod 2^m$ $= g(\dots, g(g(gS_0 + c) + c) + c) \dots) + c \mod 2^m$ $= g^k \cdot S_0 + c \cdot (g^{k-1} + g^{k-2} + \dots + g + 1) \mod 2^m$ $= g^k \cdot S_0 + c \cdot (g^{k-1})/(g-1) \mod 2^m$

- Periodic sequence: negative skip k_n equivalent to positive skip (period - k_n)
- Can skip from any seed directly to any other:
 initial seed → ith seed for jth particle on mth processor in nth batch particle i → particle j
 batch i → batch j
- All arithmetic must be performed mod 2^m, without truncation or roundoff

$$S_{k} = G(k) \cdot S_{0} + C(k) \mod 2^{m}$$



Define $G(k) = g^k \mod 2^m$

m = 32 or 48 (typical), based on the size of a computer word $-2^m < k < +2^m$, based on desired "stride"

Denote the jth bit of k by k_[i], so that

$$k = 2^{m-1} k_{[m-1]} + 2^{m-2} k_{[m-2]} + \dots + 2^{1} k_{[1]} + 2^{0} k_{[0]}$$

Substituting into G(k) yields

$$G(k) = g^{k} \mod 2^{m} = g^{j=0} \mod 2^{m}$$
$$= \prod_{j=0}^{m-1} (g^{2^{j}})^{k} (j) \mod 2^{m}$$

Efficient algorithms for evaluating G(k) can be formulated using only m steps

• Los Alamos

Enumerating a few terms of G(k) makes the algorithm obvious

$$G(k) = \left(g^{1}\right)^{k} [0] \bullet \left(g^{2}\right)^{k} [1] \bullet \left(g^{4}\right)^{k} [2] \bullet \left(g^{8}\right)^{k} [3] \cdots \left(g^{2^{m-1}}\right)^{k} [m-1] \mod 2^{m}$$

Note that $k_{[j]}=0$ or $k_{[j]}=1$, so that each term $(g^n)^{k}{}^{[j]}$ evaluates to either 1 or g^n

Algorithm G:		
G ← 1	1, h ← g,	$i \leftarrow k + 2^m \mod 2^m$
while	i > 0	
	if i = odd:	$G \leftarrow Gh \mod 2^m$
	$h \leftarrow h^2 \mod 2$	m
	i ← [i/2]	

Remarks

- Algorithm G terminates after m steps, rather than k steps
- Negative strides are trivial, due to periodicity: G(-s) = G(2^m-s)

Random Number Generators – Skip Ahead

Define

$$C(k) = c\left(\frac{g^{k}-1}{g-1}\right) \mod 2^{m}$$
$$= c \cdot \left(1+g+g^{2}+g^{3}+\ldots+g^{k-1}\right) \mod 2^{m}$$

Los Alamos

The series for C(k) can be evaluated recursively, similar to G(k), in m steps:

- Since most of the common random number generators use c = 0, Algorithm C is generally not required.
- Algorithm C can be included with Algorithm A, at very little extra cost

RNG & Skip Ahead – Example



R. N. Generator for 32-bit ma	chines	(sparc2,	rs6000, indigo,)
s ← 69069・s + 1	mod 2 ³²		
	static unsig static doub	ned long le	seed_c=1; norm=(1./4294967296.);
Random Number Generator→	double cra unsig seed return }	nf_(void) { ned long _ c = g*se n ((double	g=69069, c=1; ed_c + c;) seed_c * norm);
Routine for Arbitrary Skips ->	void cran	fjump_(unsigned long *seed, double *jump, unsigned long *newseed) {
Compute: gen = g ^k inc = c(g ^k -1)/(g-1)	unsig if(*ju else for(} * nev	<pre>ined long imp < 0) ; j; j>>=1) if(j&1) { c *= g+1; g *= g; /seed = ge</pre>	<pre>j, gen=1, inc=0, g=69069, c=1; j = *jump + 4294967296.; j = *jump;) { inc = inc*g + c; gen = gen*g; en * (*seed) + inc;</pre>
	}		

RNG &	Skip Ahea	d – Exa	ample	e	• Los Alamos
Fortran, 48-bit genera	ator: g=5	¹⁹ , (c=0,	m=48	(VIM & MCNP)
C, 32-bit generator:	g=69	9069, (c=1,	m=32	(from Marsaglia)
C. 32-bit				<u>Sparc2</u>	<u>rs6000/350</u>
random number				1.0 μs	.7 μs
skip forward, skip backward,	average for + average for -	⊦110 ⁵ ∙110 ⁵		7.4 μs 4.0 μs	10 μs 20 μs
Fortran, 48-bit					
random number				3.6 µs	2.3 μs
skip forward, skip backward,	+152,917 -152,917			163 μs 458 μs	78 μs 215 μs
skip forward, skip backward,	average for + average for -	+110 ⁵ ·110 ⁵		160 μs 695 μs	75 μs 232 μs
skip forward, skip forward,	+1,152,917 +1,152,917,	brute for	ce	189 μs 4.1 sec	90 μs 2.6 sec
skip backward, skip backward,	-1,152,917 -1,152,917, k	brute for	e	456 μs 8 yea	210 μs r 5 year



- Algorithms for direct skip-ahead in the random sequence are simple, fast, convenient,, for modern Monte Carlo codes
- Arbitrary positive or negative strides can be taken, without precomputing or hardwiring specific constants
- Direct skip-ahead simplifies the initialization of random numbers for each particle, especially for parallel processing
- Algorithms described are currently used in:

MCNP5	LANL	—- all machines
KENO-Va	— CSN (Spain)	— Convex-C3440
RACER	— KAPL	- Cray, Meiko CS1 & CS2, Sun, SGI,
parallel VIM	— ANL	— Sun, rs6000, SP1,







MCNP5

Random Number

Generation & Testing

- Knuth statistical tests
- Marsaglia's DIEHARD test suite
- Spectral test
- Performance test
- Results

F.B. Brown & Y. Nagaya, "The MCNP5 Random Number Generator", *Trans. Am. Nucl. Soc.* [also, LA-UR-02–3782] (November, 2002). 2-25

MCNP5 RNG: History



MCNP & related precursor codes

- 40+ years of intense use
- Many different computers & compilers
- Modern versions are parallel: MPI + threads
- History based: Consecutive RNs used for primary particle, then for each of it's secondaries in turn, etc.
- RN generator is small fraction of total computing time (~ 5%)

Traditional MCNP RN Algorithm

- Linear congruential, multiplicative

 $S_{n+1} = g S_n \mod 2^{48}, \quad g = 5^{19}$

- 48-bit integer arithmetic, carried out in 24-bit pieces
- Stride for new histories: 152,917
- Skip-ahead: crude, brute-force
- Period / stride = 460×10^6 histories
- Similar RN generators in RACER, RCP, MORSE, KENO, VIM



Algorithm

- Robust, well-proven
- Long period: > 10^9 particles x stride 152,917 = 10^{14} RNs
- >10⁹ parallel streams
- High-precision is **not** needed, low-order bits not important
- Must have fast skip-ahead procedure
- Reasonable theoretical basis, no correlation within or between histories

Coding

- Robust !!!! Must never fail.
- Rapid initialization for each history
- Minimal amount of state information
- Fast, but portable must be exactly reproducible on any computer/compiler



Linear congruential generator (LCG) •

 $S_{n+1} = g S_n + c \mod 2^m$,

Period = 2^{m} (for c>0) or 2^{m-2} (for c=0)

Traditional MCNP: m=63, c=1 MCNP5:

m=48, c=0 Period=10¹⁴, 48-bit integers Period=10¹⁹, 63-bit integers

How to pick g and c ???

RN Sequence & Particle Histories •

••••	•••••	•••••	
1	2	3	etc.

– Stride for new history: 152,917



• RN Generation in MCNP-5

- RN module, entirely replaces all previous coding for RN generation
- Fortran-90, using INTEGER(I8) internally, where I8=selected_int_kind(18)
- All parameters, variables, & RN generator state are PRIVATE, accessible only via "accessor" routines
- Includes "new" skip-ahead algorithm for fast initialization of histories, greatly simplifies RN generation for parallel calculations
- Portable, standard, thread-safe
- Built-in unit test, compile check, and run-time test
- Developed on PC, tested on SGI, IBM, Sun, Compaq, Mac, alpha

Extended generators : 63-bit LCGs



• Selection of multiplier, increment and modulus



• Mixed LCG(g, c, 2^{β}) g = 1 mod 4, c = odd Period : 2^{β}

• MCNP5 – Extension of multiplier

- $-5^{19} = 45$ -bit integer in the binary representation
- 5¹⁹ seems to be slightly small in 63-bit environment.
- Odd powers of 5 satisfy both conditions above.
- Try these: $(5^{19},0,2^{63}),$ $(5^{23},0,2^{63}),$ $(5^{25},0,2^{63}),$ $(5^{19},1,2^{63}),$ $(5^{23},1,2^{63}),$ $(5^{25},1,2^{63})$



- L'Ecuyer suggested 63-bit LCGs with good lattice structures. Math. Comp., 68, 249–260 (1999)
 - Good multipliers were chosen based on the **spectral test**.
 - Multiplicative LCGs
 - LCG(3512401965023503517, 0, 263)
 - LCG(2444805353187672469, 0, 263)
 - LCG(1987591058829310733, 0, 263)

Mixed LCGs

- LCG(9219741426499971445, 1, 2⁶³)
- LCG(2806196910506780709, 1, 2⁶³)
- LCG(3249286849523012805, 1, 2⁶³)

Tests for RNGs



- 13 different LCGs were tested:
 - Traditional MCNP RNG, $(5^{19}, 0, 2^{48})$
 - 6 Extended 63-bit LCGs
 - 6 L'Ecuyer's 63-bit LCGs
- Theoretical tests :
 - Analyze the RNG algorithm of based on number theory and the theory of statistics.
 - Theoretical tests depend on the type of RNG. (LCG, Shift register, Lagged Fibonacci, etc.)
 - For LCGs, the Spectral test is used

• Empirical tests :

- Analyze the uniformity, patterns, etc. of RNs generated by RNGs.
- Standard tests reviewed by D. Knuth, SPRNG test routines
- **DIEHARD tests** Bit level tests by G. Marsaglia, more stringent
- Physical tests RNGs are used in a practical application. The exact solutions for the tests are known. (not performed in this work)



- LCGs have regular patterns (lattice structures) when overlapping *t*tuples of a random number sequence are plotted in a hypercube. (Marsaglia, 1968).
- all the *t*-tuples are covered with families of parallel (*t-1*)-dimensional hyperplanes.
- The spectral test determines the maximum distance between adjacent parallel hyperplanes.

Illustration of the spectral test



2 -34



• μ value proposed by Knuth

- Represent the effectiveness of a multiplier.

Knuth's criterion

$\mu_t(m,g)$ for $2 \le t \le 6$	Result
$\mu_{t}(m,g) > 1$	Pass with flying colors
$0.1 \le \mu_t(m,g) \le 1$	Pass
μ _t (m,g) ≤ 0.1	Fail

• S value

- Normalized maximum distance $S_{t} = \frac{d_{t}^{*}(m)}{d_{t}(m,g)} \qquad \begin{array}{c} \text{Maximum distance between adjacent parallel} \\ \text{hyperplanes.} \\ d_{t}^{*}(m) : \text{ Lower bound on } d_{t}(m,g). \end{array}$
- The closer to 1 the S value is, the better the RNG is.

Spectral test for extended LCGs



Dimension(t)	2	3	4	5	6	7	8
LCG(5 ¹⁹ ,0,2 ⁶³)							
$\mu_{t}(m,g)$	1.7321	2.1068	2.7781	1.4379	0.0825	2.0043	5.9276
$S_t(m,g)$	0.6910	0.7085	0.7284	0.6266	0.3888	0.6573	0.7414
LCG(5 ²³ ,0,2 ⁶³)							
$\mu_{t}(m,g)$	0.0028	1.9145	2.4655	5.4858	0.3327	0.2895	6.6286
$S_t(m,g)$	0.0280	0.6863	0.7070	0.8190	0.4906	0.4986	0.7518
LCG(5 ²⁵ ,0,2 ⁶³)							
$\mu_{t}(m,g)$	0.3206	1.8083	0.0450	3.0128	0.3270	3.1053	0.4400
$S_t(m,g)$	0.2973	0.6733	0.2598	0.7265	0.4892	0.6998	0.5356
LCG(5 ¹⁹ ,1,2 ⁶³)							
$\mu_{t}(m,g)$	1.7321	2.9253	2.4193	0.3595	0.0206	0.5011	1.6439
$S_t(m,g)$	0.6910	0.7904	0.7036	0.4749	0.3086	0.5392	0.6316
LCG(5 ²³ ,1,2 ⁶³)							
$\mu_{t}(m,g)$	0.0007	2.8511	2.5256	3.1271	4.5931	1.8131	4.2919
$S_t(m,g)$	0.0140	0.7837	0.7112	0.7319	0.7598	0.6480	0.7121
LCG(5 ²⁵ ,1,2 ⁶³)							
$\mu_{t}(m,g)$	0.0801	3.4624	1.3077	1.0853	1.4452	0.7763	1.3524
$S_t(m,g)$	0.1486	0.8361	0.6033	0.5923	0.6266	0.5740	0.6163

Spectral test for L'Ecuyer's 63-bit LCGs



Dimension(t)	2	3	4	5	6	7	8
LCG(351240196	LCG(3512401965023503517,0,2 ⁶³)						
$\mu_{t}(m,g)$	2.9062	2.9016	3.1105	4.0325	5.3992	6.7498	7.2874
$S_t(m,g)$	0.8951	0.7883	0.7493	0.7701	0.7806	0.7818	0.7608
LCG(244480535	531876724	69,0,2 ⁶³)		-			
$\mu_{t}(m,g)$	2.2588	2.4430	6.4021	2.9364	3.0414	5.4274	4.6180
$S_t(m,g)$	0.7891	0.7443	0.8974	0.7228	0.7094	0.7579	0.7186
LCG(198759105	588293107	33,0,2 ⁶³)					
$\mu_{t}(m,g)$	2.4898	3.4724	1.7071	2.5687	2.1243	2.0222	4.1014
$S_t(m,g)$	0.8285	0.8369	0.6449	0.7037	0.6682	0.6582	0.7080
LCG(921974142	264999714	45,1,2 ⁶³)					
$\mu_t(m,g)$	2.8509	2.8046	3.5726	3.8380	3.8295	6.4241	6.8114
$S_t(m,g)$	0.8865	0.7794	0.7757	0.7625	0.7371	0.7763	0.7544
LCG(280619691	05067807	09,1,2 ⁶³)					
$\mu_{t}(m,g)$	1.9599	4.0204	4.4591	3.1152	3.0728	3.0111	3.7947
$S_t(m,g)$	0.7350	0.8788	0.8199	0.7314	0.7106	0.6967	0.7012
LCG(3249286849523012805,1,2 ⁶³)							
$\mu_{t}(m,g)$	2.4594	2.4281	3.7081	2.8333	3.7633	3.0844	1.9471
$S_t(m,g)$	0.8234	0.7428	0.7829	0.7176	0.7350	0.6991	0.6451



Results for the traditional MCNP RNG

Dimension(t)	2	3	4	5	6	7	8
$\mu_t(m,g)$	3.0233	0.1970	1.8870	0.9483	1.8597	0.8802	1.2931
$S_t(m,g)$	0.9129	0.3216	0.6613	0.5765	0.6535	0.5844	0.6129

- All extended 63-bit LCGs fail with Knuth's criterion.
- All L'Ecuyer's 63-bit LCGs pass with flying colors.
- Comparison of minimum S values

RNG	Minimum $S_t(m,g)$
LCG(5 ¹⁹ ,0,2 ⁴⁸)	0.3216
LCG(3512401965023503517,0,2 ⁶³)	0.7493
LCG(2444805353187672469,0,2 ⁶³)	0.7094
LCG(1987591058829310733,0,2 ⁶³)	0.6449
LCG(9219741426499971445,1,2 ⁶³)	0.7371
LCG(2806196910506780709,1,2 ⁶³)	0.6967
LCG(3249286849523012805,1,2 ⁶³)	0.6451



• SPRNG (Scalable Parallel Random Number Generators)

- Test programs are available. http://sprng.cs.fsu.edu

Standard test suite (Knuth)

- Equidistribution
- Serial
- Gap
- Poker
- Coupon collector's
- Permutation
- Runs-up
- Maximum-of-t
- Collision tests

Choice of test parameters

- L'Ecuyer's test suite : Comm. ACM 31 p.742 (1988)
- Vattulainen's test suite : Comp. Phys. Comm. 86 p.209 (1995)
- Mascagni's test suite : Submitted to Parallel Computing



- Check whether RNs are uniformly generated in [0, 1).
- Generate random integers in [0,d-1].
- Each integer must have the equal probability 1/d.



Criterion of "Pass or Failure"



- All empirical tests score a statistic.
- A goodness-of-fit test is performed on the test statistic and yield a p-value. (Chi-sqaure or Kolmogorov-Smirnov test)
- If the p-value is close to 0 or 1, a RNG is suspected to fail.
- Significance level : 0.01(1%)
- Repeat each test 3 times.
- All 3 p-values are suspicious, then the RNG fails.



DIEHARD test suite



- A battery of tests proposed by G. Marsaglia.
- Test all bits of random integers, not only the most significant bits.

Los Alamos

- More stringent than standard Knuth tests.
- Default test parameters were used in this work.
- Test programs are available. http://stat.fsu.edu/~geo/diehard.html

Included tests:

- Birthday spacings
- Overlapping 5-permutation
- Binary rank
- Bitstream
- Overlapping-pairs-sparse-occupancy (OPSO)
- Overlapping-quadruples-sparse-occupancy (OQSO)
- DNA
- Count-the-1's test on a stream of bytes
- Count-the-1's test for specific bytes
- Parking lot
- Minimum distance
- 3-D spheres
- Squeeze
- Overlapping sums
- Runs
- Craps

Overlapping-pairs-sparse-occupancy test (1)



- OPSO = Overlapping-Pairs-Sparse-Occupancy test
- Preparation of 32-bit integers

0.10574, 0.66509, 0.46622, 0.93925, 0.26551, 0.11361, ...

 $\left\lfloor 2^{32} * \xi_i \right\rfloor$

 $454158374,\ 2856527213,\ 2002411287,\ 4034027575,\ldots$

Binary representation

11011000100011110100000100110, 101010010000110010010101101,...

38, 365, 791, 55, ... Cumulat

Overlapping-pairs-sparse-occupancy test (2)

2-letter word 2-letter word

•

- Count the number of <u>missing</u> words (=j).
- The number of missing words should be very closely normally distributed with mean 141,909, standard deviation 290.

2-letter words are formed from an alphabet of 1024 letters. 0000100110, 0101101101, 1100010111, 0000110111, ... Decimal representation





Overlapping-quadruples-sparse-occupancy test



- OQSO = Overlapping-Quadraples-Sparse-Occupancy test
- Similar to the OPSO test.

Letter :
$$2^5 = 32$$
 letters

• 4-letter words are formed from an alphabet of 32 letters. 00110, 01101, 10111, 10111, ...



• The number of missing words should be very closely normally distributed with mean 141909, standard deviation 295.

DNA test



- · Similar to the OPSO and OQSO tests.

Letter : $2^2 = 4$ letters

10-letter words are formed from an alphabet of 4 letters.
10, 1, 11, 11, 11, 10, 0, 11, 10, ...

10-letter word

• The number of missing words should be very closely normally distributed with mean 141909, standard deviation 399.



- Criterion for DIEHARD test
 - If the p-value is close to 0 or 1, a RNG is suspected to fail.
 - Significance level : 0.01(1%)
 - A RNG fails the test if we get six or more p-values less than 0.01 or more than 0.99.
- Results for standard & DIEHARD tests
 - All 13 RNGs pass all standard tests with L'Ecuyer's, Vattulainen's and Mascagni's test parameters.
 - Extended and L'Ecuyer's 63-bit LCGs pass all the DIEHARD tests.
 - The traditional MCNP RNG fails the OPSO, OQSO and DNA tests in the DIEHARD test suite.
Result of OPSO test for traditional MCNP RNG



Tested bits	p-value	Tested bits	p-value
bits 23 to 32	0.0000	bits 11 to 20	0.7457
bits 22 to 31	0.0000	bits 10 to 19	0.0598
bits 21 to 30	0.0000	bits 9 to 18	0.1122
bits 20 to 29	0.0000	bits 8 to 17	0.4597
bits 19 to 28	0.0001	bits 7 to 16	0.0011
bits 18 to 27	0.6639	bits 6 to 15	0.6319
bits 17 to 26	0.0445	bits 5 to 14	0.7490
bits 16 to 25	0.0125	bits 4 to 13	0.2914
bits 15 to 24	0.7683	bits 3 to 12	0.1792
bits 14 to 23	0.9712	bits 2 to 11	0.3253
bits 13 to 22	0.1077	bits 1 to 10	0.7277
bits 12 to 21	0.0717		

Result of OQSO test for traditional MCNP RNG



Tested bits	p-value	Tested bits	p-value
bits 28 to 32	1.0000	bits 14 to 18	0.6487
bits 27 to 31	1.0000	bits 13 to 17	0.5575
bits 26 to 30	1.0000	bits 12 to 16	0.1634
bits 25 to 29	1.0000	bits 11 to 15	0.6600
bits 24 to 28	1.0000	bits 10 to 14	0.2096
bits 23 to 27	1.0000	bits 9 to 13	0.3759
bits 22 to 26	0.0000	bits 8 to 12	0.9191
bits 21 to 25	0.0000	bits 7 to 11	0.8554
bits 20 to 24	0.0000	bits 6 to 10	0.5535
bits 19 to 23	0.1906	bits 5 to 9	0.4955
bits 18 to 22	0.0011	bits 4 to 8	0.0868
bits 17 to 21	0.3823	bits 3 to 7	0.1943
bits 16 to 20	0.8394	bits 2 to 6	0.8554
bits 15 to 19	0.2518	bits 1 to 5	0.7421

Result of DNA test for traditional MCNP RNG



Tested bits	p-value	Tested bits	p-value	Tested bits	p-value
bits 31 to 32	1.0000	bits 20 to 21	0.4937	bits 9 to 10	0.4550
bits 30 to 31	1.0000	bits 19 to 20	0.0613	bits 8 to 9	0.4737
bits 29 to 30	1.0000	bits 18 to 19	0.2383	bits 7 to 8	0.7834
bits 28 to 29	1.0000	bits 17 to 18	0.4831	bits 6 to 7	0.4063
bits 27 to 28	1.0000	bits 16 to 17	0.0925	bits 5 to 6	0.8959
bits 26 to 27	0.1777	bits 15 to 16	0.0197	bits 4 to 5	0.3438
bits 25 to 26	0.0000	bits 14 to 15	0.7377	bits 3 to 4	0.3972
bits 24 to 25	0.0000	bits 13 to 14	0.7171	bits 2 to 3	0.8986
bits 23 to 24	0.0000	bits 12 to 13	0.0309	bits 1 to 2	0.5407
bits 22 to 23	0.0000	bits 11 to 12	0.2803		
bits 21 to 22	0.0000	bits 10 to 11	0.8440		



- Less significant (lower) bits of RNs fail the tests.
- These failures in less significant bits are caused by the shorter period than the significant bits.

Drawback of LCGs with power-of-two modulus

The (r+1)-th most significant bit has period length at most 2^{-r} times that of the most significant bit.

• However, these failures do not have a significant impact in the practical use.



• Test program

```
integer(8) :: i
integer(8), parameter :: NumGeneratedRNs = 1000000000
!real(8) :: rang ! For MCNP4
real(8) :: RN_initial, RN_last
real(8) :: dummy
```

```
!call random ! For MCNP4
call RN_init_problem( new_standard_gen = 1 )
```

```
RN_initial = rang()
```

```
do i = 2, NumGeneratedRNs-1
dummy = rang()
end do
```

```
RN_last = rang()
```

.



- Comparison between MCNP4 and MCNP5
- Generate 1 billion RNs.

	MCNP4	MCNP5	MCNP4/MCNP5
CPU (sec) No optimization (/optimization:0)	290.0	97.1	3.0
CPU (sec) Local optimization (/optimization:1)	191.7	77.2	2.5
CPU (sec) Full optimization (/optimization:4)	188.4	78.1	2.4

Platform : Windows 2000, Intel Pentium III 1GHz Compiler : Compaq Visual Fortran Ver.6.6



- The 63-bit LCGs extended from the MCNP RNG fail the spectral test.
- L'Ecuyer's 63-bit LCGs pass all the tests and their multipliers are excellent judging from the spectral test.
- These 63-bit LCGs are implemented in the RNG package for MCNP5
- The MCNP5 RNG is ~2.5 times faster than the MCNP4 RNG.





Fundamentals of Monte Carlo Particle Transport



Lecture 3

Random Sampling

"Anyone who considers arithmetical methods of producing random digits is, of course, in a state of sin." John Von Neuman, 1951

Forrest B. Brown

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• Los Alamos

Probability ?

What are the odds of

•	Being audited by the IRS this year	100	to	1
•	Losing your luggage on a U.S. flight	176	to	1
•	Being dealt 4 aces on an opening poker hand	4,164	to	1
•	Being struck by lightning in your lifetime	9,100	to	1
•	Being hit by a baseball at a major league game	300,000	to	1
•	Drowning in your bathtub this year	685,000	to	1
•	Winning Lotto in the Illinois lottery with 1 ticket	12,900,000	to	1
•	Winning the grand prize in the Reader's			
	Digest sweepstakes	199,500,000	to	1





Continuous Probability Density

f(x) = probability density function (PDF) f(x) ≥ 0 Probability {a ≤ x ≤ b} = $\int_{a}^{b} f(x) dx$ Normalization: $\int_{-\infty}^{\infty} f(x) dx = 1$

Discrete Probability Density

$$\{ f_k \}, k = 1,...,N, \text{ where } f_k = f(x_k)$$

$$f_k \ge 0$$

$$\text{Probability}\{ x = x_k \} = f_k$$

$$\text{Normalization: } \sum_{k=1}^N f_k = 1$$



а

b



X



The key to Monte Carlo methods is the notion of *random sampling*.

• The problem can be stated this way:

Given a probability density, f(x), produce a sequence of \hat{X} 's. The \hat{X} 's should be distributed in the same manner as f(x).



- The use of random sampling distinguishes Monte Carlo from other methods
- When Monte Carlo is used to solve the integral Boltzmann transport equation:
 - Random sampling models the outcome of physical events (e.g., neutron collisions, fission process, sources,)
 - Computational geometry models the arrangement of materials

• Probability Density Function (PDF)

f(x) = probability density function (PDF)
f(x) ≥ 0
Probability {a ≤ x ≤ b} =
$$\int_{a}^{b} f(x) dx$$

Normalization: $\int_{-\infty}^{\infty} f(x) dx = 1$



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• Cumulative Distribution Function (CDF)

$$F(x) = \int_{-\infty}^{x} f(x') dx'$$

$$0 \le F(x) \le 1$$

$$\frac{dF(x)}{dx} \ge 0$$

$$F(-\infty) = 0, \quad F(\infty) = 1$$



Monte Carlo & Random Sampling



Monte Carlo Codes

Categories of random sampling

- Random number generator \rightarrow uniform PDF on (0,1)
- Sampling from analytic PDFs → normal, exponential, Maxwellian, …
 - Sampling from tabulated PDFs → angular PDFs, spectrum, ...

For Monte Carlo codes...

•

- Random numbers, ξ , are produced by the RN generator on (0,1)
- Non-uniform random variates are produced from the ξ's by:
 - Direct inversion
 - Rejection methods
 - Transformations
 - Composition (mixtures)
 - Sums, products, ratios, ...
 - Table lookup + interpolation
 - Lots (!) of other tricks
- Typically < 10% of total CPU time

Random Sampling Methods



Pseudo-Random Numbers

- Not strictly "random", but good enough
 - Pass statistical tests for randomness
 - Reproducible sequence
- Uniform PDF on (0,1)
- Must be easy to compute

Linear Congruential Method

– Algorithm

• Usage

- In algorithms, usually denote RN uniform on (0,1) by ξ
- In codes, invoke basic RN generator by: r = ranf()
- Each new usage of ξ or ranf() generates a **new** RN





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-∞



Direct Sampling

• **Direct solution of** $\hat{\mathbf{x}} = \mathbf{F}^{-1}(\boldsymbol{\xi})$

Solve for
$$\hat{\mathbf{x}}$$
: $\xi = \int f(\mathbf{x}) d\mathbf{x}$

- Sampling procedure
 - Generate ξ
 - Determine $\hat{\chi}$ such that $F_{\mathbf{k}}$) = ξ



Advantages

- Straightforward mathematics & coding
- "High-level" approach

Disadvantages

- Often involves complicated functions
- In some cases, F(x) cannot be inverted (e.g., Klein-Nishina)



Rejection Sampling

Von Neumann

"..... it seems objectionable to compute a transcendental function of a random number."

Select a bounding function, g(x), such that

- $c \cdot g(x) \ge f(x)$ for all x
- g(x) is an easy-to-sample PDF

Sampling Procedure:

• sample $\hat{\mathbf{x}}$ from $g(\mathbf{x})$: $\hat{\mathbf{x}} \leftarrow \mathbf{G}^{-1}(\boldsymbol{\xi}_1)$

• test:
$$\xi_2 \cdot cg(\hat{\mathbf{x}}) \leq f(\hat{\mathbf{x}})$$

if **true** \rightarrow accept \hat{x} , done if **false** \rightarrow reject \hat{x} , try again

f(x) f(x) $x \rightarrow$

Advantages

Simple computer operations

Disadvantages

"Low-level" approach, sometimes hard to understand

Random Sampling – Discrete PDFs



F₁

Discrete PDF's

Discrete PDF



F_N = 1





Sampling from Discrete PDF's — Conventional Procedure



Step (2) requires a table search

- linear table searches require O(N) time use when N small
- binary table searches require O(ln₂N) time u

- use when N large

For some discrete PDFs, Fk's are not precomputed.

· linear search, with Fk's computed on-the-fly as needed

Random Sampling – Discrete PDFs

Example — Sampling from Discrete Uniform PDF

Discrete Uniform PDF

 $f_k = 1 / N$, k = 1, ..., N

 $F_k = k / N$, $F_0 = 0$, $F_N = 1$

Sampling procedure:

Could use table search method,



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Easier, for this special case:

$$K \leftarrow \lfloor 1 + N\xi \rfloor$$
, where $\lfloor y \rfloor$ is the "floor" function,

largest integer < y

f(x)

Note: must be sure that $\lfloor 1 + N\xi \rfloor \le N$



- Multigroup Scattering
 - Scatter from group \mathbf{g} to group $\mathbf{g'}$, where $1 \le g' \le G$

$$f_{g'} = \frac{\sigma_{g \to g'}}{\sum_{k=1}^{G} \sigma_{g \to k}}$$

- Selection of scattering nuclide for a collision
 - K = number of nuclides in composition

$$f_{k} = \frac{N^{(k)}\sigma_{s}^{(k)}}{\sum_{j=1}^{K}N^{(j)}\sigma_{s}^{(j)}}$$



Sampling from Discrete PDF's — Alias Method

Any discrete PDF can be converted into "Alias sampling" form

original PDF: { f_k }, k=1, ..., N

where f_k = probability of selecting $x = x_k$

aliased PDF: { q_k, i_k, }, k=1, ..., N
where
$$\frac{1}{N} \cdot q_k$$
 = prob. of selecting $\hat{x} = x_k$
 $\frac{1}{N} \cdot (1 - q_k)$ = prob. of selecting $\hat{x} = x_{i_k}$

Alias sampling procedure:

Select <u>uniformly</u> for \hat{k} : $\hat{k} \leftarrow \lfloor 1 + N\xi_1 \rfloor$

Select <u>either</u> \hat{k} or its "alias" $i_{\hat{k}}$: if $\xi_2 < q_{\hat{k}}$, $\hat{x} \leftarrow x_{\hat{k}}$, otherwise, $\hat{x} \leftarrow x_{i_c}$

.....(continued on next page)



Sampling from Discrete PDF's — Alias Method (continued)

Why bother with "alias sampling" ?

- → No table search needed, requires O(1) time
- → Sampling time is constant & independent of size of PDF
- → Vectorizes completely & efficiently
- → Fastest possible way to sample discrete PDFs
- → Invented by Brown (who later found out Walker did it 3 yr earlier)

Creating the "aliased PDF" amounts to converting an N-way tree from arbitrary branching probabilities with single outcomes

to

uniform branching probabilities with dual outcomes

(See FB Brown & RACER coding for the set up algorithm)





Example – Sampling from uniform PDF in range (a,b), Histogram with 1 bin



 $\mathbf{x} \leftarrow \mathbf{a} + \boldsymbol{\xi} \cdot (\mathbf{b} - \mathbf{a})$



Example – Sampling from histogram with 2 bins



$$p_{1} = Prob\{ x_{0} < x < x_{1} \} = A_{1} / (A_{1}+A_{2})$$

$$p_{2} = Prob\{ x_{1} < x < x_{2} \} = A_{2} / (A_{1}+A_{2})$$

$$p_{1} + p_{2} = 1$$

Two-step sampling procedure:

1. Select a bin, b:

If $\xi_1 < \mathbf{p}_1$,	select	b = bin 1
otherwise,	select	b = bin 2

2. Sample x within bin:

$$\mathbf{x} \leftarrow \mathbf{x}_{b-1} + \xi_2 \cdot (\mathbf{x}_b - \mathbf{x}_{b-1})$$



Example – Sampling from Histogram PDF



Two-step sampling:

(1) Sample from discrete PDF to select a bin(2) Sample from uniform PDF within bin

• Discrete PDF:

$$p_k = f_k (x_k - x_{k-1}), \quad k = 1, ..., N, \quad \Sigma p_k = 1$$

- Generate ξ_1
- Use table search or alias method to select k
- Uniform sampling within bin k
 - Generate ξ_2

- Then,
$$x \leftarrow x_{k-1} + (x_k - x_{k-1}) \cdot \xi_2$$



Examples — Sampling from Linear PDF on (0,1)

$$f(x) = 2x, \quad 0 \le x \le 1$$

F(x) = $\int_{0}^{x} f(x') dx' = \int_{0}^{x} 2x' dx' = x^{2}$

Direct Sampling:

solving $F(\hat{x}) = \xi$ or $\hat{x} \leftarrow F^{-1}(\xi)$

gives:

x ← √ξ





Examples — Sampling from xⁿ PDF on (0,1)

$$\begin{split} f(\mathbf{x}) &= (n+1) \, \mathbf{x}^n \,, \qquad 0 \leq \mathbf{x} \leq 1 \\ F(\mathbf{x}) &= \mathbf{x}^{n+1} \\ \text{Solving } F(\hat{\mathbf{x}}) &= \xi \text{ gives:} \quad \hat{\mathbf{x}} \leftarrow \xi^{\frac{1}{n+1}} \end{split}$$

(Note: only for 0<x<1, does not apply to general intervals !)

Random Sampling – Continuous PDFs



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Examples – Sampling from an Exponential PDF

$$f(x) = \frac{1}{\lambda} \cdot e^{-x/\lambda}, \quad 0 \le x \le \infty$$
$$F(x) = \int_{0}^{x} f(x') dx' = 1 - e^{-x/\lambda}$$

Direct sampling:

Solve for x: $F(x) = \xi$

$$\begin{array}{ll} \text{Solving} \quad \xi = 1 - e^{-x/\lambda} & \text{gives:} \quad x \leftarrow -\lambda \cdot \ln(1 - \xi) \\ & \text{or} \\ & x \leftarrow -\lambda \cdot \ln\xi \end{array}$$

Although $(1-\xi) \neq \xi$, both ξ and $(1-\xi)$ are uniformly distributed on (0,1), so that we can use either in the random sampling procedure. (I.e., the numbers are different, but the distributions are the same)

Random Sampling – Direct vs. Rejection



Example — 2D Isotropic

$$f(\vec{p}) = \frac{1}{2\pi}$$
, $\vec{p} = (u, v)$

Rejection (old vim)

SUBROUTINE AZIRN_VIM(S, C) IMPLICIT DOUBLE PRECISION (A-H, O-Z) 100 R1=2.*RANF() - 1. R1SQ=R1*R1 R2=RANF() R2SQ=R2*R2 RSQ=R1SQ+R2SQ IF(1.-RSQ)100,105,105 105 S=2.*R1*R2/RSQ C=(R2SQ-R1SQ)/RSQ RETURN END

Direct (racer, new vim)

```
subroutine azirn_new( s, c )
implicit double precision (a-h,o-z)
parameter ( twopi = 2.*3.14159265 )
phi = twopi*ranf()
c = cos(phi)
s = sin(phi)
return
end
```



Random Sampling – Direct vs. Rejection



Example — Watt Spectrum

$$f(x) = \frac{2e^{-ab/4}}{\sqrt{\pi a^3 b}} e^{-x/a} \sinh \sqrt{bx} , \quad 0 < x$$

Rejection (mcnp)

- Based on Algorithm R12 from <u>3rd Monte Carlo Sampler</u>, Everett & Cashwell
- Define K = 1 + ab/8, $L = a \{K + (K^2 1)_{1/2}\}$, M = L/a 1

• Set
$$x \leftarrow -\log \xi_1$$
, $y \leftarrow -\log \xi_2$

1

• If $\{y - M(x+1)\}^2 \le bLx$, accept: return (Lx) otherwise, reject

Direct (new vim)

Sample from Maxwellian in C-of-M, transform to lab

$$\mathbf{w} \leftarrow \mathbf{a} \left(-\log \xi_1 - \log \xi_2 \cos^2 \frac{\pi}{2} \xi_3 \right)$$
$$\mathbf{x} \leftarrow \mathbf{w} + \frac{\mathbf{a}^2 \mathbf{b}}{4} + (2\xi_4 - 1) \sqrt{\mathbf{a}^2 \mathbf{b} \mathbf{w}}$$

(assume isotropic emission from fission fragment moving with constant velocity in C-of-M)

Unpublished sampling scheme, based on original Watt spectrum derivation



Example — Linear PDF

 $f(\mathbf{x}) = 2\mathbf{x}, \qquad 0 \le \mathbf{x} \le 1$

Rejection

(strictly - this is not "rejection", but has the same flavor)

- $\begin{array}{ll} \text{if } \xi_1 \geq \xi_2, \quad \text{then} \quad \hat{\mathbf{x}} \leftarrow \xi_1 \\ & \text{else} \quad \hat{\mathbf{x}} \leftarrow \xi_2 \end{array}$
- or
- $\hat{\mathbf{x}} \leftarrow \max{(\xi_1, \xi_2)}$
- or

$$\hat{\mathbf{x}} \leftarrow |\boldsymbol{\xi}_1 - \boldsymbol{\xi}_2|$$

Direct

.

$$F(\mathbf{x}) = \mathbf{x}^2, \qquad 0 \le \mathbf{x} \le 1$$
$$\hat{\mathbf{x}} \leftarrow \sqrt{\xi}$$

Random Sampling – Direct vs. Rejection



Pi	robability Density Function	Direct Sampling Method
Linear:	f(x) = 2x, $0 < x < 1$	x ← √ξ
Exponential:	$f(x) = e^{-x}, 0 < x$	x ← −logξ
2D Isotropic:	$f(\vec{p}) = \frac{1}{2\pi}$, $\vec{p} = (u, v)$	$u \leftarrow \cos 2\pi \xi_1$ $v \leftarrow \sin 2\pi \xi_1$
3D Isotropic:	$f(\vec{\Omega}) = \frac{1}{4\pi}$, $\vec{\Omega} = (u, v, w)$	$u \leftarrow 2\xi_1 - 1$ $v \leftarrow \sqrt{1 - u^2} \cos 2\pi \xi_2$ $w \leftarrow \sqrt{1 - u^2} \sin 2\pi \xi_2$
Maxwellian:	$f(x) = \frac{2}{T\sqrt{\pi}} \sqrt{\frac{x}{T}} e^{-x/T}, \qquad 0 < x$	$\mathbf{x} \leftarrow \mathbf{T} \left(-\log \xi_1 - \log \xi_2 \cos^2 \frac{\pi}{2} \xi_3 \right)$
Watt Spectrum:	$f(x) = \frac{2e^{-ab^{4}}}{\sqrt{\pi a^{3}b}} e^{-x/a} \sinh \sqrt{bx} , 0 < x$	$w \leftarrow a \left(-\log \xi_1 - \log \xi_2 \cos^2 \frac{\pi}{2} \xi_3 \right)$ $x \leftarrow w + \frac{a^2 b}{4} + (2\xi_4 - 1) \sqrt{a^2 b w}$
Normal:	$f(x) = \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{x-\mu}{\sigma}\right)^2}$	$x \leftarrow \mu + \sigma \sqrt{-2\log \xi_1} \cos 2\pi \xi_2$



Machine Considerations

Vector Hardware

- Since ~1980, direct methods have been recommended for vectorization & high performance on *cray*, *cyber-205*, *sx-3*, *cm-2*,
 - Vector concepts apply directly to pipelined RISC cpu's (e.g., *rs6000*, *i860*, *Fujitsu* μ-*vp*,...)

Math Libraries

Many routines in math libraries are now table-driven, hence very fast

→ fast *sin*, *cos*, *sqrt*, *log*, & *exp* functions

RISC + Compiler Technology

- Pipelining, concurrent ops, simple instructions, register-to-register ops, 64-bit hardware, better instruction scheduling,
 - → fast arithmetic (even for double-precision)
 - ➔ Today, the most expensive operations are

- load/store	(memory access)
— IFGOTO	(flush/fill instruction stack)


Software Considerations

"Rules of thumb" for M.C. algorithm design have changed

- Never take the square root of a random number
- Avoid using sin, cos, log, exp,
- · Use IF ... GOTO ... to avoid arithmetic
- Random numbers are cheap, arithmetic is expensive

Direct sampling methods have advantages

- Clear, succinct coding easier to verify & maintain
- Cpu time is comparable to rejection
- Direct methods vectorize efficiently

If a specific number of samples, M, is needed from a single distribution:

- Naive approach repeat the sampling procedure M times
- Stratified sampling approach
 - partition the sample space into M disjoint regions of equal probability
 - produce 1 sample from each region



- Stratified sampling considerations
 - F(x) must be known & easy to partition
 - The number of partitions, M, must be known in advance
 - Must be relatively easy to sample within each given partition
 - Stratification improves the "coverage"
 - Stratified sampling reduces variance, at little or no computing cost



Random Sampling – Rejection Method

- Rejection sampling methods are useful when it is difficult or impossible to invert F(x), or when F(x) is no known
- Example Selection of initial source sites in a reactor
 - Select a trial site:

$$\mathbf{x}' \leftarrow \mathbf{x}_1 + (\mathbf{x}_2 - \mathbf{x}_1) \cdot \boldsymbol{\xi}_1$$
$$\mathbf{y}' \leftarrow \mathbf{y}_1 + (\mathbf{y}_2 - \mathbf{y}_1) \cdot \boldsymbol{\xi}_2$$

- If (x',y') is inside a fuel pin (shaded region), then <u>accept</u> (x',y').
- Otherwise, <u>reject</u> (x',y') and repeat
- Efficiency of rejection sampling ~ (volume source region) / (total volume)







It is sometimes useful to sample from an alternate PDF

$$f(x) dx = \left[\frac{f(x)}{g(x)}\right]g(x) dx = h(x)g(x) dx$$

& then "correct" the result via either weight factors or a 2nd sampling stage

Weighted Sampling

- To sample x from f(x),
 - first, sample \hat{x} from g(x)
 - then, multiply the "weight" assigned to \hat{x} by $\frac{\text{right answer}}{\text{wrong answer}} = \frac{f(\hat{x})}{g(\hat{x})}$
- Note that g(x) must be >0 whenever f(x)>0.
- Also, g(x) must be normalized so that $\int g(x) dx = 1$

Example — survival biasing of collisions

If a collision occurs, $P_{survive} = \frac{\Sigma_S}{\Sigma_T}$ is the probability of surviving. Instead of sampling the outcome, always choose survival & multiply the "weight" by $P_{survive}$



Combined Russian Rouletting & Splitting

- Russian Roulette kill off some particles, but conserve total weight
- Splitting create extra identical particles, but conserve total weight
- Definitions

wgt = Particle weight

For the region containing the particle:

w_{high} = upper bound on weight, if wgt larger — split
 w_{low} = lower bound on weight, if wgt lower — roulette
 w_{ave} = weight to assign survivors, w_{low} < w_{ave} < w_{high}

Then,

wgt / wave = probability of surviving split/roulette

· Combined game for split/roulette:

if
$$wgt < w_{low}$$
 or $w_{high} < wgt$,
create **n** particles of weight w_{ave} , where $\mathbf{n} \leftarrow \left| \frac{wgt}{w_{ave}} + \xi \right|$

Random Sampling – Example



Weighted Sampling Example — Effective Free-gas Model for Scatter with Bound Hydrogen

- Given a neutron with initial energy E₀, E₀ > .625 eV
- · For scattering with free hydrogen (target-at-rest), PDF for scatter to E is

$$f_{\mathsf{FREE}}(\mathsf{E}_0 \to \mathsf{E}) = \frac{1}{\mathsf{E}_0}, \qquad 0 \le \mathsf{E} \le \mathsf{E}_0$$

· For scattering with bound hydrogen (free-gas), PDF for scatter to E is

down-scatter:
$$f_{BOUND}(E_0 \rightarrow E) = \frac{erf\sqrt{E/kT}}{E_0 - kT/2}, \quad 0 \le E \le E_0$$

up-scatter:
$$f_{BOUND}(E_0 \rightarrow E) = \delta(E - E_0), \quad E > E_0$$

P(upscatter) = $\frac{kT}{E_0 + kT/2}$

Sampling scheme for Ê, û:

First, sample \hat{E} , $\hat{\mu}$ using target-at-rest scattering model. Then,

If $\xi < P(upscatter)$, set $\hat{E} \leftarrow E_0$, $\hat{\mu} \leftarrow 1$, then exit

Otherwise, modify weight by factor

$$\frac{f_{\text{BOUND}}\left(E_{0} \rightarrow \hat{E}\right)}{f_{\text{FREE}}\left(E_{0} \rightarrow \hat{E}\right)} = \frac{\text{erf}\sqrt{\hat{E}/kT}}{1 - kT/(2E_{0})}$$

and set $\hat{\mu} \leftarrow \bar{\mu}_{\text{BOUND}}\left(E_{0} \rightarrow \hat{E}\right)$



Random Sampling

- Source
 - Fixed sites uniform PDF + rejection
 - Fission sites discrete PDF + stratified sampling
 - Energy piecewise-linear PDF (binary table search + linear PDF)
 - Direction isotropic 3D PDF
- Tracking
 - free-flight distance exponential PDF
 - delta-tracking rejection sampling of pseudo- & real collisions
- Russian Roulette & Splitting
 - discrete PDF + weights
- Collisions
 - Survival biasing weights
 - Select phase & nuclide discrete PDF (on-the-fly)
 - Epithermal
 - Scattering angle equally-probable histograms (uniform discrete PDF + uniform in bin)
 - Inelastic: energy discrete PDF (aliased), then uniform within group
 - n2n weights
 - modified free-gas weights
 - Thermal
 - multigroup discrete PDF for group-to-group (aliased), linear PDF for μ
 - S(α,β) discrete PDF (aliased) & uniform PDF sampling
 - Direction polar angle from uniform PDF
 - Fission bank discrete PDF + weights





Every Monte Carlo code developer who works with random sampling should own & read these references:

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Fundamentals of Monte Carlo Particle Transport



Lecture 4



Forrest B. Brown Diagnostics Applications Group (X-5) Los Alamos National Laboratory

Engineering Model vs. Computational Model





Model Generation

Engineering Model



Large-scale Computation

Computational Model

Post-processing

Engineering Model

- Model Generation
 - Focus on engineering productivity
 - Describes "reality" to computer
 - Interactive, batch, or CAD
- Large-scale Computation
 - Focus on efficiency & capabilities
 - Data structures should be compact & regular
 - Computational model often hidden from user
- Post-Processing
 - Interpretation of results
 - Visualization

Modeling vs. Computation • Los Alamos Element geometry • Geometry computation $\mathbf{O}\mathbf{O}$ $(\bigcirc$ Model construction 0 Elements —> assemblies • $\left[O\right]$ Ο $\mathbf{O}\mathbf{O}$ 0 0 $\mathbf{O}\mathbf{O}$ \bigcirc \bigcirc Ο (O) Assemblies -> core • Core + peripherals • -> 3D model

Monte Carlo Geometry





mcnp, rcp, vim, racer, sam-ce, tart, morse, keno, tripoli, mcbend, monk, o5r, recap, andy,.....

Development of particular geometric capabilities is driven by applications:

- Shielding & experiment analysis
 - Irregular geometry
 - Moderate number of regions & compositions
- Reactor core analysis
 - Regular geometry
 - Very many regions & compositions

Computational Algorithm – Geometric View



Repe	eat for all	cycles						
	Repeat	t for all his	stories in c	ycle				
		Repea	Repeat until collision					
			Repea	Repeat for each universe level				
				. Repeat for surfaces of 3D region				
					Distar	nce calculation		
	•							
			Bound	lary crossing	g			
			Neigh	bor search				
			Roule	tte/split				
		Collisi	on analysis	5				
		Roule	tte/split			1 reactor calculation requires		
						~10 ¹⁰ distance calculations		



- Every point in space that a particle could possibly reach must be defined in the geometry model
- 3D **volumes** are defined by their bounding **surfaces**
 - Boundary representation
 - Combinatorial geometry, with either surfaces or primitive bodies
 - CSG constructive solid geometry, tree structure with boolean operators
 - Mesh geometry
- A cell number is assigned to each 3D volume
 - For some codes, disjoint volumes must have different cell numbers
 - For MCNP & others, disjoint volumes may have the same cell number
- A material number is assigned to each cell
 - Composition is assumed to be uniform & homogeneous within cell
- **Tallies** are defined for particular cells or surfaces, reaction types, & estimator types



Locate

Given a point in space, determine what cell it is in

• Distance to surface

Given a point & direction in a particular cell, determine the distance to the next surface of that cell

Neighbor search

For a particle which has hit a bounding surface of a cell, determine the cell to be entered next

Boundary conditions

For a particle which has hit a cell bounding surface declared to be periodic or reflecting, determine the new position & direction and cell to be entered next



Particle

Position = (x,y,z), Direction = (u,v,w)

Cell number

(i,j,k), indices in mesh

- Locate
 - i: binary search to find x-interval containing x
 - j: binary search to find y-interval containing y
 - k: binary search to find z-interval containing z
- Distance
 - Use signs of (u,v,w) to select surfaces, then compute 3 distances:
 - if u>0, $d_x = (x_{i+1}-x)/u$, otherwise $d_x = (x_i-x)/u$... similar for $d_y \& d_z$
 - Distance: $d = min(d_x, d_y, d_z)$
- Neighbor search
- Boundary conditions







- MCNP uses a "combinatorial geometry" based on surfaces
 - Define **surfaces**
 - Define **cells** using surfaces & operators (intersection, union, complement)
 - Can also group cells together into a **universe**, and embed that universe inside another cell
 - Can also group cells together into a universe, repeat that universe in a **lattice** arrangement, and embed that universe inside another cell
 - Assign **materials** to cells
 - Assign **other properties** to cells (e.g., importance weights)
 - Define **tallies** using cell or surface numbers



In MCNP, surface types include:

1st order:	planes
2nd order:	spheres, cylinders, cones, ellipsoid,
	hyperboloid, paraboloid, general quadric
4th order:	elliptical & circular torus (axes parallel to x, y, or z)

[see tables on next 2 slides]

• Quadratic polynomial for surface:

 $F(x,y,z) = ax^{2} + by^{2} + cz^{2} + dxy + eyz + fzx + gx + hy + jz + k$

- Surface is defined by: F(x,y,z) = 0
- Surface is either infinite or closed
- Normalization convention: factor of leading 2nd order term is positive

MCNP Surfaces



Mnemonic	Туре	Description	Equation	Card Entries		
Р	Plane	General	Ax + By + Cz - D = 0	ABCD		
PX		Normal to X-axis	x - D = 0	D		
PY		Normal to Y-axis	y - D = 0	D		
PZ		Normal to Z-axis	z - D = 0	D		
SO	Sphere	Centered at Origin	$x^2 + y^2 + z^2 - R^2 = 0$	R		
S		General	$(x-\bar{x})^{2} + (y-\bar{y})^{2} + (z-\bar{z})^{2} - R^{2} = 0$	$\overline{x} \ \overline{y} \ \overline{z} \ R$		
SX		Centered on X-axis	$(x-\bar{x})^2 + y^2 + z^2 - R^2 = 0$	$\bar{x} R$		
SY		Centered on Y-axis	$x^{2} + (y - \overline{y})^{2} + z^{2} - R^{2} = 0$	$\overline{y} R$		
SZ		Centered on Z-axis	$y^{2} + y^{2} + (z - \bar{z})^{2} - R^{2} = 0$	$\overline{z} R$		
C/X	Cylinder	Parallel to X-axis	$(y - \bar{y})^2 + (z - \bar{z})^2 - R^2 = 0$	<u>y</u> z R		
C/Y		Parallel to Y-axis	$(x - \bar{x})^{2} + (z - \bar{z})^{2} - R^{2} = 0$	$\overline{x} \ \overline{z} \ R$		
C/Z		Parallel to Z-axis	$(x - \bar{x})^{2} + (y - \bar{y})^{2} - R^{2} = 0$	$\overline{x} \ \overline{y} R$		
CX		On X-axis	$y^2 + z^2 - R^2 = 0$	R		
CY		On Y-axis	$x^2 + z^2 - R^2 = 0$	R		
CZ		On Z-axis	$x^2 + y^2 - R^2 = 0$	R		

Table 3.1: MCNP Surface Cards

MCNP Surfaces



				2	
K/X	Cone	Parallel to X-axis	$(y - \bar{y})^{2} + (z - \bar{z})^{2} - t(x - \bar{x}) = 0$	$\bar{x} \bar{y} \bar{z} t^2 \pm 1$	
K/Y		Parallel to Y-axis	$\sqrt{\left(x-\overline{x}\right)^2 + \left(z-\overline{z}\right)^2} - t(y-\overline{y}) = 0$	$\bar{x} \bar{y} \bar{z} t^2 \pm 1$	
K/Z		Parallel to Z-axis	$\sqrt{(x-\bar{x})^{2}+(y-\bar{y})^{2}}-t(z-\bar{z}) = 0$	$\bar{x} \bar{y} \bar{z} t^2 \pm 1$	
KX		On X-axis	$\sqrt{y^2 + z^2} - t(x - \overline{x}) = 0$	$\bar{x} t^2 \pm 1$	
KY		On Y-axis	$\sqrt{x^2 + z^2} - t(y - \overline{y}) = 0$	$\bar{y} t^2 \pm 1$	
ΚZ		On Z-axis	$\sqrt{x^2 + y^2} - t(z - \overline{z}) = 0$	$\overline{z} t^2 \pm 1$ ± 1 used only for 1 sheet cone	
SQ	Ellipsoid Hyperboloid Paraboloid	Axis parallel to X-, Y-, or Z-axis	$A(x-\overline{x})^{2} + B(y-\overline{y})^{2} + C(z-\overline{z})^{2}$ $+ 2D(x-\overline{x}) + 2E(y-\overline{y})$ $+ 2F(z-\overline{z}) + G = 0$	А В С D Е F G x ў ž	
GQ	Cylinder Cone Ellipsoid Hyperboloid Paraboloid	Axes not parallel to X-, Y-, or Z-axis	$Ax^{2} + By^{2} + Cz^{2} + Dxy + Eyz$ $+Fzx + Gx + Hy + Jz + K = 0$	A	
ТХ	Elliptical or circular torus.	$(x-\overline{x})^2/B^2 + (\sqrt{x})$	$(y-\bar{y})^2 + (z-\bar{z})^2 - A)^2 / C^2 - 1 = 0$	$\bar{x} \ \bar{y} \ \bar{z} \ A \ B \ C$	
TY	X-, Y-, or Z-axis	$(y-\bar{y})^2/B^2 + (\sqrt{(x-\bar{x})^2 + (z-\bar{z})^2} - A)^2/C^2 - 1 = 0$ $\bar{x}\bar{y}\bar{z}ABC$			
ΤZ		$(z-\bar{z})^2/B^2 + (\sqrt{(x-\bar{x})^2 + (y-\bar{y})^2} - A)^2/C^2 - 1 = 0 \qquad \bar{x}\bar{y}\bar{z}ABC$			
XYZP Surfaces defined by points See pages 3–15 and 3–17					

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For a given point in space, (x,y,z), and surface equation, F(x',y',z')=0, the sense of the point with respect to the surface is defined as:

Inside the surface,	sense < 0,	if	F(x,y,z) < 0
Outside the surface,	sense > 0,	if	F(x,y,z) > 0
On the surface,	sense = 0,	if	F(x,y,z)=0

[Must be careful to consider computer roundoff!]





- A surface divides space into positive & negative sides
 - MCNP convention: +1 = positive side of surface 1
 - -1 = negative side of surface 1



- If not sure which side is + or -, pick a point & substitute into surface function, F(x,y,z) — see if result is + or -



MCNP convention: +1 -2 == intersection of positive side of surface 1 and negative side of surface 2



Union of Sides



- MCNP convention: colon signifies a union operator
 - -1:2 == union of negative side of surface 1 with positive side of surface 2



Cells

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- A cell is defined to be the
 - Intersection of half-spaces defined by a list of signed surface numbers



- Union of half-spaces defined by signed surface numbers

Example: cell 1 +1 : -2

- The complement of another cell (i.e., volume NOT in other cell)

Example: cell 1 #5

- A combination of the above

Example: cell 1 (+1 -2): 3 #5

Cells



• Cells do not have to be convex



Cells may involve discontiguous regions





Given point (x,y,z), determine which cell it is contained in:

```
For( cell = 1 ... n_cells ) {
    Foreach surf in cell {
        Evaluate S<sub>surf</sub> = sign{ F<sub>surf</sub>(x,y,z) }
        Does S<sub>surf</sub> match the sense from the cell definition?
    }
    If all surface-senses for (x,y,z) matched the cell definition,
        then exit & return cell as the result
}
```



Given point (x,y,z) in cell I, determine the distance to the cell boundary

```
d <-- infinity
```

```
Foreach surf in cell I {
```

```
If surf is part of the external boundary of cell I {
```

```
Evaluate d<sub>surf</sub> = smallest positive root of
F<sub>surf</sub>(x+du, y+dv, z+dw) = 0
d = min(d, d<sub>surf</sub>)
}
return the value of d
```



• When a cell boundary is reached, what's on the other side?



- Most codes build "neighbor lists" during tracking
 - For each bounding surface of cell, remember list of neighbors
 - Initially, neighbor lists are empty
 - Check all cells having surface in common, until one is found satisfying all sense conditions for the particle position
 - Save it
 - Later, check neighbor lists first, only do search if necessary
- Neighbor search is expensive at first, cheap later
- Tracking speeds up as calculation progresses

Embedded Geometry – Universes



• In most real-world applications, there is a need for modeling detailed geometry with many repeating units



- All production Monte Carlo codes provide capabilities for multiple levels of nested geometry
 - Called "universes" in MCNP
 - A collection of cells may be grouped into a "universe"
 - Universe may be embedded in another cell, with the universe 'clipped' by the cell boundaries

Universes & Lattices



Universe 1 – cells for detailed fuel pin



Universe 2 – lattice of cells for fuel assembly

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Universe 2, with cells filled by Universe 1

Universe 3 - lattice of cells for reactor

"Real world" - final geometry



• Some Monte Carlo codes use primitive bodies rather than surfaces for defining cells (e.g., MORSE, KENO, ITS, VIM)

SPH – sphere	ELL – ellipsoid
BOX – box	REC – right elliptic cylinder
RPP – box	RHP – hexagonal prism
RCC – cylinder	HEX – hexagonal prism
WED – wedge	ARB – arbitrary polyhedron
	TRC – truncated cone

- Usually called "combinatorial geometry"
 - Invented by MAGI in ~1956, used in SAM-CE & other codes
 - Space **inside** the body has a **negative** sense, **outside** a **positive** sense
 - Boolean operators AND, OR, NOT may be used to combine bodies (like MCNP's intersection, union, & complement operators)
 - MCNP allows body geometry input (calls them "macrobodies"), but internally converts them to lists of surfaces



- **Simple cells** are those which can be constructed using only **intersections**, with no union operators
- Some Monte Carlo codes require that all cells be simple cells. Union operators are not allowed.
- Tracking through simple cells is fast, at the expense of more complex geometry input & setup
 - For simple cells, the logic to find the distance to boundary is simple check the distance to each of the cell surfaces & keep only the smallest positive distance





Consider the example at the left.

Using the union operator, the cell is described by: +1 : -2

Without the union operator, separate cells must be defined & then assigned the same material properties:

Special Topics - Distance Calculations



- 3D Surface
 - F(x,y,z) = 0
 - Linear: $\nabla F = constant$
 - Quadratic: $\nabla F = f(x,y,z), \nabla^2 F = constant$
- Distance calculation
 - **S** = directed distance from (x_0, y_0, z_0) along (u, v, w) to F(x, y, z)=0= smallest positive root of **F** $(x_0+su, y_0+sv, z_0+sw) = 0$
 - General form:

 $As^{2} + 2Bs + C = 0$, $D = B^{2} - AC$

- 27 combinations of A, B, C >0, =0, <0
- Only 12 yield valid solutions:

s = -C/(2B)	if	(A=0, C<0, B>0)	or	(A=0, C>0, B<0)
s = (-B-√D)/A	if	(A>0, C>0, B<0, D>0)	or	(A<0, C>0, B>0, D>0)
	or	(A<0, C>0, B<0, D>0)	or	(A<0, C>0, B=0, D>0)
s = (-B+√D)/A	if	(A>0, C<0, B>0, D>0)	or	(A>0, C<0, B<0, D>0)
	or	(A>0, C<0, B=0, D>0)	or	(A<0, C<0, B>0, D>0)
	or	(A>0, C=0, B<0, D>0)	or	(A<0, C=0, B>0, D>0)
Special Topics - Distance Calculations



- Noting that $C = F(x_0, y_0, z_0) =$ sense at (x_0, y_0, z_0) , the valid solutions can be simplified using the known surface sense §:

s' = -C/(2B)	if	(A=0, D>0)
s' = (-B-√D)/A	if	(A≠0, D>0, §>0)
s' = (-B+√D)/A	if	(A≠0, D>0, §<0)
S' = ∞	otherwise	

And

s = s'	if s'>0
=∞	otherwise



- Most MC codes check for coincident surfaces & eliminate one of them (replacing it by the other)
- The tolerance for coincident surfaces usually defaults to a small separation distance (e.g., 1.e-4 cm). For problems with unusual geometry (very small or very large), this may have to be changed in the code or code input.





Stochastic Geometry & HTGR Modeling

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Introduction



• Much interest lately in analyzing HTGRs

- Fuel kernels with several layers of coatings
- Very high temperatures
- Contain fission products
- Safety aspects ...

Double heterogeneity problem

- Fuel kernels randomly located within fuel elements
- Fuel elements may be "compacts" or "pebbles" (maybe random)
- Challenging computational problem

Monte Carlo codes can faithfully model HTGRs

- Full 3D geometry
- Multiple levels of geometry, including embedded lattices
- Random geometry ?????

Example – Very High Temperature Gas Cooled Reactor





Pyrolytic Carbon Silicon Carbide Porous Carbon Buffer Uranium Oxycarbide

TRISO Coated fuel particles (left) are formed into fuel rods (center) and inserted into graphite fuel elements (right).



P. E. MacDonald, et al., "NGNP Preliminary Point Design – Results of the Initial Neutronics and Thermal-Hydraulic Assessments During FY-03", INEEL/EXT-03–00870 Rev. 1, Idaho National Engineering and Environmental Laboratory (2003).

Example – GT-MHR Modeling



Fig. 4. Fragments of double-heterogeneous GT-MHR (HTR3): (A) fuel element (compact) cross section with coated fuel particles; (B) magnified view of coated fuel particles: spherical kernels of PuO2-x are surrounded by protective coatings made of PyCbuffer, PyC I, SiC and PyC II layers correspondingly. The same structure is valid for particles containing burnable poison-natural Er2O3.

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Fig. 3. Fragments of single-heterogeneous GT-MHR (HTR2): (A) an active core structure: three rings of hexagonal fuel columns; (B) magnified view of a separate fuel assembly. Fuel compacts are presented in small grey circles, burnable poison compacts in light grey. Bigger diameter holes stand for He channels, while the rest material represents the graphite matrix.

core

(A)

Fig. 4

(B)

Example – Pebble Bed Experiments at Proteus Facility





Fig. 3. Cross section of the odd layers of the hcp config rations, case 4 with polyethylene rods. The fit with the visualization tools of the MCNP program.



Difilippo, F.C., Monte Carlo Calculations of Pebble Bed Benchmark Configurations of the PROTEUS Facility. Nucl. Sci. Eng. 143, 240–253 (2003).



Fuel kernel lattice



Fuel kernel



Fig. 8. Fuel kernel with the four coatings at each location of the cubic lattice and inside the fuel region of the pebble shown in Fig. 6.

Pebbles



4 - 35

MCNP Models for HTGRs



• Existing MCNP geometry can handle:

- 3D description of core
- Fuel compacts or lattice of pebbles
 - Typically, hexagonal lattice with close-packing of spherical pebbles
 - Proteus experiments: ~ 5,000 fuel pebbles
 - \sim 2,500 moderator pebbles
- Lattice of fuel kernels within compact or pebble
 - · Typically, cubic lattice with kernel at center of lattice element
 - Proteus experiments: ~ 10,000 fuel kernels per pebble
 - ~ 50 M fuel kernels, total
- Could introduce random variations in locations of a few thousand cells in MCNP input, but **not** a few million.
- See papers by: Difilippo, Plukiene et al, Ji-Conlin-Martin-Lee, etc.



- When a neutron enters a new lattice element, a transformation is made to the neutron's position & direction to the local coordinates of the universe embedded in that lattice element. [standard MCNP]
- Users can flag selected universes as "stochastic" [new]
 - A neutron entering a lattice element containing a stochastic universe undergoes the normal transformations.
 - Then, additional **random translations** are made:

$$x \leftarrow x + (2\xi_1 - 1) \cdot \delta_x$$
$$y \leftarrow y + (2\xi_2 - 1) \cdot \delta_y$$
$$z \leftarrow z + (2\xi_3 - 1) \cdot \delta_z$$

 Then, tracking proceeds normally, with the universe coordinates fixed until the neutron exits that lattice element

MCNP5 Stochastic Geometry



• Neutron on lattice edge, about to enter embedded universe



• Embedded universe,

before random translation after random translation





• Track normally, until neutron exits the lattice element





• On-the-fly random translations of embedded universes in lattice

- Does not require any extra memory storage
- Very little extra computing cost only 3 random numbers for each entry into a stochastic universe

• For K-effective calculations (KCODE problems)

- If fission occurred within fuel kernel, should have source site in next cycle be at same position within fuel kernel
- Need to save $\delta_x, \delta_v, \delta_z$ along with neutron coordinates in fission bank
- On source for next cycle, apply δ_x , δ_y , δ_z after neutron pulled from bank
- To preserve mass exactly, rather than on the average stochastically, must choose δ_x , δ_y , δ_z so that fuel kernels are not displaced out of a lattice element



Numerical Results — HTGR Fuel Kernels



• Infinite array of TRISO fuel kernels in graphite matrix

 Fuel kernel geometry & composition taken from the NGNP Point Design (MacDonald et al. 2003)

TRISO Fuel Kernel Geometry and Composition

Region	Name	Outer radius	Composition	Density
#		(μ)		(g/cc)
1	Uranium oxycarbide	175	UCO (UC ^{.5} O ^{1.5})	10.5
2	Porous carbon buffer	275	С	1.0
3	Inner pyrolytic carbon	315	С	1.9
4	Silicon carbide	350	SiC	3.2
5	Outer pyrolytic carbon	390	С	1.9

Calculations run 4 ways:

- 1. Fixed lattice with centered kernels
- 2. Fixed lattice with random kernels [MCNP stochastic geometry]
- 3. Multiple lattice realizations
- 4. Box of randomly place kernels



Fixed lattice with centered kernels

- 5x5x5 cubical lattice
- Lattice edge chosen to preserve the specified packing fraction.
- Fuel kernels centered within the cubical cells
- Reflecting boundaries on the outer surfaces
- Essentially same as Difilipo, Plukiene et al, Ji-Conlin-Martin-Lee
- No random geometry, standard MCNP5 calculations





• Fixed lattice with random kernels [MCNP stochastic geometry]

- 5x5x5 cubical lattice
- Lattice edge chosen to preserve the specified packing fraction.
- Fuel kernels randomly placed **on-the-fly** within the cubical cells
- Reflecting boundaries on the outer surfaces
- Uses new MCNP5 stochastic geometry



Fuel kernel displaced randomly within lattice element each time that neutron enters



- Multiple lattice realizations
 - 5x5x5 cubical lattice
 - Lattice edge chosen to preserve the specified packing fraction.
 - Fuel kernels randomly placed in job input within the cubical cells
 - Reflecting boundaries on the outer surfaces
 - Uses standard MCNP5
 - 25 separate calculations, each with different location of kernels in the input files



1 realization, fixed lattice with kernel locations chosen randomly in problem input & held constant during each MCNP calculation



- Box of randomly placed fuel kernels
 - Single box with 125 fuel kernels
 - Box size chosen to preserve the specified packing fraction.
 - Fuel kernels randomly placed in job input within the box (using RSA algorithm, Random Sequential Addition)
 - Reflecting boundaries on the outer surfaces
 - Uses standard MCNP5
 - 25 separate calculations, each with different location of kernels in the input files

2 different realizations of "truly random" cases:







MCNP5 Results for Infinite Lattices of Fuel Kernels

#	Method	K-effective
1	Fixed 5x5x5 lattice with centered spheres	1.1531 ± 0.0004
2	Fixed 5x5x5 lattice with randomly located spheres ("on the fly")	1.1515 ± 0.0004
3	Multiple (25) realizations of 5x5x5 lattice with randomly located spheres	1.1513 ± 0.0004
4	Multiple (25) realizations of randomly packed (RSA) fuel "box"	1.1510 ± 0.0003

Conclusions



- The new stochastic geometry treatment for MCNP5 provides an accurate and effective means of modeling the particle heterogeneity in TRISOL particle fuel
 - Same results as (brute-force) multiple realizations of random geometry input with standard MCNP
 - Negligible difference from "truly random" multiple realizations

The results indicate that:

- The neutronic effect of using a fixed lattice is negligible
- The effect of choosing either a centered spheres or randomly located spheres is also small, at least for the specific fuel geometry that was analyzed during this study

• Future work

- Examination of finite geometries, including cylindrical fuel compacts, hexagonal fuel blocks, and full core configurations.
- We will also consider lattices other than simple cubic lattices, such as BCC, FCC, and HCP lattices.

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Fundamentals of Monte Carlo Particle Transport





Collision Physics

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Monte Carlo Calculations





mcnp, rcp, vim, racer, sam-ce, tart, morse, keno, tripoli, mcbend, monk, o5r, recap, andy,.....

- Geometry routines determine the cell & material in that cell
- Collision routines model the physical interactions with the material
 - Random sampling from PDFs determined by cross-section data
 - Continuous: flight distance, exit E & direction,
 - Discrete: select nuclide, select interaction type, secondaries,

Collision Physics







- After a particle emerges from source or collision, or if the particle is on a cell bounding surface:
 - Randomly sample the free-flight distance to the next interaction
 - If the distance-to-interaction is less than the distance to cell boundary, then move the particle to the interaction point
 - Collision physics at the interaction point:
 - Determine which isotope the interaction is with
 - Determine which interaction type for that isotope
 - Determine the energy & direction of the exiting particle
 - Determine if secondary particles were produced
 - Biasing + weight adjustments
 - Tallies of quantities of interest



- Given a particle at (x_0, y_0, z_0) with direction (u, v, w) in cell I containing material M, sample the free-flight distance to the next interaction
 - Σ_T = total macroscopic cross-section in material M
 = sum{ N^jσ_T^j }, where j = isotopes in material M
 = probability of any interaction per unit distance, units cm⁻¹
 - **PDF for flight distance s**, where $0 \le s \le \infty$, f(s) = {prob interaction p.u.d} • {prob travelling dist s w/o interact} = $\Sigma_T \exp(-\Sigma_T s)$
 - Sampling procedure $F(s) = 1 - \exp(-\Sigma_T s) \quad -> \quad s = -\ln(1-\xi) / \Sigma_T$

We are assuming here that material M is uniform & homogeneous



•
$$\Sigma_{\mathrm{T}} = \sum_{j} \mathsf{N}^{(j)} \sigma_{\mathrm{T}}^{(j)}$$

٠

where j = isotopes in material M

Probability that collision is with isotope j

$$p_{j} = \frac{N^{(j)}\sigma_{T}^{(j)}}{\sum_{k} N^{(k)}\sigma_{T}^{(k)}}$$

- $\{p_i\}$ = set of discrete probabilities for selecting collision isotope
- { P_i } = discrete CDF, P_i = sum{ p_i , i=1,j }, P_0 =0
- **Discrete sampling** for collision isotope **k** table search to determine **k** such that $P_{k-1} \le \xi \le P_k$



• For collision isotope k,

 $\sigma_{T} = \sigma_{elastic} + \sigma_{inelastic} + \sigma_{capture} + \sigma_{fission} + \dots$

- $p_i = \sigma_i / \sigma_T$ = probability of reaction type j for isotope k
- { p_i } = set of discrete probabilities for selecting reaction type j
- { P_j } = discrete CDF, P_j = sum{ p_i , i=1,j }, P_0 =0
- **Discrete sampling** for reaction type **j** table search to determine **j** such that $P_{j-1} \le \xi \le P_j$



- In many applications, **survival biasing** is an effective variance reduction technique
 - Survival biasing is also called implicit absorption, nonabsorption weighting, or (loosely) implicit capture
 - $\sigma_{\rm T} = \sigma_{\rm absorption} + \sigma_{\rm scatter} \qquad (absorption = disappearance)$
 - Probability that particle survives collision = $P_{surv} = \sigma_{scatter}/\sigma_T$
 - Probability that particle is absorbed (killed) = $1 P_{surv}$
- Disallow absorption of particle, & then adjust particle weight to ensure a fair game
 - Tally absorption of wgt• $(1-P_{surv})$
 - Multiply particle weight by P_{surv}
 - When selecting reaction type, don't consider probability of absorption

Sampling Exit Energy & Direction



- Given a collision isotope k & reaction type j, the random sampling techniques used to determine the exit energy and direction, E' and (u',v',w'), depend on
 - Conservation of energy & momentum
 - Scattering laws either equations or tabulated data

• Examples

- Isotropic scattering in lab system
- Multigroup scattering
- Elastic scattering, target-at-rest
- Inelastic scattering, MCNP
- Other collision physics, MCNP



- Elastic scattering from infinite-mass target nucleus
 - No change in energy:

E' = E

- Sample direction from isotropic scattering PDF, $f(u',v',w') = 1 / 4\pi$

$$\phi = 2\pi \, \xi_1$$

$$u' = 2\xi_2 - 1$$

$$v' = sqrt(1-u'^2) cos(\phi)$$

$$w' = sqrt(1-u'^2) sin(\phi)$$



Multigroup approach

- Divide energy range into intervals (groups)
- Use average cross-sections for each group,

 σ_{Tg} = total cross-section for group g

- Use discrete transfer matrix for group-to-group scatter, $\sigma_{aa'}$ = cross-section for scatter from group g to group g'

Multigroup scattering

- For particle with energy E, determine initial energy group g
- Select exit energy group g' by discrete sampling from $\sigma_{qq'}$

$$p_{g'} = \frac{\sigma_{g \to g'}}{\sum_{k=1}^{G} \sigma_{g \to k}}$$

- Sample exit energy uniformly within bound of group g'
- Direction
 - For P0 scattering use procedure for isotropic lab scatter
 - For P1 scattering sample mu from linear PDF, then select new direction (see next section on elastic scatter)

Elastic Scattering, Target-at-rest





- Sample μ_{cm} from tabulated PDF data, f(μ_{cm})
- Use kinematics to get E'_{lab} & μ_{lab}
- Sample polar angle ϕ uniformly on $(0,2\pi)$
- Rotate particle direction using μ_{lab} & ϕ

Sampling the Scattering Direction-cosine, μ_{cm}



- Typical representations for $f(\mu_{cm})$
 - Histogram or Equiprobable Histogram PDF



- Piecewise linear PDF







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$$E' = E \bullet \frac{A^2 + 2A\mu_{cm} + 1}{(A+1)^2}$$

$$\mu_{\text{lab}} = \frac{\mathbf{I} + \mathbf{A}\mu_{\text{cm}}}{\sqrt{\mathbf{A}^2 + 2\mathbf{A}\mu_{\text{cm}} + 1}}$$

Where A = (mass target)/(mass particle)



..........

- Rotation from (u,v,w) to (u',v',w') using μ_{lab} & ϕ

 $\mu = \mu_{lab}$

 $\phi = 2\pi\xi$

$$\mathbf{u}' = \mu \mathbf{u} + \frac{\sqrt{1 - \mu^2} (\mathbf{u} \mathbf{w} \cos \phi - \mathbf{v} \sin \phi)}{\sqrt{1 - \mathbf{w}^2}}$$

$$\mathbf{v}' = \mu \mathbf{v} + \frac{\sqrt{1 - \mu^2} (\mathbf{v} \mathbf{w} \cos \phi + \mathbf{u} \sin \phi)}{\sqrt{1 - \mathbf{w}^2}}$$

If μ close to 1, special coding may be used to avoid roundoff

 θ_{lab}

$$\mathbf{w}' = \mu \mathbf{w} - \sqrt{1 - \mu^2} \sqrt{1 - \mathbf{w}^2} \cos \phi$$


- Law 1 ENDF law 1 Equiprobable energy bins
- Law 2 Discrete photon energies
- Law 3 ENDF law 3 Inelastic scatter from nuclear levels
- Law 4 ENDF law 4 Tabular distribution
- Law 5 ENDF law 5 General evaporation spectrum
- Law 7 ENDF law 7 Simple Maxwell fission spectrum
- Law 9 ENDF law 9 Evaporation spectrum
- Law 11 ENDF law 11 Energy dependent Watt spectrum
- Law 22 UK law 2 Tabular linear functions of incident energy out
- Law 24 UK law 6 Equiprobable energy multipliers
- Law 44 ENDF law 1, lang 2, Kalbach-87 correlated energy-angle scatter
- Law 61 ENDF law 11, lang 0,12, or 14 correlated energy-angle scatter
- Law 66 ENDF law 6 N-body phase space distribution
- Law 67 ENDF law 7 correlated energy-angle scatter

Other Collision Physics – MCNP



- Emission from fission
- Delayed neutron emission
- $S(\alpha,\beta)$ scattering for thermal neutrons
- Free-gas scattering for neutrons
- Probability tables for the unresolved resonance energy range for neutrons
- Photoelectric effect
- Pair production
- Compton scattering (incoherent)
- Thomson scattering (coherent)
- Fluorescent emission
- Photonuclear reactions
- Electron interactions condensed history approach
 - Stopping power, straggling, angular deflections
 - Bremsstrahlung
 - K-shell impact ionization & Auger transitions
 - Knock-on electrons



• Consider a collision which results in fission

wgt • $v\sigma_F/\sigma_T$ = expected number of fission neutrons produced per collision

• To sample the number of neutrons produced in the collision

Let $r = wgt \cdot v\sigma_F / \sigma_T$ n = int[r]

Then, Produce n fission neutrons with probability 1 and an additional fission neutron with probability r-n

Assign a weight of r/n to each

Example: wgt• $v\sigma_F/\sigma_T = 1.75$ If $\xi < .75$, produce 2 neutrons, otherwise produce 1 or

```
Produce int[ 1.75 + \xi ] neutrons
```



Alternative Schemes for Flights/Collisions





Conventional scheme

- Particle weight constant during flight
- Use Σ_{T} to determine distance-to-collision, ~s = -In\xi / Σ_{T}
- Change weight only on collisions
- For pathlength absorption estimator, tally wgt·s· Σ_A
- Most common scheme for reactors & shielding applications

Continuous absorption

- Particle weight decreases continuously during flight, due to absorption

$$wgt(s) = wgt_0 \cdot e^{-\Sigma_A s}$$

- Use $\Sigma_{\rm S}$ to determine distance-to-scattering, $~{\rm s}$ = -In ξ / $\Sigma_{\rm s}$
- For pathlength absorption estimator, tally wgt_o \cdot (1 e^{- $\Sigma_A s$})
- No absorption in collision
- Typical use in astrophysics (Implicit Monte Carlo codes)



Random Sampling – Flight Distance



Sampling the free-flight distance, s

Samp

- To simulate the free-flight of particles through the problem geometry, need to randomly sample the distance to collision
- PDF for free-flight distance, s, along the current direction:

$$f(s) = \Sigma_{T}(s) \cdot \exp\left(-\int_{0}^{s} \Sigma_{T}(x) dx\right)$$

- If $\Sigma_{T}(x)$ is constant within a region, the PDF simplifies to

$$f(s) = \Sigma_{T} \cdot exp(-\Sigma_{T} \cdot s)$$

ling procedure is then:
$$\hat{\mathsf{s}} \leftarrow \frac{-\mathsf{ln}\xi}{\Sigma_{\tau}}$$

- For multiple regions, can stop particle at each boundary & resample s. Why is this OK ?
 - Note that prob. of traversing region is $Prob{\hat{s}}$

$$\geq s \Big\} = 1 - \int_{0}^{s} \Sigma_{\mathsf{T}} e^{-\Sigma_{\mathsf{T}} x} dx = e^{-\Sigma_{\mathsf{T}} s}$$

- For 2 regions, note that $e^{-\Sigma x_1} \cdot e^{-\Sigma x_2} = e^{-\Sigma(x_1+x_2)}$ = prob. of traversing <u>both</u> regions



"Regular" Tracking

- Move particles through one region at a time, until collision occurs
- Can be expensive if many regions must be traversed before collision



- "Regular" tracking procedure, when Σ_{T} constant within each region:
 - Sample a flight distance, s', using Σ_T for current region: $s' \leftarrow \frac{-\ln \xi}{\Sigma_T}$
 - If $s' < d_{boundary}$, move particle by s', then analyze the collision
 - Otherwise, move particle by d_{boundary}, enter next region, repeat until collision occurs



Delta Tracking

- A type of rejection method for sampling the free-flight distance
- Also called Woodcock tracking, fast tracking, or hole tracking
- Useful when Σ_{T} varies rapidly over the flight path





- For delta tracking, a fictitious cross-section Σ^* is used, rather than $\Sigma_T(s)$
 - $-\Sigma^*$ should be chosen to be $\geq \Sigma_T(s)$ for all possible points along path
 - $-\Sigma^*$ may be a function of energy, or region, or not
 - $\Sigma^{\star} = \Sigma_{\mathsf{T}}(\mathsf{s}) + \Sigma_{\delta}(\mathsf{s}) = \text{ constant}, \qquad \Sigma_{\delta}(\mathsf{s}) \geq 0 \ \text{ for all } \mathsf{s}{>}0$

where $\Sigma_{\delta}(s) = cross-section$ for "delta-scattering", i.e., scatter with no change in energy or direction, a fictitious scattering event, or "pseudo-collision"



• For many problems of interest, Σ_{T} varies within a cell

- Charged particle transport continuous slowing down along the flight path due to interactions with electron field in material
- $-\Sigma_{T}$ increases along the flight path



- For most MC codes, a procedure called delta tracking is used in sampling the free-flight distance
 - Also called Woodcock tracking, fast tracking, pseudo-collision method, hole tracking, …
 - Involves biased sampling using a larger Σ_T , followed by rejection sampling to assure a fair game



- **Basic idea:** Sample flight distance using Σ^* , then reject collision point if $\xi > \Sigma_T(s) / \Sigma^*$
- Using Σ* rather than Σ_T(s) gives an interaction probability per unit distance that is too large, hence a flight distance that is too short. Rejection scheme compensates for this.



Sampling procedure

- Sample s' from $f(s) = \Sigma^* \exp(-\Sigma^* s)$: $s' = -\ln(1-\xi_1)/\Sigma^*$
- Move the particle a distance s'
- $\begin{array}{ll} & \mbox{if} \quad \xi_2 < \Sigma_{\rm T}({\rm s}')/\Sigma^* \;, & \mbox{"real" collision: do collision physics} \\ & \mbox{otherwise,} & \mbox{"delta" collision: no change in E, (u,v,w), wgt} \end{array}$
- Repeat until a real collision occurs
- Delta tracking can be effective if Σ^* is not too different from the "average" $\Sigma_T(s)$
- Delta tracking can be ineffective if Σ* >> Σ_T(s) for most values of s, so that sampling efficiency is low
- Delta tracking is also frequently used for tracking through reactor fuel assemblies, where the geometry is a regular lattice.



Proof: Delta tracking is an unbiased method for sampling the free-flight distance

Consider the probability of traversing a distance **s** along the flight path without undergoing a (real) collision, **P(s)**

- $\Sigma^* = \Sigma_T(s) + \Sigma_{\delta}(s) = constant, \qquad \Sigma^* \ge \Sigma_T(s) \text{ and } \Sigma_{\delta}(s) \ge 0 \text{ for all } s > 0$
- For convenience, define optical thickness for real & delta scatter:

$$\tau(s) = \int_{0}^{s} \Sigma_{T}(x) dx \qquad \qquad \tau_{\delta}(s) = \int_{0}^{s} \Sigma_{\delta}(x) dx$$

Note that, by definition,

$$\Sigma^* s = \tau(s) + \tau_{\delta}(s), \qquad \Sigma^* s \ge \tau(s)$$



- For a particular flight, there could be <u>exactly</u> 0, 1, 2, ..., ∞ delta-collisions before a real collision occurs
- Let P(sln) = probability of traversing distance s along the flight path with <u>exactly</u> n delta collisions

Then,

$$\mathbf{P}(\mathbf{s}) = \sum_{n=0}^{\infty} \mathbf{P}(\mathbf{s} \mid \mathbf{n})$$

$$P(s \mid 0) = e^{-\Sigma^{*}s}$$

$$P(s \mid 1) = \int_{0}^{s} P(x \mid 0)\Sigma_{\delta}(x)P(s - x \mid 0)dx = \int_{0}^{s} e^{-\Sigma^{*}x}\Sigma_{\delta}(x)e^{-\Sigma^{*}(s - x)}dx = \tau_{\delta}(s)e^{-\Sigma^{*}s}$$

$$P(s \mid 2) = \int_{0}^{s} P(x \mid 1)\Sigma_{\delta}(x)P(s - x \mid 0)dx = \int_{0}^{s} \tau_{\delta}(x)e^{-\Sigma^{*}x}\Sigma_{\delta}(x)e^{-\Sigma^{*}(s - x)}dx$$

$$= \int_{0}^{s} \tau_{\delta}(x)\Sigma_{\delta}(x)e^{-\Sigma^{*}s}dx = \frac{[\tau_{\delta}(x)]^{2}}{2}e^{-\Sigma^{*}s}$$

$$P(s \mid n) = \int_{0}^{s} P(x \mid n - 1)\Sigma_{\delta}(x)P(s - x \mid 0)dx = \frac{[\tau_{\delta}(s)]^{n}}{n!}e^{-\Sigma^{*}s}$$

Special Topic – Delta Tracking



Then, the total probability of traversing a distance s without undergoing a (real) collision is

$$\mathsf{P}(\mathsf{s}) = \sum_{\mathsf{n}=0}^{\infty} \mathsf{P}(\mathsf{s} | \mathsf{n}) = \sum_{\mathsf{n}=0}^{\infty} \frac{[\tau_{\delta}(\mathsf{s})]^{\mathsf{n}}}{\mathsf{n}!} e^{-\Sigma^{\star} \mathsf{s}} = e^{\tau_{\delta}(\mathsf{s}) - \Sigma^{\star} \mathsf{s}} = e^{-\tau(\mathsf{s})} = \exp\left(-\int_{0}^{\mathsf{s}} \Sigma_{\mathsf{T}}(\mathsf{x}) d\mathsf{x}\right)$$

This is the correct result, identical to the normal sampling of the flight path (without delta tracking)





For many problems of interest, Σ_{T} varies within a cell

- Charged particle transport
 - Continuous slowing down along the flight path due to interactions with electron field in material
 - $-\Sigma_{T}$ increases along the flight path
- Atmospheric transport
 - Air density varies with altitude
- Depleted reactor
 - Fuel & poison distribution varies due to burnup





- Stepwise approximation
 - Subdivide geometry
 - Constant material properties within each step



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Woodcock tracking

- Also called delta tracking, fast tracking, pseudo-collision method, hole tracking, ...
- Involves biased sampling the flight distance using a larger Σ_T , followed by rejection sampling to assure a fair game



• Optical depth along flight path

$$-\tau(s) = \int_{x}^{x+s} \Sigma_T(x') dx' \qquad \Sigma T(x) \text{ is finite,} \quad \Sigma T(x) \ge 0$$

- Note that
$$\frac{d\tau(s)}{ds} = \Sigma_T(x+s), \qquad 0 \le \frac{d\tau}{ds} \le \infty$$

- To explicitly allow for the case of no collision,
 - PNC = probability of no collision

$$-P_{NC}=e^{-\tau(\infty)}$$

• Probability density function (pdf) for the flight distance s:

$$f(s) = P_{NC} \cdot \delta(s = \infty) + (1 - P_{NC}) \cdot \frac{1}{G} \frac{d\tau}{ds} e^{-\tau(s)}$$

- Where
$$G = \int_{0}^{\infty} \frac{d\tau(s)}{ds} e^{-\tau(s)} ds = 1 - e^{-\tau(\infty)} = 1 - P_{NC}$$

Sampling the flight distance in varying media



• Random sampling of the Monte Carlo free-flight path requires solving the following equation for s, the flight path:

$$\xi = \int_{0}^{\xi} f(x) dx$$

$$\xi = P_{NC} \cdot H(s, \infty) + (1 - P_{NC}) \cdot \frac{1}{G} \cdot \left(1 - e^{-\tau(s)}\right)$$

• Common case: Σ_T independent of x

$$\tau(s) = \Sigma_T \cdot s, \quad \frac{d\tau}{ds} = \Sigma_T, \quad P_{NC} = 0, \quad G = 1, \quad f(s) = \Sigma_T \cdot e^{-\Sigma_T \cdot s}$$

- With solution:

$$s = -\frac{\ln(1-\xi)}{\Sigma_T}$$

Sampling the flight distance in varying media



Direct Numerical Sampling for the free-flight distance:

Step [1] If $\xi < PNC$, Then: No collision, set s= ∞ , exit Otherwise: Do Steps 2 & 3 **Step** [2] Define $\hat{\tau} = \tau(s)$ Sample $\hat{\tau}$ by solving $\xi = \frac{1}{G} \int_{0}^{\hat{\tau}} e^{-\tau} d\tau$, with $0 \le \hat{\tau} \le \tau(\infty)$ That is, sample from a truncated exponential PDF: $\hat{\tau} = -\frac{\ln(1-\xi \cdot G)}{\Sigma_{\tau}}$ Solve for s: $\hat{\tau} = \tau(s) = \int_{0}^{s} \Sigma_T(x+s') ds'$ **Step** [3] Analytic solution if possible, otherwise use Newton iteration



• Newton iteration to numerically solve for s:

$$s_0 = \hat{\tau} / \Sigma_T(x_0)$$
$$n = 0$$

Iterate:

$$n = n+1$$

$$g = \hat{\tau} - \tau(s_{n-1})$$

$$g' = dg/ds = -\Sigma_T(x_0 + s_{n-1})$$

$$s_n = s_{n-1} - g/g'$$

$$Stop if |s_n - s_{n-1}| < \varepsilon$$

- Notes:
 - Because g'<0, g(s) is monotone & there can be only one root
 - For cases where $\Sigma_T > 0$, Newton iteration guaranteed to converge
 - If $\Sigma_T(x)=0$ or very small, g' may be 0, leading to numerical difficulties
 - Remedied by combining Newton iteration with bisection if g' near zero
 - Typically only 1–5 iterations needed to converge s to within 10–6



- Represent material density by high-order, orthogonal polynomial expansion within each cell
 - Legendre polynomial representation for material density in cell

$$p(x) = \sum_{n=0}^{N} \frac{2n+1}{2} \cdot a_n \cdot P_n \left[\frac{2}{\Delta x} (x - x_{\min}) - 1 \right]$$

$$a_n = \frac{2}{\Delta x} \int_{x_{\min}}^{x_{\max}} \rho(x) P_n \left[\frac{2}{\Delta x} (x - x_{\min}) - 1 \right] dx$$

• Sample the free-flight distance to next interaction using a direct numerical sampling scheme (Brown & Martin)

$$\Sigma(x) = \frac{\rho(x)}{\rho_0} \cdot \Sigma_0, \qquad \tau(s) = \frac{\Sigma_0}{\rho_0} \cdot \int_x^{x \to s} \rho(x') \frac{dx'}{\mu}$$

- Use Newton iteration to solve nonlinear equation for flight path

References – Continuous Materials & Tallies



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Fundamentals of Monte Carlo Particle Transport

Lecture 6



Tallies & Statistics

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Monte Carlo Calculations





mcnp, rcp, vim, racer, sam-ce, tart, morse, keno, tripoli, mcbend, monk, o5r, recap, andy,.....

- During a history, tally the events of interest
- Upon completing a history, accumulate total scores & squares
- After completing all histories, compute mean scores & standard deviations



Given a function R(x), where x is a random variable with PDF f(x),

- Expected value of R(x) is
- Variance of R(x) is

$$\mu = \int \mathsf{R}(\mathsf{x}) f(\mathsf{x}) \, \mathsf{d}\mathsf{x}$$

$$\sigma^2 = \int R^2(x) f(x) \, dx - \mu^2$$

Monte Carlo method for estimating μ

- make N random samples $\hat{\mathbf{x}}_{i}$ from f(x)
- Then

$$\overline{\mathsf{R}} \approx \frac{1}{\mathsf{N}} \sum_{j=1}^{\mathsf{N}} \mathsf{R}(\hat{\mathsf{X}}_j)$$

- Central Limit Theorem states that for large N, the PDF of $\overline{\mathbf{R}}$ approaches a Gaussian distribution
- That is, if the Monte Carlo problem is repeated, **R** will be normally distributed

Laws of Large Numbers



Let $x_1, x_2, ..., x_N$ be a sequence of independent, identically distributed random variables each with a finite mean $E[x_i]=\mu$ and let

$$\overline{\mathbf{X}}_{\mathbf{N}} = \frac{1}{\mathbf{N}} \sum_{j=1}^{\mathbf{N}} \mathbf{X}_{j}$$

Weak Law of Large Numbers

For any $\varepsilon > 0$

$$\lim_{N\to\infty} |\mathbf{P}(|\overline{\mathbf{X}}_N - \boldsymbol{\mu}| > \epsilon) = \mathbf{0}$$

Tells how a sequence of probabilities converges

• Strong Law of Large Numbers

$$P\!\!\left(\lim_{N\to\infty}\left|\overline{\boldsymbol{x}}_{\!N}-\boldsymbol{\mu}\right| \! > \! \boldsymbol{\epsilon} \right) = \boldsymbol{0}$$

Tells how the sequence of IID random variables behaves in the limit



Central Limit Theorem

$$\lim_{N\to\infty} \operatorname{Prob}\left\{ \mu - a\frac{\sigma}{\sqrt{N}} \leq \overline{x} \leq \mu + b\frac{\sigma}{\sqrt{N}} \right\} = \frac{1}{\sqrt{2\pi}} \int_{-a}^{b} e^{-t^{2}} dt$$

- If
$$a = b = 1$$
, $\operatorname{Prob}\left\{\mu - \frac{\sigma}{\sqrt{N}} \le \overline{x} \le \mu + \frac{\sigma}{\sqrt{N}}\right\} = 68\%$

Note: 32% of the time, $\overline{\mathbf{X}}$ should be <u>outside</u> range $\mu \pm \frac{\sigma}{\sqrt{N}}$

$$- \text{ If } a = b = 2, \quad \text{Prob} \left\{ \mu - \frac{2\sigma}{\sqrt{N}} \le \overline{x} \le \mu + \frac{2\sigma}{\sqrt{N}} \right\} = 95\%$$

Note: 5% of the time, $\overline{\mathbf{x}}$ should be <u>outside</u> range $\mu \pm \frac{2\sigma}{\sqrt{N}}$



- For a given history, tally events of interest
 - Example surface crossings
 - For each particle crossing surface A, accumulate the weight each time a particle crosses that surface
 - A particular particle may cross the surface more than once
 - Progeny of that particle (e.g., another particle created by splitting) may also cross that surface one or more times
- When the history is complete, add the score & score² to accumulators for the problem

$$\begin{split} & \text{S1}_{\text{problem}} = \text{S1}_{\text{problem}} \ + (\text{S}_{\text{history}}) \\ & \text{S2}_{\text{problem}} = \text{S2}_{\text{problem}} + (\text{S}_{\text{history}})^2 \end{split}$$

 When all N histories are complete, compute final mean score & standard deviation

mean score =
$$\frac{1}{N} \cdot S1$$

std dev of mean = $\sqrt{\frac{1}{N-1} \left[\frac{S2}{N} - \left(\frac{S1}{N}\right)^2\right]}$

Variance of the Population vs. Mean



• Given a set of random samples, x₁, x₂, ..., x_N,

– Mean

$$\overline{\mathbf{x}} = \frac{1}{N} \sum_{j=1}^{N} \mathbf{x}_{j}$$

- Population variance

$$\sigma^{2} = \frac{1}{N} \sum_{j=1}^{N} x_{j}^{2} - \left(\frac{1}{N} \sum_{j=1}^{N} x_{j}\right)^{2} = \frac{1}{N} \sum_{j=1}^{N} x_{j}^{2} - \overline{x}^{2}$$

- Variance of the mean

$$\sigma_{\overline{x}}^2 = \frac{\sigma^2}{N}$$



• Tallies can be made for selected events & portions of phase space:

- Range of energies, $E_1 E_2$
- Range of particle times, $t_1 t_2$
- Specified cells
- Specified surfaces
- Specified range of $n \cdot \Omega$ for surface crossings
- Specified reaction cross-sections Σ_x
- Secondary particle production
- Energy deposited in cell
- Conditional events, e.g., absorption in cell B due to source in cell A
- Energy of neutrons causing fission
- Scattering from energy range $E_1 E_2$ to range $E_3 E_4$
- Etc.

Flux & Current

- Angular flux $\Psi(\mathbf{r}, \mathbf{E}, \Omega)$
- Flux

$$\phi(\mathbf{r}) = \int_{\mathsf{E}_1}^{\mathsf{E}_2} d\mathsf{E} \int_{4\pi} d\Omega \, \Psi(\mathbf{r}, \mathsf{E}, \Omega)$$

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- Scalar quantity
- Total distance traveled by all particles in a cm³ per second
- Units: distance / cm^3 -sec = 1 / cm^2 -sec
- Current
 - Number of particles crossing surface per second per unit area
 - Units: 1 / cm²-sec
 - Partial current: in + or direction only, J^+ or J^-
 - Net current = $J = J^+ J^-$

$$\mathbf{J}(\mathbf{r}) = \int_{\mathbf{E}_{1}}^{\mathbf{E}_{2}} d\mathbf{E} \int_{4\pi} d\Omega \ \vec{\mathbf{n}} \bullet \Omega \ \Psi(\mathbf{r}, \mathbf{E}, \Omega)$$

$$J^{+}(\mathbf{r}) = \int_{\mathsf{E}_{1}}^{\mathsf{E}_{2}} d\mathsf{E} \int_{\vec{n} \bullet \Omega > 0} d\Omega \ \vec{n} \bullet \Omega \ \Psi(\mathbf{r}, \mathsf{E}, \Omega) \qquad \qquad J^{-}(\mathbf{r}) = \int_{\mathsf{E}_{1}}^{\mathsf{E}_{2}} d\mathsf{E} \int_{\vec{n} \bullet \Omega < 0} d\Omega \ \vec{n} \bullet \Omega \ \Psi(\mathbf{r}, \mathsf{E}, \Omega)$$

Reaction Rates



$$\mathbf{R}_{\mathbf{X}}(\mathbf{r}) = \int_{\mathbf{E}_{1}}^{\mathbf{E}_{2}} d\mathbf{E} \int_{4\pi} d\Omega \, \Psi(\mathbf{r}, \mathbf{E}, \Omega) \, \Sigma_{\mathbf{X}}(\mathbf{r}, \mathbf{E})$$

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- Reactions per cm³ per sec
- Collision density $C(\mathbf{r}) = \int_{E_1}^{E_2} dE \int_{4\pi} d\Omega \Psi(\mathbf{r}, \mathbf{E}, \Omega) \Sigma_{\mathsf{T}}(\mathbf{r}, \mathbf{E})$
- Energy deposition (average per collision)

$$\mathbf{E}_{\text{deposited}}(\mathbf{r}) = \int_{\mathbf{E}_{1}}^{\mathbf{E}_{2}} d\mathbf{E} \int_{4\pi} d\Omega \Psi(\mathbf{r}, \mathbf{E}, \Omega) \Sigma_{\mathsf{T}}(\mathbf{r}, \mathbf{E}) K(\mathbf{r}, \mathbf{E})$$

where K(r, E) = average E deposited per collision



Analog Monte Carlo

- Faithful simulation of particle histories
- No alteration of PDFs (I.e., no biasing or variance reduction)
- At collision, particle is killed if absorption occurs
- Particle is born with weight = 1.0
- Weight unchanged throughout history until particle is killed
- Score 1.0 when tallying events of interest

Weighted Monte Carlo (non-analog)

- Alter the PDFs to favor events of interest
- Particle is born with weight = 1.0
- Weight is altered if biased PDF is used
- Typically, particle always survives collision & weight is reduced by P_{surv}
- Weight can also be changed by Russian roulette/splitting & other variance reduction techniques
- Score wgt when tallying events of interest
Tally Types

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Current tallies

- Surface crossing estimator

• Flux tallies

- Pathlength estimator
- Collision estimator
- Surface crossing estimator
- Next event estimator (point detector)

Reaction rate tallies

- Any of the above flux estimators times a cross-section
- Energy deposition tallies
 - Any of the above flux estimators times $\boldsymbol{\Sigma}_{T}$ times energy deposited per collision

Current Tallies



- For each particle crossing surface, tally the particle weight
- Divide by total starting weight & surface area to get current



- Typically, keep separate tally for outward partial current for each surface of a cell
- Can get net current by combining partial currents



- For each particle flight within a cell, tally (pathlength*weight)
- Divide by cell volume & total starting weight to get flux estimate



Flux Tally – Collisions



- Since $(\Sigma_T \phi)$ is collision rate, for each collision, tally (wgt/Σ_T) to estimate flux
- Divide by total starting weight & cell volume



Flux Tally – Surface Crossing



- Consider particles crossing a surface
 - Put a "box" of thickness a around the surface
 - Pathlength estimate of flux in the box

$$\phi = \frac{1}{W aA} \bullet \sum_{\substack{all \\ particles \\ crossing \\ surface}} where \quad \mu_j = \cos \theta_j$$

- Note that **a** cancels out
- Take the limit as **a**->0
- Surface crossing estimate of flux

$$\phi = \frac{1}{W A} \bullet \sum_{\substack{all \\ particles \\ crossing \\ surface}} \frac{wgt_j}{|\mu_j|}$$

wat



- Complication: wgt_i/μ_i can be very large for small μ_i
 - Usual solution, based on theory from FH Clark, "Variance of Certain Flux Estimators Used in Monte Carlo Calculations", Nucl.Sci. Eng. 27, 235–239 (1967)
 - For small lµl, that is, -ε<µ<ε, (where ε is small), if it is assumed that the flux is only isotropic or linearly anisotropic, then the expected value of 1/lµl is 2/ε.
- Actual tally procedure:
 - If $|\mu| < \epsilon$, then replace $|\mu|$ by $\epsilon/2$ to score an expected flux.
 - This results in a reliable variance, without affecting the flux estimate significantly.
- MCNP uses $\varepsilon = .1$. Many other codes use $\varepsilon = .01$



- Instead of estimating flux for a cell or surface, it may be useful to estimate flux at a point
 - Probability of a history trajectory going through a particular point is zero
- Use a "next event estimator" to get flux at a point
 - Regardless of the actual outcome of simulating a collision, estimate what would happen if the particle scattered exactly in the direction of a point detector

Expected
$$\phi$$
 score = wgť • $\frac{\mathsf{p}_{sc}(\mu)}{2\pi\mathsf{R}^2}$ • $\exp\left\{-\int_0^\mathsf{R} \Sigma_\mathsf{T}(\mathsf{E}')ds\right\}$





- Expected score has 1/R² singularity collisions close to detector can result in large scores
 - Point detector estimator has finite mean, but infinite variance due to 1/R² singularity
- To keep variance finite:
 - For collisions within radius \Re of detector, replace the factor

$$\frac{\exp\left\{-\int_{0}^{R} \Sigma_{T}(E')ds\right\}}{R^{2}}$$

by volume average assuming uniform collisions inside sphere

$$\frac{\int\limits_{0}^{\Re} e^{-\Sigma_{\mathsf{T}}(\mathsf{E}')\mathsf{s}} d\mathsf{s}}{\int\limits_{0}^{\Re} \mathsf{s}^{2} d\mathsf{s}} = \frac{1 - e^{\Sigma_{\mathsf{T}}(\mathsf{E}')\Re}}{\frac{1}{3}\Re^{3}\Sigma_{\mathsf{T}}(\mathsf{E}')}$$

– Typically choose ${\mathfrak R}$ to be ~half a mean free path



• Example – pathlength tallies

After each flight, tally

- Flux wgt∙d_i
- Total absorption $wgt \bullet d_i \bullet \Sigma_A$
- Nu-fission $\mathbf{wgt} \bullet \mathbf{d}_{j} \bullet \mathbf{v} \Sigma_{F}$
- U235 absorption $wgt \bullet d_i \bullet N^{U^{235}} \sigma_A^{U^{235}}$





Mesh Tallies

Impose a grid over the problem & tally flux or reaction rates in each grid cell



• Fission matrix

- Impose a grid over problem
- Tally F(I->J) for source in cell I causing fission in cell J
- For N cells in grid, N² tallies

RE & FOM

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$$\mathsf{RE} = \frac{\sigma_{\overline{x}}}{\overline{x}}$$

• Some codes report a Figure-of-Merit for selected tallies

$$= OM = \frac{1}{RE^2 \bullet T}$$

Where T =computer time used

- $RE^2 \sim 1/N$, where N is the total number of histories
- − T ~ N
- Therefore, FOM should be roughly constant
- Used for comparing effectiveness of different variance reduction schemes



- RE should decrease smoothly with $1/\sqrt{N}$ dependence as more histories are run
- Tallies are reliable only if "enough" histories traverse the portions of problem phase space being tallied
 - Undersampling can lead to questionable or erroneous values of the mean score & relative error
 - Indicators of undersampling:
 - Large RE, RE > .1
 - RE does not decrease smoothly as $1/\sqrt{N}$
 - A few histories have very large scores
- MCNP performs statistical checks on selected tallies to try to detect undersampling effects
 - Large RE
 - Variance of the variance (VOV)
 - Tally fluctuation charts (distribution of scores)
 - Slope of tails in tally fluctuation charts
 - Etc.



Given N sets of (mean,std-dev) for independent Monte Carlo calculations, (x_1,σ_1) , (x_2,σ_2) , ..., how should the results be combined?

$$\mathbf{w}_{j} = \frac{1}{\sigma_{j}^{2}} \qquad \qquad \mathbf{W} = \sum_{j=1}^{N} \frac{1}{\sigma_{j}^{2}}$$

$$\overline{\mathbf{x}} = \sum_{j=1}^{N} \frac{\mathbf{w}_{j}}{\mathbf{W}} \mathbf{x}_{j}$$

$$\sigma_{\overline{x}}^2 = \sum_{j=1}^{N} \frac{w_j}{W^2} = \frac{1}{W}$$

Weighting factors ~ $1/\sigma^2$



• Suppose 2 estimators, x and y, are correlated, such as the path & collision estimator for Keff

$$\overline{\mathbf{x}} = \frac{1}{N} \sum_{j=1}^{N} \mathbf{x}_{j} \qquad \overline{\mathbf{y}} = \frac{1}{N} \sum_{j=1}^{N} \mathbf{y}_{j}$$
$$\sigma_{\mathbf{x}}^{2} = \frac{1}{N} \sum_{j=1}^{N} \mathbf{x}_{j}^{2} - \overline{\mathbf{x}}^{2} \qquad \sigma_{\mathbf{y}}^{2} = \frac{1}{N} \sum_{j=1}^{N} \mathbf{y}_{j}^{2} - \overline{\mathbf{y}}^{2} \qquad \sigma_{\mathbf{xy}}^{2} = \frac{1}{N} \sum_{j=1}^{N} \mathbf{x}_{j} \mathbf{y}_{j} - \overline{\mathbf{x}} \cdot \overline{\mathbf{y}}$$

Minimum variance combination of x & y

$$\alpha = \frac{\sigma_y^2 - \sigma_{xy}^2}{\sigma_x^2 - 2\sigma_{xy}^2 + \sigma_y^2}$$

$$mean_{\mathbf{x},\mathbf{y}} = \alpha \,\overline{\mathbf{x}} + (\mathbf{1} - \alpha) \,\overline{\mathbf{y}}$$

std-dev_{x,y} =
$$\sqrt{\frac{\alpha^2 \sigma_x^2 + 2\alpha (1-\alpha) \sigma_{xy}^2 + (1-\alpha)^2 \sigma_y^2}{N-1}}$$



Continuously Varying Tallies



- Conventional Monte Carlo codes tally integral results
 - Tallies summed into bins
 - Zero-th order quantities
 - Stepwise approximation to results



• Higher order tallies

- Represent results by high-order, orthogonal polynomial expansion within each cell
- Make tallies for expansion coefficients
- Legendre polynomial representation for continuous tallies

$$\Phi(x) = \sum_{n=0}^{N} \frac{2n+1}{2} \cdot b_n \cdot P_n \left[\frac{2}{\Delta x} (x - x_{\min}) - 1 \right]$$

$$b_n = \frac{2}{\Delta x} \int_{x_{\min}}^{x_{\max}} \Phi(x) P_n \left[\frac{2}{\Delta x} (x - x_{\min}) - 1 \right] dx$$



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$$b_n = \frac{2}{\Delta x} \int_{x_{\min}}^{x_{\max}} \Phi(x) P_n \left[\frac{2}{\Delta x} (x - x_{\min}) - 1 \right] dx$$

• At collisions, tally $\frac{wgt}{\Sigma_T} \cdot P_n \left[\frac{2}{\Delta x} (x - x_{\min}) - 1 \right]$ for n=1..N

• At flights, tally
$$wgt \cdot \frac{1}{\mu} \int_{x}^{x+s} P_n \left[\frac{2}{\Delta x}(x'-x_{\min})-1\right] dx'$$
 for n=1,N

• Reconstruct $\Phi(x)$ and $\sigma_{\Phi}^2(x)$ from tallied coefficients

Continuous 2D Tallies – Example





Figure 2a. 9x9 Legendre expansion tally for thermal neutron flux across the fuel pin obtained in a 2 million history simulation.



Figure 2b. MCNP5 20×20 mesh tally for thermal neutron flux across the fuel pin obtained in a 2 million history simulation.

DP Griesheimer & WR Martin, **"Two Dimensional Functional Expansion Tallies for Monte Carlo Simulations,"** PHYSOR-2004, Chicago, IL (2004)

References – Continuous Materials & Tallies



- FB Brown, D Griesheimer, & WR Martin, "Continuously Varying Material Properties and Tallies for Monte Carlo Calculations", PHYSOR-2004, Chicago, IL (April, 2004)
- FB Brown & WR Martin, "Direct Sampling of Monte Carlo Flight Paths in Media with Continuously Varying Cross-sections", ANS Mathematics & Computation Topical Meeting, Gatlinburg, TN (April, 2003).
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- ER Woodcock, T Murphy, PJ Hemmings, TC Longworth, "Techniques Used in the GEM Code for Monte Carlo Neutronics Calculations in Reactors and Other Systems of Complex Geometry," *Proc. Conf. Applications of Computing Methods to Reactor Problems*, ANL-7050, p. 557, Argonne National Laboratory (1965).
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- J. Spanier, "Monte Carlo Methods for Flux Expansion Solutions of Transport Problems," Nucl. Sci. Eng., **133**, 73 (1999).





Fundamentals of Monte Carlo Particle Transport



Lecture 7

Eigenvalue Calculations Part I

Forrest B. Brown Diagnostics Applications Group (X-5) Los Alamos National Laboratory



• Time-dependent neutron transport with (prompt) fission source

$$\frac{1}{v} \frac{\partial \psi(\vec{r}, \textbf{E}, \vec{\Omega}, t)}{\partial t} = \left[-\vec{\Omega} \cdot \nabla - \Sigma_{\mathsf{T}}(\vec{r}, \textbf{E}) \right] \psi + \iint \psi(\vec{r}, \textbf{E}', \vec{\Omega}', t) \Sigma_{\mathsf{S}}(\vec{r}, \textbf{E}' \to \textbf{E}, \vec{\Omega} \cdot \vec{\Omega}') d\vec{\Omega}' d\textbf{E}' \right. \\ \left. + \frac{\chi(\textbf{E})}{4\pi} \iint v \Sigma_{\mathsf{F}}(\vec{r}, \textbf{E}') \psi(\vec{r}, \textbf{E}', \vec{\Omega}', t) d\vec{\Omega}' d\textbf{E}' \right. + \left. S(\vec{r}, \textbf{E}, \vec{\Omega}, t) \right.$$

This equation can be solved directly by Monte Carlo

- Simulate time-dependent transport for a neutron history
- If fission occurs, bank any secondary neutrons. When original particle is finished, simulate secondaries till done.
- Tallies for time bins, energy bins, cells, ...

Overall time-behavior $\psi(\mathbf{r}, \mathbf{E}, \Omega, \mathbf{t}) = \Psi(\mathbf{r}, \mathbf{E}, \Omega) e^{\alpha t}$ can be estimated by

$$\alpha \approx \frac{\ln W_2 - \ln W_1}{t_2 - t_1} \qquad \text{where} \quad W_j = \sum_{k=1}^{N_{particles}} wgt_k(t_j)$$

. .

Alpha Eigenvalue Equations



- For problems which are separable in space & time, it may be advantageous to solve a static eigenvalue problem, rather than a fully time-dependent problem
- If it is assumed that $\psi(\mathbf{r}, \mathbf{E}, \Omega, \mathbf{t}) = \Psi_{\alpha}(\mathbf{r}, \mathbf{E}, \Omega) \mathbf{e}^{\alpha \mathbf{t}}$, then substitution into the time-dependent transport equation yields

$$\begin{split} \left[\vec{\Omega} \cdot \nabla + \Sigma_{T}(\vec{r}, \mathsf{E}) + \frac{\alpha}{\nu} \right] \Psi_{\alpha}(\vec{r}, \mathsf{E}, \vec{\Omega}) &= \iint \Psi_{\alpha}(\vec{r}, \mathsf{E}', \vec{\Omega}') \Sigma_{S}(\vec{r}, \mathsf{E}' \to \mathsf{E}, \vec{\Omega} \cdot \vec{\Omega}') d\vec{\Omega}' d\mathsf{E}' \\ &+ \frac{\chi(\mathsf{E})}{4\pi} \iint \nu \Sigma_{\mathsf{F}}(\vec{r}, \mathsf{E}') \Psi_{\alpha}(\vec{r}, \mathsf{E}', \vec{\Omega}') d\vec{\Omega}' d\mathsf{E}' \end{split}$$

- This is a static equation, an eigenvalue problem for α and Ψ_α without time-dependence
- α is often called the time-eigenvalue or time-absorption
- α -eigenvalue problems can be solved by Monte Carlo methods

K_{eff} **Eigenvalue Equations**



- Another approach to creating a static eigenvalue problem from the time-dependent transport equation is to introduce K_{eff} , a scaling factor on the multiplication (v)
- Setting $\partial \psi / \partial t = 0$ and introducing the K_{eff} eigenvalue gives

$$\begin{split} \left[\vec{\Omega} \cdot \nabla + \Sigma_{T}(\vec{r},E) \right] \Psi_{k}(\vec{r},E,\vec{\Omega}) &= \iint \Psi_{k}(\vec{r},E',\vec{\Omega}') \Sigma_{S}(\vec{r},E' \to E,\vec{\Omega} \cdot \vec{\Omega}') d\vec{\Omega}' dE' \\ &+ \frac{1}{K_{eff}} \cdot \frac{\chi(E)}{4\pi} \iint v \Sigma_{F}(\vec{r},E') \Psi_{k}(\vec{r},E',\vec{\Omega}') d\vec{\Omega}' dE' \end{split}$$

- This is a static equation, an eigenvalue problem for $\mathsf{K}_{\mathsf{eff}}$ and Ψ_k without time-dependence
- K_{eff} is called the effective multiplication factor
- K_{eff} and Ψ_k should **never** be used to model time-dependent problems. [Use α and Ψ_{α} instead]
- K_{eff}-eigenvalue problems can be solved by Monte Carlo methods

Comments on K_{eff} and $\alpha~$ Equations



Criticality

Supercritical:	α > 0	or	$K_{eff} > 1$
Critical:	$\alpha = 0$	or	$K_{eff} = 1$
Subcritical:	α < 0	or	K _{eff} < 1

• K_{eff} vs. α eigenvalue equations

- $\Psi_{k}(\mathbf{r}, \mathbf{E}, \Omega) \neq \Psi_{\alpha}(\mathbf{r}, \mathbf{E}, \Omega)$, except for a critical system
- $-\alpha$ eigenvalue & eigenfunction used for time-dependent problems
- K_{eff} eigenvalue & eigenfunction used for reactor design & analysis
- Although $\alpha = (K_{eff}-1)/\Lambda$, where $\Lambda = \text{lifetime}$, there is **no** direct relationship between $\Psi_k(r, E, \Omega)$ and $\Psi_{\alpha}(r, E, \Omega)$
- K_{eff} eigenvalue problems can be simulated directly using Monte Carlo methods
- α eigenvalue problems are solved by Monte Carlo indirectly using a series of K_{eff} calculations



• Eigenvalue problems – reactor analysis & criticality safety

$$\begin{split} \Psi(p) &= \int \Psi(p') R(p' \to p) dp' + \frac{1}{K_{eff}} \int \Psi(p') F(p' \to p) dp' \\ \Psi &= R \bullet \Psi + \frac{1}{K_{eff}} F \bullet \Psi \end{split}$$

Iterative solution, using power iteration method

$$\begin{split} \Psi^{(i+1)} &= \mathbf{R} \bullet \Psi^{(i+1)} + \frac{1}{\mathsf{K}_{eff}^{(i)}} \mathbf{F} \bullet \Psi^{(i)} \\ \Psi^{(i+1)} &= \frac{1}{\mathsf{K}_{eff}^{(i)}} [\mathbf{I} - \mathbf{R}]^{-1} \, \mathbf{F} \bullet \Psi^{(i)} \qquad \qquad \mathsf{K}_{eff}^{i} = \int \mathbf{F} \bullet \Psi^{(i)} d\mathbf{p} d\mathbf{p}' \end{split}$$

- Monte Carlo approach:
 - Guess $\Psi^{(0)}$, $K_{eff}^{(0)}$
 - Follow a "batch" of histories, estimate $\Psi^{(i)}$, $K_{eff}^{(i)}$
 - Repeat until converged (discard tallies)
 - After converging, begin tallies, iterate until variances small enough



Random Walk for particle



Fixed-source Monte Carlo Calculation









Note: batch = cycle = iteration = generation

Initialize

- Assume a value for the initial K_{eff} (usually, $K_0 = 1$)
- Sample **M** fission sites from the initial source distribution

• For each cycle n, $n = 1 \dots N+D$

- Follow histories for all source particles in cycle
 - If fissions occur, bank the sites for use as source in next cycle
 - Make tallies for K_{cycle}⁽ⁿ⁾ using path, collision, & absorption estimators
 - If $n \le D$, discard any tallies
 - If n > D, accumulate tallies
- Estimate K_{cycle}⁽ⁿ⁾
- Compute final results & statistics using last **N** cycles



- Guess an initial source distribution
- Iterate until converged (How do you know ???)
- Then
 - For Sn code: done, print the results
 - For Monte Carlo: start tallies, keep running until uncertainties small enough
- Batch size? Convergence? Stationarity? Bias? Statistics?



• During a particle random walk,

wgt $\cdot \frac{\nu \Sigma_F}{\Sigma_T}$ = expected number of fission neutrons created at collision point

- Averaged over all collisions for all histories, the expected value for $~wgt\cdot\nu\Sigma_{\rm F}/~\Sigma_{\rm T}$ is $~K_{\rm eff}.$
- In order to bank approximately the same number of fission sites in each cycle, the current value of Keff is used to bias the selection of fission sites at a collision:

$$\begin{split} &\mathsf{R} = \mathsf{wgt} \cdot \frac{\mathsf{v}\Sigma_{\mathsf{F}}}{\Sigma_{\mathsf{T}}} \cdot \frac{1}{\mathsf{K}}, \qquad \mathsf{n} = \left\lfloor \mathsf{R} \right\rfloor \\ &\mathsf{If} \quad \xi < \mathsf{R} - \mathsf{n}, \qquad \mathsf{store} \quad \mathsf{n} + \mathsf{1} \quad \mathsf{sites} \ \mathsf{in} \ \mathsf{bank} \ \mathsf{with} \quad \mathsf{wgt'} = \mathsf{K} \\ &\mathsf{Otherwise}, \qquad \mathsf{store} \ \mathsf{n} \qquad \mathsf{sites} \ \mathsf{in} \ \mathsf{bank} \ \mathsf{with} \quad \mathsf{wgt'} = \mathsf{K} \end{split}$$



- N_J = number of particles starting cycle J,
 - N'_J = number of particles created by fission in cycle J (number of particles stored in fission bank)
 - The expected value for N'_J is: $E[N'_J] = K_{eff} \cdot N_J$
 - (N'_J/N_J) is a single-cycle estimator for K_{eff}
- To prevent the number of particles per cycle from growing exponentially (for K>1) or decreasing to 0 (for K<1), the particle population is renormalized at the end of each cycle:
 - In some Monte Carlo codes, the number of particles starting each cycle is a constant N. Russian roulette or splitting are used to sample N particles from the N' particles in the fission bank. (All particles in fission bank have a weight of 1.0)
 - In other codes, the total weight W starting each cycle is constant. The particle weights in the fission bank are renormalized so that the total weight is changed from W' to W. (Particles in fission bank have equal weights, but not necessarily 1.0)



• Pathlength estimator for Keff

$$\mathbf{K}_{\text{path}} = \left(\frac{\sum_{\substack{all \\ flights}} \mathbf{wgt}_{j} \cdot \mathbf{d}_{j} \cdot \mathbf{v} \Sigma_{F} \right) / \mathbf{W}$$

W = total weight starting each cycle

Collision estimator for Keff

$$K_{\text{collision}} = \left(\sum_{\substack{\text{all} \\ \text{collisions}}} wgt_j \cdot \frac{\nu \Sigma_F}{\Sigma_T} \right) / W$$

Absorption estimator for Keff

$$\mathbf{K}_{absorption} = \left(\sum_{\substack{all \\ absorptions}} \mathbf{wgt}_{j} \cdot \frac{\mathbf{v}\Sigma_{F}}{\Sigma_{A}} \right) / \mathbf{W}$$



- The Keff estimators from each cycle (K_{path}, K_{collision}, K_{absorption}) are used to compute the overall K_{path}, K_{collision}, & K_{absorption} for the problem & the standard deviations.
- The Keff estimators from each cycle (K_{path}, K_{collision}, K_{absorption}) can also be combined to produce a minimum-variance combined result, K_{combination}. This combination must account for correlations between the path, collision, & absorption estimators


- The renormalization procedure used at the end of each cycle introduces a small bias into the computed Keff
 - Renormalization involves multiplying particle weights by (W/W'), where W = total weight starting a cycle,

W'= total weight at the end of a cycle.

- W' is a random variable, due to fluctuations in particle random walks.
- Theoretical analysis of the MC iteration process & propagation of history fluctuations gives

bias in $K_{eff} = -\frac{\sigma_k^2}{K_{eff}} \cdot \begin{pmatrix} \text{sum of correlation coeff's} \\ \text{between batch K's} \end{pmatrix}$

- M = histories/cycle
- Bias in Keff ~ 1/M
 - Smaller M \Rightarrow larger cycle correlation \Rightarrow larger bias in Keff & source
 - Larger M \Rightarrow smaller cycle correlation \Rightarrow smaller bias

[T Ueki, "Intergenerational Correlation in Monte Carlo K-Eigenvalue Calculations", Nucl. Sci. Eng. (2002)]



• For a simple Godiva reactor calculation:



Keff vs 1/M



• Observed PDF for single-cycle Keff, for varying M



- Bias in Keff is negative: $K_{calc} < K_{true}$
- Bias is significant for M < 10 particles/cycle for $M \sim 100$ negligible for M > 10000 for M > 1000
- Recommendation:

Always use 1000 or more particles/cycle, preferably 5000, 10000, or more



• Some number of initial cycles must be discarded

- The source distribution & Keff are not known initially
- Guess at the source & Keff
- Iterate, discarding tallies
- When converged, iterate to accumulate tallies

• Number of iterations to discard depends on the dominance ratio

- Dominance Ratio = K_1 / K_{eff}
 - K_{eff} = eigenvalue of fundamental eigenmode
 - K_1 = eigenvalue of first higher eigenmode, $K_1 < K_{eff}$
- If DR close to 1 (e.g., .999...), 100s or 1000s of initial iterations may be required for initial source distribution errors to die away
- Most statistical tests for convergence are *ex post facto* tests to look for trends
- Most common practice is to examine plots of Keff vs. cycles



Plots of single-cycle Keff vs. cycle number • koode data from file runtpe ŝ, 1 à 1 \dot{s} k (trk longth) 1.2 1 ì 1 о è о 'n. 20 80 100 o 40 60 keff cycle number











• Keff is an integral quantity – converges faster than source shape





- Choose the number of cycles to discard by examining convergence plots
- Then, choose the total number of cycles to be large enough so that relative errors are "small enough"
 - Always run >25 cycles for tallies, to get good estimates of σ^2
 - Always try to run a few 100 or 1000 cycles for tallies
 - Statistical tests on convergence more reliable if more cycles
 - Better plots for assessing convergence

• Summary

- Particles per cycle > 1000
- Discarded cycles varies, check plots
- Tally cycles > 100



- Eigenvalue equation with both ${\rm K}_{\rm eff}$ & α
 - α is a fixed number, not a variable

$$\begin{bmatrix} \vec{\Omega} \cdot \nabla + \Sigma_{T}(\vec{r},E) + \max(\frac{\alpha}{\nu},0) \end{bmatrix} \Psi_{\alpha}(\vec{r},E,\vec{\Omega})$$

= $\max(\frac{-\alpha}{\nu},0)\Psi_{\alpha}(\vec{r},E,\vec{\Omega}) + \iint \Psi_{\alpha}(\vec{r},E',\vec{\Omega}')\Sigma_{S}(\vec{r},E' \to E,\vec{\Omega} \cdot \vec{\Omega}')d\vec{\Omega}'dE'$
+ $\frac{1}{K_{eff}} \cdot \frac{\chi(E)}{4\pi} \iint v\Sigma_{F}(\vec{r},E')\Psi_{\alpha}(\vec{r},E',\vec{\Omega}')d\vec{\Omega}'dE'$

- Note on the max(α/v , 0) and max($-\alpha/v$, 0) terms
 - If $\alpha < 0$, real absorption plus time absorption could be negative
 - If $\alpha < 0$, move α/v to right side to prevent negative absorption,
 - If $\alpha < 0$, $-\alpha/v$ term on right side is treated as a delta-function source
- Select a fixed value for α
- Solve the K-eigenvalue equations, with fixed time-absorption α/v
- Select a different α and solve for a new Keff
- Repeat, searching for value of α which results in Keff = 1

Special Topic – Stationarity Tests



Plots of single-cycle Keff or cumulative Keff are difficult to interpret when assessing convergence





The MCNP team has been investigating new stationarity tests

Progressive relative entropy



Special Topic – Stationarity Tests







One cycle delay embedding plot of relative entropy wrt initial source



Special Topic – Stationarity Tests



• In a series of related papers, we have significantly extended the theory of Monte Carlo eigenvalue calculations, explicitly accounting for correlation effects.

LA-UR-02-0190:	T Ueki, "Intergenerational Correlation in Monte Carlo K-Eigenvalue Calculations", Nucl. Sci. Eng. (2002)
LA-UR-01-6770:	T Ueki & FB Brown, "Autoregressive Fitting for Monte Carlo K-effective Confidence Intervals", ANS Summer Meeting, (June 2002)
LA-UR-02-3783:	T Ueki & FB Brown, "Stationarity Diagnostics Using Shannon Entropy in Monte Carlo Criticality Calculations I: F Test", ANS Winter Meeting (Nov 2002)
LA-UR-02-6228:	T Ueki & FB Brown, "Stationarity and Source Convergence in Monte Carlo Criticality Calculations", ANS Topical Meeting on Mathematics & Computation, Gatlinburg, TN (April, 2003)
LA-UR-03-0106:	T Ueki, FB Brown, DK Parsons, " Dominance Ratio Computation via Time Series Analysis of Monte Carlo Fission Sources" , ANS Annual Meeting (June 2003)
LA-UR-02-5700:	T Ueki, FB Brown, DK Parsons, & DE Kornreich, "Autocorrelation and Dominance Ratio in Monte Carlo Criticality Calculations", Nucl. Sci. Eng. (Nov 2003)
LA-UR-03–3949:	T Ueki & FB Brown, "Informatics Approach to Stationarity Diagnostics of the Monte Carlo Fission Source Distribution", ANS Winter meeting (Nov 2003)
LA-UR-03-5823:	T Ueki, FB Brown, DK Parsons, JS Warsa, "Time Series Analysis of Monte Carlo Fission Source: I. Dominance Ratio Calculation", Nucl. Sci. Eng. (Nov 2004)
LA-UR-03-????:	T Ueki & FB Brown, "Stationarity Modeling and Informatics-Based Diagnostics in Monte Carlo Criticality Calculations," submitted to Nucl. Sci. Eng.





Fundamentals of Monte Carlo Particle Transport



Lecture 8

Eigenvalue Calculations Part II

Forrest B. Brown Diagnostics Applications Group (X-5) Los Alamos National Laboratory

Eigenvalue Calculations – Part II



- K-eigenvalue equation
- Solution by power iteration
- Convergence of power iteration
- Stationarity Diagnostics
- Weilandt acceleration method
- Superhistory method



$$\begin{split} \left[\vec{\Omega} \cdot \nabla + \Sigma_{T}(\vec{r},E) \right] \Psi_{k}(\vec{r},E,\vec{\Omega}) &= \iint \Psi_{k}(\vec{r},E',\vec{\Omega}') \Sigma_{S}(\vec{r},E' \to E,\vec{\Omega} \cdot \vec{\Omega}') d\vec{\Omega}' dE' \\ &+ \frac{1}{K_{eff}} \cdot \frac{\chi(E)}{4\pi} \iint v \Sigma_{F}(\vec{r},E') \Psi_{k}(\vec{r},E',\vec{\Omega}') d\vec{\Omega}' dE' \end{split}$$

where

 K_{eff} = k-effective, eigenvalue for fundamental mode $\Psi_{k}(\vec{r}, E, \vec{\Omega})$ = angular flux, for fundamental k-eigenmode

 $\vec{\Omega} \cdot \nabla \Psi_{k}(\vec{r}, E, \vec{\Omega}) = \text{loss term, leakage}$ $\Sigma_{\tau}(\vec{r}, E) \Psi_{k}(\vec{r}, E, \vec{\Omega}) = \text{loss term, collisions}$

$$\begin{split} &\iint \Psi_{k}(\vec{r},E',\vec{\Omega}')\Sigma_{S}(\vec{r},E'\to E,\vec{\Omega}\cdot\vec{\Omega}')d\vec{\Omega}'dE' \\ &\frac{1}{K_{eff}}\cdot\frac{\chi(E)}{4\pi} \iint \nu\Sigma_{F}(\vec{r},E')\Psi_{k}(\vec{r},E',\vec{\Omega}')d\vec{\Omega}'dE' \end{split}$$

- = gain term, scatter from E', Ω' into E, Ω
- = gain term, production from fission

 \Rightarrow Jointly find K_{eff} and $\Psi_{k}(\mathbf{r},\mathbf{E},\Omega)$ such that equation balances



• Use operator (or matrix) form to simplify notation

$$(L+T)\Psi = S\Psi + \frac{1}{K_{eff}}M\Psi$$

where

S = scatter-in operator M = fission multiplication operator

• Rearrange

$$\begin{split} (\mathsf{L} + \mathsf{T} - \mathsf{S}) \Psi &= \frac{1}{\mathsf{K}_{\text{eff}}} \mathsf{M} \Psi \\ \Psi &= \frac{1}{\mathsf{K}_{\text{eff}}} \cdot (\mathsf{L} + \mathsf{T} - \mathsf{S})^{-1} \mathsf{M} \Psi \\ \Psi &= \frac{1}{\mathsf{K}_{\text{eff}}} \cdot \mathsf{F} \Psi \end{split}$$

 \Rightarrow This eigenvalue equation will be solved by power iteration

Eigenvalue equation

$$\Psi = \frac{1}{\mathsf{K}_{eff}} \cdot \mathsf{F} \Psi$$

1. Assume that k_{eff} and Ψ on the right side are known for iteration n, solve for Ψ on left side (for iteration n+1)

$$\Psi^{(n+1)} = \frac{1}{K_{\text{eff}}^{(n)}} \cdot F \Psi^{(n)}$$

Note: This requires solving the equation below for $\Psi^{(n+1)},$ with $K_{eff}{}^{(n)}$ and $\Psi^{(n)}$ fixed

$$(\mathsf{L} + \mathsf{T} - \mathsf{S})\Psi^{(n+1)} = \frac{1}{\mathsf{K}_{eff}^{(n)}}\mathsf{M}\Psi^{(n)}$$

2. Then, compute $K_{eff}^{(n+1)}$

$$K_{\text{eff}}^{(n+1)} = K_{\text{eff}}^{(n)} \cdot \frac{\int M \Psi^{(n+1)} d\vec{r}}{\int M \Psi^{(n)} d\vec{r}}$$

(other norms could be used)

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4. Repeat 1–3 until both $K_{eff}^{(n+1)}$ and $\Psi^{(n+1)}$ have converged



• Power iteration for Monte Carlo k-effective calculation



Power Iteration



Diffusion Theory or Discrete-ordinates Transport

1. Initial guess for K_{eff} and Ψ

 $K_{eff}^{(0)}, \Psi^{(0)}$ 2. Solve for $\Psi^{(n+1)}$

Inner iterations over space or space/angle to solve for $\Psi^{(n+1)}$

$$(L + T - S)\Psi^{(n+1)} = \frac{1}{\kappa_{eff}^{(n)}}M\Psi^{(n)}$$

3. Compute new K_{eff}

$$K_{\text{eff}}^{(n+1)} = K_{\text{eff}}^{(n)} \cdot \frac{1 \text{-} M \Psi^{(n+1)}}{1 \text{-} M \Psi^{(n)}}$$

4. Repeat 1–3 until both $K_{eff}^{(n+1)}$ and $\Psi^{(n+1)}$ have converged

 Initial guess for K_{eff} and Ψ K_{eff}⁽⁰⁾, Ψ⁽⁰⁾
 Solve for Ψ⁽ⁿ⁺¹⁾ Follow particle histories to solve for Ψ⁽ⁿ⁺¹⁾

$$(L+T-S)\Psi^{(n+1)} = \tfrac{1}{K_{\text{eff}}^{(n)}}M\Psi^{(n)}$$

Monte Carlo

During histories, save fission sites to use for source in next iteration

3. Compute new K_{eff}

During histories for iteration (n+1), estimate $K_{eff}^{(n+1)}$

$$K_{\text{eff}}^{(n+1)} = K_{\text{eff}}^{(n)} \cdot \frac{\int M \Psi^{(n+1)} d\vec{r}}{\int M \Psi^{(n)} d\vec{r}}$$

- 4. Repeat 1–3 until both $K_{eff}^{(n+1)}$ and $\Psi^{(n+1)}$ have converged
- 5. Continue iterating, to compute tallies





- Guess an initial source distribution
- Iterate until converged

(How do you know ???)

- Then
 - For S_n code: done, print the results
 - For Monte Carlo: start tallies, keep running until uncertainties small enough
- Convergence? Stationarity? Bias? Statistics?



• Expand Ψ in terms of eigenfunctions $u_j(r, E, \Omega)$

$$\Psi = \sum_{j=0}^{\infty} a_j \vec{u}_j = a_0 \vec{u}_0 + a_1 \vec{u}_1 + a_2 \vec{u}_2 + a_3 \vec{u}_3 + \dots$$

$$\int \vec{u}_{j}\vec{u}_{k}dV=\delta_{jk}$$

$$a_{j}=\int\Psi\cdot\vec{u}_{j}dV$$

$$\vec{u}_{j} = \frac{1}{k_{j}} F \cdot \vec{u}_{j} \qquad k_{0} > k_{1} > k_{2} > \dots$$
$$k_{0} \equiv k_{effective}$$



• Expand the initial guess in terms of the eigenmodes

$$\Psi^{(0)} = \sum_{j=0} a_j^{(0)} \vec{u}_j$$

• Substitute the expansion for Ψ into eigenvalue equation

$$\Psi^{(n+1)} = \frac{1}{K^{(n)}} F \cdot \Psi^{(n)} = \frac{1}{k^{(n)}} \cdot \frac{1}{k^{(n-1)}} \dots \frac{1}{k^{(0)}} \cdot F^n \cdot \Psi^{(0)}$$

$$= \left[\prod_{m=0}^{n} \frac{k_{0}}{K^{(m)}}\right] \cdot a_{0}^{(0)} \cdot \left[\vec{u}_{0} + \sum_{j=1}^{n} \left(\frac{a_{j}^{(0)}}{a_{0}^{(0)}}\right) \cdot \left(\frac{k_{j}}{k_{0}}\right)^{n+1} \cdot \vec{u}_{j}\right]$$

$$\approx \left[\text{constant}\right] \cdot \left[\vec{u}_0 + \left(\frac{a_1^{(0)}}{a_0^{(0)}}\right) \cdot \left(\frac{k_1}{k_0}\right)^{n+1} \cdot \vec{u}_1 + \left(\frac{a_2^{(0)}}{a_0^{(0)}}\right) \cdot \left(\frac{k_2}{k_0}\right)^{n+1} \cdot \vec{u}_2 + \dots\right]$$



$$\mathsf{K}^{(n+1)} \approx \mathsf{k}_{0} \cdot \frac{\left[1 + \left(\frac{a_{1}^{(0)}}{a_{0}^{(0)}}\right) \cdot \left(\frac{\mathsf{k}_{1}}{\mathsf{k}_{0}}\right)^{n+1} \cdot \mathsf{G}_{1} + \left(\frac{a_{2}^{(0)}}{a_{0}^{(0)}}\right) \cdot \left(\frac{\mathsf{k}_{2}}{\mathsf{k}_{0}}\right)^{n+1} \cdot \mathsf{G}_{2} + \dots\right]}{\left[1 + \left(\frac{a_{1}^{(0)}}{a_{0}^{(0)}}\right) \cdot \left(\frac{\mathsf{k}_{1}}{\mathsf{k}_{0}}\right)^{n} \cdot \mathsf{G}_{1} + \left(\frac{a_{2}^{(0)}}{a_{0}^{(0)}}\right) \cdot \left(\frac{\mathsf{k}_{2}}{\mathsf{k}_{0}}\right)^{n} \cdot \mathsf{G}_{2} + \dots\right]}\right]$$

where
$$G_m = \frac{\int M \vec{u}_m d\vec{r}}{\int M \vec{u}_0 d\vec{r}}$$



- After n iterations, the J-th mode error component is reduced by the factor $(k_J/k_0)^n$
- Since $1 > k_1/k_0 > k_2/k_0 > k_3/k_0 > ...,$ after the initial transient, error in $\Psi^{(n)}$ is dominated by first mode:

$$\Psi^{(n)} \approx [\text{constant}] \cdot \left[\vec{u}_0 + \left(\frac{a_1^{(0)}}{a_0^{(0)}} \right) \cdot \left(\frac{k_1}{k_0} \right)^n \cdot \vec{u}_1 + \dots \right]$$

- (k_1 / k_0) is called the dominance ratio, DR or ρ
 - Errors die off as $\sim (DR)^n$
 - − To reduce 10% error \rightarrow .1% error
 - DR~.9 \rightarrow 44 iterations DR~.99 \rightarrow 458 iterations DR~.999 \rightarrow 2301 iterations



Typical K-effective convergence patterns

- Higher mode error terms die out as $(k_J / k_0)^n$, for n iterations
- When initial guess is concentrated in center of reactor, initial K_{eff} is too high (underestimates leakage)
- When initial guess is uniformly distributed, initial K_{eff} is too low (overestimates leakage)
- The Sandwich Method uses 2 K_{eff} calculations one starting too high & one starting too low.
 Both calculations should converge to the same result.









Power Iteration – Convergence



$$\Psi^{(n+1)} \approx [\text{constant}] \cdot \left[\vec{u}_0 + \left(\frac{a_1^{(0)}}{a_0^{(0)}} \right) \cdot \left(\frac{k_1}{k_0} \right)^{n+1} \cdot \vec{u}_1 + \dots \right]$$

$$\begin{split} \mathsf{K}^{(n+1)} \approx \mathsf{k}_{0} \cdot \frac{\left[1 + \left(\frac{a_{1}^{(0)}}{a_{0}^{(0)}}\right) \cdot \left(\frac{\mathsf{k}_{1}}{\mathsf{k}_{0}}\right)^{n+1} \cdot \mathsf{G}_{1} + \ldots\right]}{\left[1 + \left(\frac{a_{1}^{(0)}}{a_{0}^{(0)}}\right) \cdot \left(\frac{\mathsf{k}_{1}}{\mathsf{k}_{0}}\right)^{n} \cdot \mathsf{G}_{1} + \ldots\right]} \approx \mathsf{k}_{0} \cdot \left[1 + \left(\frac{a_{1}^{(0)}}{a_{0}^{(0)}}\right) \cdot \left(\frac{\mathsf{k}_{1}}{\mathsf{k}_{0}}\right)^{n+1} \cdot \mathsf{G}_{1}\right] \cdot \left[1 - \left(\frac{a_{1}^{(0)}}{a_{0}^{(0)}}\right) \cdot \left(\frac{\mathsf{k}_{1}}{\mathsf{k}_{0}}\right)^{n} \cdot \mathsf{G}_{1}\right] \right] \cdot \left[1 + \left(\frac{a_{1}^{(0)}}{a_{0}^{(0)}}\right) \cdot \left(\frac{\mathsf{k}_{1}}{\mathsf{k}_{0}}\right)^{n} \cdot \mathsf{G}_{1}\right] \cdot \left[1 + \left(\frac{\mathsf{k}_{1}}{\mathsf{k}_{0}}\right] \cdot \mathsf{K}_{1}\right] \cdot \left[1 + \left(\frac{\mathsf{k}_{1}}{\mathsf{k}_{0}}\right] \cdot \mathsf{K}_{1}\right] \cdot \mathsf{K}_{1}\right] \cdot \mathsf{K}_{1} \cdot \mathsf{K}_{1}\right] \cdot \mathsf{K}_{1} \cdot \mathsf{K}_{1}\right] \cdot \mathsf{K}_{1} \cdot \mathsf{K}_{1} \cdot \mathsf{K}_{1} \cdot \mathsf{K}_{1}\right] \cdot \mathsf{K}_{1} \cdot \mathsf{K}_{1} \cdot \mathsf{K}_{1} \cdot \mathsf{K}_{1} \cdot \mathsf{K}_{1}\right] \cdot \mathsf{K}_{1} \cdot$$

$$\approx \mathbf{k}_{0} \cdot \left[\mathbf{1} + \left(\frac{\mathbf{a}_{1}^{(0)}}{\mathbf{a}_{0}^{(0)}} \right) \cdot \left(\frac{\mathbf{k}_{1}}{\mathbf{k}_{0}} \right)^{n} \cdot \left(\frac{\mathbf{k}_{1}}{\mathbf{k}_{0}} - \mathbf{1} \right) \cdot \mathbf{G}_{1} + \ldots \right]$$

- For problems with a high dominance ratio (e.g., DR ~ .99), the error in K_{eff} may be small, since the factor $(k_1/k_0 1)$ is small.
 - ⇒ K_{eff} may appear converged, even if the source distribution is <u>not</u> converged



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Power Iteration – Convergence





- For Monte Carlo power iteration, statistical fluctuations in source shape die out gradually over a number of successive iterations.
 - Persistence of the noise over successive iterations gives correlation among source distributions in successive iterations. (<u>Positive</u> correlation)
 - Correlation directly affects confidence intervals:
 Serial correlation in the source distribution → larger confidence intervals
- ⇒ Most Monte Carlo codes ignore these correlation effects & incorrectly <u>underestimate</u> the confidence intervals



Summary

- Local errors in the source distribution decay as $(k_J/k_0)^n$
 - Higher eigenmodes die out rapidly, convergence dominated by k_1/k_0
 - High DR \rightarrow slow convergence
 - High DR \rightarrow large correlation \rightarrow large error in computed variances
- Errors in K_{eff} decay as $(k_J/k_0 1) * (k_J/k_0)^n$
 - High DR $\rightarrow k_J/k_0 \sim 1 \rightarrow \text{small error}$
- K_{eff} errors die out faster than local source errors
 - K_{eff} is an integral quantity positive & negative fluctuations cancel
- High DR is common for
 - Large reactors, with small leakage
 - Heavy-water moderated or reflected reactors
 - Loosely-coupled systems
- \Rightarrow If <u>local</u> tallies are important (e.g., assembly power, pin power, ...), examine their convergence – not just K_{eff} convergence

Keff Calculations – Stationarity Diagnostics



Plots of single-cycle Keff or cumulative Keff are sometimes difficult to interpret when assessing convergence


Keff Calculations – Stationarity Diagnostics



- Initial cycles of a Monte Carlo K-effective calculation should be discarded, to avoid contaminating results with errors from initial guess
 - How many cycles should be discarded?
 - How do you know if you discarded enough cycles?



- Analysis of the power iteration method shows that Keff is not a reliable indicator of convergence — K_{eff} can converge faster than the source shape
- Based on concepts from information theory, Shannon entropy of the source distribution is useful for characterizing the convergence of the source distribution

K_{eff} **Calculations – Stationarity Diagnostics**



- Divide the fissionable regions of the problem into N_s spatial bins
 - Spatial bins should be consistent with problem symmetry
 - Typical choices: -1 bin for each assembly

regular grid superimposed on core



- Rule-of-thumb for number of spatial bins:

N_s ~ (histories/batch) / 25 or less

Why?

- Would like to have >25 fission source sites per bin to get good statistics
- If source distribution were uniform, ~25 sites would be in each bin

Shannon entropy of the source distribution

$$H(S) = -\sum_{J=1}^{N_S} p_J \cdot \ln_2(p_J), \text{ where } p_J = \frac{(\text{# source particles in bin J})}{(\text{total # source particles in all bins})}$$



Shannon entropy of the source distribution

$$H(S) = -\sum_{J=1}^{N_S} p_J \cdot ln_2(p_J), \text{ where } p_J = \frac{(\text{# source particles in bin J})}{(\text{total # source particles in all bins})}$$

- $0 \leq H(S) \leq \ln_2(N_S)$
- Note that $p_J \ln_2(p_J) = 0$ if $p_J=0$
- For a uniform source distribution, $p_1 = p_2 = ... = p_{Ns} = 1/N_s$, so that $H(S) = In_2(N_s)$
- For a point source (in a single bin), H(S) = 0
- H(S⁽ⁿ⁾) provides a single number to characterize the source distribution for iteration n
 - As the source distribution converges in 3D space,
 a line plot of H(S⁽ⁿ⁾) vs. n (the iteration number) converges

K_{eff} Calculations – Stationarity Diagnostics





Figure 3: Posterior computation of relative entropy assuming the true source is the mean source over 1501-2500 cycles (problem 1)









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• Example – Reactor core (Problem inp24)





• Example – Loosely-coupled array of spheres (Problem test4s)





K_{eff} Calculations – Stationarity Diagnostics



• Example – Fuel Storage Vault (Problem OECD_bench1)





• Basic transport equation for eigenvalue problems

$$(L + T - S)\Psi = \frac{1}{K_{eff}}M\Psi$$

L = loss to leakageT = loss to collisions S = gain from scatter-in M = gain from fission multiplication

- Define a fixed parameter \mathbf{k}_{e} such that $\mathbf{k}_{e} > \mathbf{k}_{0}$ (\mathbf{k}_{0} = exact eigenvalue)
- Subtract $\frac{1}{k_0}M\Psi$ from each side of the transport equation

$$(L + T - S - \frac{1}{k_e}M)\Psi = (\frac{1}{K_{eff}} - \frac{1}{k_e})M\Psi$$

• Solve the modified transport equation by power iteration

$$(\mathsf{L} + \mathsf{T} - \mathsf{S} - \frac{1}{\mathsf{k}_{\mathsf{e}}}\mathsf{M})\Psi^{(\mathsf{n}+1)} = (\frac{1}{\mathsf{K}_{\mathsf{eff}}^{(\mathsf{n})}} - \frac{1}{\mathsf{k}_{\mathsf{e}}})\mathsf{M}\Psi^{(\mathsf{n})}$$



• Power iteration for modified transport equation

$$\begin{split} (\mathsf{L} + \mathsf{T} - \mathsf{S} - \frac{1}{k_{e}}\mathsf{M})\Psi^{(n+1)} &= (\frac{1}{\mathsf{K}_{eff}^{(n)}} - \frac{1}{k_{e}})\mathsf{M}\Psi^{(n)} \\ \Psi^{(n+1)} &= (\frac{1}{\mathsf{K}_{eff}^{(n)}} - \frac{1}{k_{e}}) \cdot (\mathsf{L} + \mathsf{T} - \mathsf{S} - \frac{1}{k_{e}}\mathsf{M})^{-1}\mathsf{M}\Psi^{(n)} \\ \Psi^{(n+1)} &= \frac{1}{\tilde{\mathsf{K}}^{(n)}} \cdot \tilde{\mathsf{F}}\Psi^{(n)} \\ & \text{where} \quad \tilde{\mathsf{K}}^{(n)} &= (\frac{1}{\mathsf{K}_{eff}^{(n)}} - \frac{1}{k_{e}})^{-1} \quad \text{or} \quad \mathsf{K}_{eff}^{(n)} &= (\frac{1}{\tilde{\mathsf{K}}^{(n)}} + \frac{1}{k_{e}})^{-1} \end{split}$$

- How to choose k_e
 - $\mathbf{k}_{\mathbf{e}}$ must be larger than $\mathbf{k}_{\mathbf{0}}$ (but, don't know \mathbf{k}_{0} !)
 - k_e must be held constant for all of the histories in a batch, but can be adjusted between batches
 - Typically, guess a large initial value for k_e , such as $k_e=5$ or $k_e=2$
 - Run a few batches, keeping k_e fixed, to get an initial estimate of K_{eff}
 - Adjust k_e to a value slightly larger than the estimated K_{eff}
 - Run more batches, possibly adjusting k_e if the estimated K_{eff} changes

Wielandt Method

Convergence

- Eigenfunctions for the Wielandt method are same as for basic power iteration

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- Eigenvalues are shifted:

$$\tilde{\mathbf{k}}_{\mathrm{J}} = \left[\frac{1}{k_{\mathrm{J}}} - \frac{1}{k_{\mathrm{e}}}\right]^{-1} \qquad \mathbf{k}_{\mathrm{e}} > \mathbf{k}_{\mathrm{0}} > \mathbf{k}_{\mathrm{1}} > \dots$$

- Expand the initial guess, substitute into Wielandt method, rearrange to:

$$\begin{split} \Psi^{(n+1)} &\approx [\text{constant}] \cdot \left[\vec{u}_0 + \left(\frac{a_1^{(0)}}{a_0^{(0)}} \right) \cdot \left(\frac{k_e - k_0}{k_e - k_1} \cdot \frac{k_1}{k_0} \right)^{n+1} \cdot \vec{u}_1 + \ldots \right] \\ K^{(n+1)} &\approx k_0 \cdot \left[1 + \left(\frac{a_1^{(0)}}{a_0^{(0)}} \right) \cdot \left(\frac{k_e - k_0}{k_e - k_1} \cdot \frac{k_1}{k_0} \right)^n \cdot \left(\frac{k_e - k_0}{k_e - k_1} \cdot \frac{k_1}{k_0} - 1 \right) \cdot G_1 + \ldots \right] \end{split}$$

- Additional factor $(k_e - k_0)/(k_e - k_1)$ is less than 1 and positive, so that the red terms die out faster than for standard power iteration



• The **dominance ratio** for this modified power iteration is

$$\mathsf{DR'} = \frac{\tilde{\mathsf{k}}_1}{\tilde{\mathsf{k}}_0} = \frac{\left[\frac{1}{\mathsf{k}_1} - \frac{1}{\mathsf{k}_e}\right]^{-1}}{\left[\frac{1}{\mathsf{k}_0} - \frac{1}{\mathsf{k}_e}\right]^{-1}} = \frac{\mathsf{k}_e - \mathsf{k}_0}{\mathsf{k}_e - \mathsf{k}_1} \cdot \frac{\mathsf{k}_1}{\mathsf{k}_0} = \frac{\mathsf{k}_e - \mathsf{k}_0}{\mathsf{k}_e - \mathsf{k}_1} \cdot \mathsf{DR}$$

- Since $k_e > k_0$ and $k_0 > k_1$, **DR' < DR**

- DR of Wielandt method is always **smaller** than standard power iteration
- Wielandt acceleration improves the convergence rate of the power iteration method for solving the k-eigenvalue equation



Iteration, n

\Rightarrow Weilandt method converges at a faster rate than power iteration



Monte Carlo procedure for Wielandt acceleration

$$(\mathsf{L} + \mathsf{T} - \mathsf{S} - \frac{1}{k_e}\mathsf{M})\Psi^{(n+1)} = (\frac{1}{\mathsf{K}_{eff}^{(n)}} - \frac{1}{k_e})\mathsf{M}\Psi^{(n)}$$

- For standard Monte Carlo (power iteration) in generation n+1
 - When a collision occurs, the expected number of fission neutrons produced is

$$\mathbf{n}_{\mathsf{F}} = \left[\mathsf{wgt} \cdot \frac{\mathbf{v} \Sigma_{\mathsf{F}}}{\Sigma_{\mathsf{T}}} \cdot \frac{1}{\mathsf{K}^{(\mathsf{n})}} + \xi \right]$$

- Store n_F copies of particle in the "fission bank"
- Use the fission bank as the source for the next generation (n+2)
- For Monte Carlo Wielandt method in generation n+1
 - When a collision occurs, compute 2 expected numbers of fission neutrons

$$n_{\mathsf{F}}' = \left\lfloor \mathsf{wgt} \cdot \frac{\mathsf{v}\Sigma_{\mathsf{F}}}{\Sigma_{\mathsf{T}}} \cdot \left(\frac{1}{\mathsf{K}^{(n)}} - \frac{1}{\mathsf{k}_{\mathsf{e}}}\right) + \xi \right\rfloor \qquad \qquad n_{\mathsf{e}}' = \left\lfloor \mathsf{wgt} \cdot \frac{\mathsf{v}\Sigma_{\mathsf{F}}}{\Sigma_{\mathsf{T}}} \cdot \frac{1}{\mathsf{k}_{\mathsf{e}}} + \xi \right\rfloor$$

- Note that $E[n'_F + n'_e] = E[n_F]$
- Store n'_F copies of particle in the "fission bank"
- Follow n'_e copies of the particle in the current generation (n+1)
- Use the fission bank as the source for the next generation (n+2)



• <u>Power iteration</u> for Monte Carlo k-effective calculation





• <u>Wielandt method</u> for Monte Carlo k-effective calculation



within generation due to Wielandt method



Summary

- <u>Wielandt Method has a lower DR than power iteration</u>
 - Faster convergence <u>rate</u> than power iteration \Rightarrow fewer iterations
 - Some of the particle random walks are moved from the next generation into the current generation ⇒ more work per iteration
 - Same total number of random walks \Rightarrow no reduction in CPU time
- Advantages
 - Reduced chance of false convergence for very slowly converging problems
 - Reduced inter-generation correlation effects on variance
 - Fission source distribution spreads more widely in a generation (due to the additional particle random walks), which should result in more interactions for loosely-coupled problems



• Standard generation model, solved by power iteration

$$\Psi^{(n+1)} = \frac{1}{K_{eff}^{(n)}} \cdot F \Psi^{(n)}$$

- Superhistory method
 - Follow several generations (L) before recomputing K_{eff} and renormalizing

$$\Psi^{(n+1)} = \frac{1}{\tilde{K}^{(n)}} \cdot \tilde{F} \Psi^{(n)}, \quad \text{ with } \tilde{F} = F^{L}, \quad \tilde{K}^{(n)} = (K_{\text{eff}}^{(n)})^{L}$$

Convergence

- Same eigenfunctions as standard power iteration
- Eigenvalues are k_0^L , k_1^L , k_2^L , ...
- DR' = DR^L, where DR = dominance ratio for power iteration
- Fewer iterations, but L generations per iteration \Rightarrow same work as power iteration
- Same convergence rate as power iteration

Advantages

- Reduced correlation between iterations
- Fewer renormalizations



Superhistory Method for Monte Carlo k-effective calculation



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Fundamentals of Monte Carlo Particle Transport





Variance Reduction

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Monte Carlo Calculations





mcnp, rcp, vim, racer, sam-ce, tart, morse, keno, tripoli, mcbend, monk, o5r, recap, andy,.....

- Variance reduction
 - Modify the PDFs for physics interactions to favor events of interest
 - Use splitting/rouletting to increase particles in certain geometric regions
 - Kill particles in uninteresting parts of problem
- May be necessary in order to sample rare events
- More samples (with less weight each) —> smaller variance in tallies



Given a function R(x), where x is a random variable with PDF f(x),

Expected value of R(x) is

$$\mu = \int \mathsf{R}(\mathsf{x}) \,\mathsf{f}(\mathsf{x}) \,\mathsf{d}\mathsf{x}$$

- Variance of R(x) is

$$\sigma^2 = \int R^2(x) f(x) \, dx - \mu^2$$

Monte Carlo method for estimating μ

- make N random samples $\hat{\boldsymbol{x}}_{j}$ from f(x)
- Then

$$\begin{split} \overline{R} &\approx \frac{1}{N} \sum_{j=1}^{N} R(\hat{x}_{j}) \\ \sigma_{\overline{R}}^{2} &\approx \frac{1}{N-1} \cdot \left(\frac{1}{N} \sum_{j=1}^{N} R^{2}(\hat{x}_{j}) - \overline{R}^{2} \right) \end{split}$$



•Expected mean score is not changed by variance reduction



•Variance is changed due to altered sampling scheme

$$\sigma^{2} = \int \left[\mathsf{R}(x) \right]^{2} \mathsf{f}(x) \, dx - \mu^{2} \qquad \qquad \sigma^{2} = \int \left[\mathsf{R}(x) \frac{\mathsf{f}(x)}{\mathsf{g}(x)} \right]^{2} \, \mathsf{g}(x) \, dx - \mu^{2}$$

Goal: Choose g(x) such that variance is reduced

Review



• Given a set of random samples, $x_1, x_2, ..., x_{N_1}$

– Mean

$$\overline{\mathbf{x}} = \frac{1}{N} \sum_{j=1}^{N} \mathbf{x}_{j}$$

- Variance of the mean

$$\sigma_{\overline{x}}^{2} = \frac{1}{N-1} \cdot \left(\frac{1}{N} \sum_{j=1}^{N} x_{j}^{2} - \overline{x}^{2} \right)$$

- Relative Error

$$\mathsf{RE} = \frac{\sigma_{\overline{x}}}{\overline{x}}$$

- Figure of Merit

$$FOM = \frac{1}{RE^2 \cdot T}$$

• Variance reduction: Reduce RE or T, to increase FOM



Analog Monte Carlo

- Faithful simulation of particle histories
- No alteration of PDFs (i.e., no biasing or variance reduction)
- Particle is born with weight = 1.0
- Weight unchanged throughout history until particle is killed
- Scores are weighted by 1.0 when tallying events of interest

Weighted Monte Carlo (non-analog)

- Alter the PDFs to favor events of interest
- Particle is born with weight = 1.0
- Weight, wgt, is altered if biased PDF is used
- Weight can also be changed by Russian roulette/splitting & other variance reduction techniques
- Scores are weighted by wgt when tallying events of interest

Variance Reduction – General Approaches



Truncation

Remove particles from parts of phase space that do not contribute significantly to the tallies

Population control

 Use particle splitting and Russian rouletting to control the number of samples taken in various regions of phase space

Modified sampling

- Modify the PDFs representing problem physics, to favor tallies of interest

Deterministic methods

 Replace portions of a particle random walk by the expected results obtained from a deterministic calculation



• MCNP has 14 variance reduction techniques

- 1. Time and energy cutoffs
- 2. Geometry splitting & roulette
- 3. Weight windows
- 4. Exponential transform
- 5. Forced collisions
- 6. Energy splitting & roulette
- 7. Time splitting & roulette
- 8. Point and ring detectors
- 9. DXTRAN
- 10. Implicit capture
- 11. Weight cutoff
- 12. General source biasing
- 13. Secondary particle biasing
- 14. Bremsstrahlung energy biasing



- Also called implicit absorption or non-absorption weighting
- Modify collision process according to expected outcome
- Particle always survives collision
 - Tally expected absorption,
 - Reduce weight of surviving particle,

wgt • (σ_A / σ_T) wgt' = wgt • $(1 - \sigma_A / \sigma_T)$

- Extends particle history so that more particles reach events which occur after many collisions
- Most effective for thermal reactor problems, but doesn't hurt in other types of problems
- Must also use some form of low-weight cutoff to eliminate particles with very low weight

Geometry Splitting & Russian Roulette



- Increase the number of particles in "important" regions, decrease the number of particles in "unimportant" regions
- Assign each cell an importance, I_{cell}
 - Arbitrary, use best guess or adjoint fluxes from deterministic calculation
 - Could use one value for all energies or separate values for different energy ranges
 - Higher value —> more important
 - $-I_{cell} > 0$
 - I_{cell}=0 is a way to declare regions as not in physical problem
 - Values of I_{cell} must not change during Monte Carlo calculation
- Modify random walk simulation at surface crossings:
 - If $(I_{enter}/I_{leave}) > 1$, perform splitting
 - If $(I_{enter}/I_{leave}) < 1$, perform Russian roulette



- If n > 1, **split** into n particles with weight (wgt/n)
 - All of the n particles emerging from splitting have identical attributes (e.g., x,y,z, u,v,w, E) including wgt' = wgt/n
 - All of the n particles from a splitting are part of the same history, and their tallies must be combined
 - Typically, (n-1) particles are banked, 1 particle is followed until its death, then a particle is removed from the bank & followed, etc.
- Avoid over-splitting
 - Splitting into a large number of particles can increase CPU-time & lead to (apparent) bias in results
 - Typically, choose cell importances to split 2-for-1 or 3-for-1
 - Typically, can limit the splitting to n-for-1 or less
- Total particle weight is exactly conserved in splitting



- With probability **r**, keep the particle & alter its weight to (wgt/r)
 - With probability (1-r), kill the particle (set its weight to 0)

- Russian roulette effectively merges a number of low-weight particles into one with higher weight
- Total particle weight is only conserved statistically (expected value)

Weight Cutoff



- If particle weight drops below
 W_{low}, play Russian roulette with
 weight of W_{ave} for survivors
 - Probability of surviving RR = wgt/W_{ave}
 - Probability of being killed = $1 wgt/W_{ave}$

```
If wgt < W_{low},

if \xi < wgt/W_{ave},

wgt' = W_{ave}

else

wgt' = 0
```



Los Alamos

• Expected value of surviving weight is conserved, (wgt/W_{ave})•W_{ave}



Los Alamos

- Let R_i = (importance of **source cell**) / (importance of **cell j**)

- Then,
$$W_{ave}(j) = W_{ave} \cdot R_j$$

 $W_{low}(j) = W_{low} \cdot R_j$

- Weight cutoffs reduce computing time, not variance
- Weight cutoffs can be applied anytime the particle weight changes after collisions, after boundary crossings, ...

Weight Windows





- Prevent particle weights from getting too large or too small
 - Weight too large —> splitting
 - Weight too small -> Russian Roulette


- Large fluctuations in particle weights contributing to a tally lead to larger variance
- Weight windows eliminate large or small weights (outside the window) by creating or destroying particles
- Weight windows can be applied any time after collisions, after surface crossings, ...
 - If wgt > W_{hi} splitting Elseif wgt < W_{low} roulette

MCNP weight window scheme

Input: W_{low} for each cell (can be energy or time dependent), $[W_{ave}/W_{low}]$, $[W_{hi}/W_{low}]$, mxspln

```
If wgt > W_{hi}

n = min(mxspln, 1 + wgt/W_{hi}) < -max splitting is mxspln-to-1

wgt = wgt/n

bank n-1 copies of particle <-- n-to-1 splitting
```

Los Alamos

```
Elseif wgt < W_{low}

P = max( 1/mxspln, wgt/W<sub>ave</sub>) <-- limits survivor to mxspln*wgt

if \xi < P

wgt = wgt/P <-- particle survives

else

wgt = 0 <-- particle killed
```

Source Biasing



- Bias the PDFs used to select the angle, energy, or position or source particles
 - Produce more source particles (with lower weights) in desired parts of phase space

True source:
$$f(R,E,\Omega)$$

- Sample (R',E',Ω') from $g(R,E,\Omega)$
 - & assign weight $f(R',E',\Omega')/g(R',E',\Omega')$ to source particle

Choose $g(R,E,\Omega)$ to favor directions more important to tallies



- Particles entering specified cells are split into collided & uncollided parts
 - For distance-to-boundary d

Prob(no collision) = $exp(-\Sigma_T d)$

Prob(collision) = $1 - \exp(-\Sigma_T d)$



• Los Alamos

• Sampling the flight distance **s** for a forced collision with max flight distance **d**

Sampling from a truncated exponential PDF:

$$\mathbf{f}(\mathbf{s}) = \boldsymbol{\Sigma}_{\mathsf{T}} \cdot \frac{\mathbf{e}^{-\boldsymbol{\Sigma}_{\mathsf{T}}\mathbf{s}}}{\mathbf{1} - \mathbf{e}^{-\boldsymbol{\Sigma}_{\mathsf{T}}\mathbf{d}}}, \qquad \mathbf{0} \le \mathbf{s} \le \mathbf{d}$$

$$\mathbf{F}(\mathbf{s}) = \frac{\mathbf{1} - \mathbf{e}^{-\Sigma_{\mathsf{T}} \mathbf{s}}}{\mathbf{1} - \mathbf{e}^{-\Sigma_{\mathsf{T}} \mathbf{d}}}$$

Solve for s:
$$\xi = F(s)$$

$$s = \frac{-\ln[1 - (1 - e^{-\Sigma_T d})\xi]}{\Sigma_T}$$



• Encourage particles to head in a certain preferred direction, Ω_0



- Replace Σ_T by $\Sigma^* = \Sigma_T [1 p \Omega \cdot \Omega_0]$
 - p = a parameter, 0
 - Ω_0 = unit vector from particle position to detector
 - Ω = actual particle direction
- Sample flight distance **s'** from $g(s) = \Sigma^* exp(-\Sigma^* s)$
- Adjust weight by factor:

$$f(s')/g(s') = \exp(-p \Omega \cdot \Omega_0 \Sigma_T s')/[1 - p \Omega \cdot \Omega_0]$$

- Paths toward detector are stretched $(\Sigma^* < \Sigma_T)$
- Paths away from detector are shortened ($\Sigma^* > \Sigma_T$)

Variance Reduction Goals & Cautions



- Maximize FOM either reduce RE or T
- Keep the number of particles per cell roughly constant from source to detector
- Reduce the number of particles in unimportant regions
- Achieve adequate sampling of all portions of phase space
- Avoid over-biasing (e.g., over-splitting)
- Ensure that tallies pass statistical checks

References



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- L.L. Carter & E.D. Cashwell, "Particle Transport Simulation with the Monte Carlo Method", TID-26607, National Technical Information Service (1975)



Fundamentals of Monte Carlo Particle Transport

Lecture 10



Parallel Monte Carlo

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Parallel Monte Carlo



Parallel Computing

- Parallel Computers
- Message Passing
- Threads
- Amdahl's Law

Parallel Monte Carlo

- Parallel Algorithms
- Histories, Random Numbers, Tallies
- Load Balancing, Fault Tolerance, ...

Parallel Monte Carlo Performance

- Performance Measures & Limits
- Parallel Scaling

MCNP5 Parallel Processing

- MCNP5 parallelism
- MPI or PVM + Threads
- Run Commands & Input Options
- Performance on ASCI Tera-scale systems
- Parallel Processing for Large-scale Calculations



Parallel Computing



Fast desktop computers

1980s super:	200 MHz	16 MB	10 GB	\$ 20 M
Today, PC:	2000 MHz	1000 MB	100 GB	\$2K

• Linux clusters + MPI

- Cheap parallel computing
- Everyone can do parallel computing, not just national labs

Mature Monte Carlo codes

- MCNP, VIM, KENO, MCBEND, MONK, COG, TART, RACER, RCP, ...
- New generation of engineers/scientists
 - Less patience for esoteric theory & tedious computing procedures
 - Computers are tools, not to be worshipped
 - What's a slide rule ???

➔ More calculations with Monte Carlo codes

Trends in Computing Technology



Commodity chips •

- Microprocessor speed \rightarrow
- Memory size \rightarrow
- Memory latency \rightarrow

High-end scientific computing •

- Key driver (or limit) \rightarrow
- Architecture \rightarrow
- **Operating systems** •
 - **Desktop & server** \rightarrow
 - Supercomputers Unix, Linux \rightarrow

High-performance scientific computing \rightarrow parallel computing

~2x gain / 18 months

- ~2x gain / 18 months
- \sim no change (getting worse)

economics: mass production of desktop PCs & commercial servers

clusters: with small/moderate number of commodity microprocessors on each node

Windows, Linux

Parallel Computers





Characterize computers by: •

- CPU: scalar, vector, superscalar, RISC,
- Memory: shared, distributed, cache, banks, bandwidth,
- Interconnects: bus, switch, ring, grid,
- **Basic types:** •

CPU

Mem





Los Alamos



Approaches to Parallel Processing



High-level	 Independent programs + message-passing Distribute work among processors Loosely-coupled Programmer must modify high-level algorithms
Mid-level	 Threads (task-level) Independent tasks (subprograms) + shared memory For shared memory access, use locks on critical regions Compiler directives by programmers
Low-level	 Threads (loop-level) Split DO-loop into pieces, compute, synchronize Compiler directives by programmers
Low-level	Pipelining or vectorizationPipelined execution of DO-loops

 Automatic vectorization by compilers &/or hardware, or compiler directives by programmers

Message-passing





- Independent programs
- Separate memory address space for each program (private memory)
- All control information & data must be passed between programs by explicit messages (SENDs & RECEIVEs)
- Can run on distributed or shared memory systems
- Efficient only when T_{computation} >> T_{messages}
- Standard message-passing:
 - MPI
 - PVM

Threading (task-level)





- Single program, independent sections or subprograms
- Each thread executes a portion of the program
- Common address space, must distinguish private & shared data
- Critical sections must be "locked"
- Can run only on shared memory systems, not distributed memory
- Thread control by means of compiler directives
- Standard threading:
 - OpenMP

Threading (loop-level)





- Single DO-loop within program
- Each loop iteration must be independent
- Each thread executes different portion of DO-loop
- Invoked via compiler directives
- Standard threading:
 - OpenMP

Domain Decomposition





- Coarse-grained parallelism, high-level
- For mesh-based programs:
 - 1. Partition physical problem into blocks (domains)
 - 2. Solve blocks separately (in parallel)
 - 3. Exchange boundary values as needed
 - 4. Iterate on global solution
- Revised iteration scheme may affect convergence rates
- Domain decomposition is often used when the entire problem will not fit in the memory of a single SMP node



If a computation has fast (parallel) and slow (scalar) components, the overall calculation time will be dominated by the slower component

Overall System	=	Single CPU		1	
Performance		Performance		1-F + F/N	

where	F = fraction of work performed in parallel			
\mathbf{N} = number of parallel processors				
Spee	dup = 1 / (1 - F + F / N)			

For N=10			Fo	For N=infinity					
	<u>F</u>	<u>S</u>	E	<u>S</u>		E	<u>S</u>	E	<u>S</u>
	20%	1.2	90%	5.3		20%	1.3	90%	10
	40%	1.6	95%	6.9		40%	1.7	95%	20
	60%	2.2	99%	9.2		60%	2.5	99%	100
	80%	3.6	99.5%	9.6		80%	5	99.5%	200



My favorite example

Which system is faster?

System A: (16 processors) • (1 GFLOP each) = 16 GFLOP total

System B: (10,000 procs)·(100 MFLOP each) = **1,000 GFLOP total**

Apply Amdahl's law, solve for F:

1 / (1-F + F/16) = .1 / (1-F + F/10000)

→ System A is faster, unless >99.3% of work is parallel

- In general, a smaller number of fatter nodes is better
- For effective parallel speedups, must parallelize everything



Parallel Monte Carlo



- Possible parallel schemes:
 - **Jobs** run many sequential MC calculations, combine results
 - **Functional** sources, tallies, geometry, collisions,
 - **Phase space** space, angle, energy
 - **Histories** Divide total number of histories among processors
- All successful parallel Monte Carlo algorithms to date have been history-based.
 - Parallel jobs always works, variation on parallel histories
 - Some limited success with spatial domain decomposition



- **Master task**: control + combine tallies from each slave
- Slave tasks: Run histories, tallies in private memory
 - Initialize:

Master sends problem description to each slave (geometry, tally specs, material definitions, ...)

- **Compute**, on each of N slaves:

Each slave task runs 1/N of total histories.

Tallies in private memory.

Send tally results back to Master.

- Combine tallies:

Master receives tallies from each slave &

combines them into overall results.

• Concerns:

- Random number usage
- Load-balancing
- Fault tolerance (rendezvous for checkpoint)
- Scaling



Control + Bookkeeping

Computation



save_files()



etc.

- Linear Congruential RN Generator $S_{k+1} = g S_k + C \mod 2^M$
- RN Sequence & Particle Histories

1 2 3 MCNP stride for new history: 152,917

• To skip ahead k steps in the RN sequence:

 $S_k = g S_{k-1} + C \mod 2^M = g^k S_0 + C (g^{k-1})/(g-1) \mod 2^M$

Initial seed for n-th history

 $S_0^{(n)} = g^{n^*152917} S_0 + C (g^{n^*152917}-1)/(g-1) \mod 2^M$

This is easy to compute quickly using exact integer arithmetic

- Each history has a unique number
 - Initial problem seed \rightarrow initial seed for nth particle on mth processor
 - If slave knows initial problem seed & unique history number, can initialize RN generator for that history



- On parallel systems with complex system software & many CPUs, interconnects, disks, memory, MTBF for system is a major concern.
- Simplest approach to fault tolerance:
 - Dump checkpoint files every M histories (or XX minutes)
 - If system crashes, restart problem from last checkpoint
- Algorithm considerations
 - Rendezvous every M histories.
 - Slaves send current state to master, master saves checkpoint files
 - Parallel efficiency affected by M.

Fault Tolerance





- For efficiency, want (compute time) >> (rendezvous time)
 - Compute time: Proportional to #histories/task
 - Rendezvous time: Depends on amount of tally data & latency+bandwidth for message-passing



– Initialize:

Master sends problem description to each slave (geometry, tally specs, material definitions, ...)

- For rendezvous = 1, L
 - **Compute**, on each of N slaves:

Each slave task runs 1/N of (total histories)/L. Tallies in private memory. Send tally results back to Master.

Combine tallies:

Master receives tallies from each slave & combines them into overall results.

• Checkpoint:

Master saves current tallies & restart info in file(s)



• Time per history may vary significantly

- For problems using **variance reduction**:
 - Particles headed in "wrong" direction may be killed quickly, leading to a short history.
 - Particles headed in "right" direction may be split repeatedly. Since the split particles created are part of the same history, may give a very long history.
- For problems run on a workstation **cluster**:
 - Workstation nodes in the cluster may have different CPU speeds
 - Workstations in the cluster may be simultaneously used for interactive work, with highly variable CPU usage on that node.
 - Node performance effectively varies continuously over time.

Naïve solution

- Monitor performance per node (e.g., histories/minute)
- Periodically adjust number of histories assigned to each node, according to node performance

histories assigned to node n \sim measured speed of node n

Better solution:

self-scheduling



- For a problem with N slave processors, divide histories into **more than** N chunks.
 - Let L = number of chunks, L > N
 - Typically, $L \sim 20 \text{ N}$ or $L \sim 30 \text{ N}$
 - Histories/chunk = (total histories) / L
 - Slave: If idle, ask master for work. Repeat until no more work.
 - Master: Send chunk of work to idle slave. Repeat until no more work.
 - On average, imbalance in workload should be < 1/L
- Additional gains:
 - Naïve master/slave algorithm is **synchronous**
 - Self-scheduling master/slave algorithm is asynchronous. More overlap of communication & computation → reduced wait times & better performance

Load Balancing – Self-Scheduling





- Much more communication with Master, but only minimal amount of control info needed (1st & last history in chunk)
- Need to handle stopping condition carefully avoid "dangling" messages

Load Balancing – Self-Scheduling





save_files()



- For clustered SMPs,
 - Use message-passing to distribute work among slaves ("boxes")
 - Use threading to distribute histories among individual processors on box



 Only the master thread (thread 0) on each slave uses MPI send/recv's
Master / Slave Algorithm, threaded & self-scheduling







Parallel Monte Carlo Performance

Parallel MC Computational Characteristics



- For master/slave algorithms (with self-scheduling, fault tolerance, & threads):
 - No communication among slave tasks
 - Occasional communication between master & slaves (rendezvous)
 - Slave tasks are compute-intensive
 - Few DO-loops
 - 40% of ops are test+branch (IF... GOTO...)
 - Irregular memory access, no repetitive patterns
 - For fixed-source problems:
 - Only 1 rendezvous is strictly necessary, at end of calculation
 - More rendezvous used in practice, for fault tolerance
 - For eigenvalue problems (K-effective):
 - Must have a rendezvous every cycle
 - Master controls iteration & source sampling
- Common-sense approach to performance:

Fewer rendezvous → better parallel performance

(cycle = batch = generation)

Parallel MC Performance Measures



- Speedup $S_N = T_1 / T_N$ N = # processors
- Efficiency $E_N = S_N / N$
- Fixed overall work

(fixed problem size)

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- Efficiency decreases with N
- Speedup (eventually) drops as N increases
- Why?

As N increases, same communication/processor, but less work/processor (fewer histories/processor) → (computation/communication) decreases

Fixed work per processor

(scaled problem size)

- Efficiency approx. constant with N
- Speedup approx. linear with N
- Why?

As N increases, same communication/processor, same work/processor (# histories \sim N) \rightarrow (computation/communication) stays approx. same

Called scaled speedup



Another way to determine efficiency

Parallel Efficiency = $T_C / (T_C + T_M)$

 T_{C} = computing time

 T_M = time for messages, not overlapped with computing

Slaves can send messages in parallel



Master receives & processes messages serially

$$\rightarrow \rightarrow \rightarrow \blacksquare$$

If enough messages are sent to master, extra wait time will limit performance



- N = # processors
- $T_1 = CPU$ time for M histories using 1 processor
 - (Depends on physics, geometry, compiler, CPU speed, memory, etc.)
- L = amount of data sent from 1 slave each rendezvous
- $T_{\rm S} = 0$ negligible, time to distribute control info
- $T_R = s + L/r$ **s** = latency for message, **r** = streaming rate
- $T_C^{fix} = T_1 / N$ fixed problem size, **M** histories/rendezvous $T_C^{scale} = T_1$ scaled problem size, **NM** histories/rendezvous



• Scaling models, for master/slave with serial rendezvous

- "fixed" = constant number of histories/rendezvous, M (constant work)
- "scaled" = M histories/slave per rendezvous, NM total (constant time)



- N = number of slaves
- $c = (s + L/r) / T_1$
- $T_1 \sim M$,more histories/rendezvous \rightarrow larger T_1 , smaller cS+L/r,fixed, determined by number of tallies,

As $M \rightarrow infinity$, $c \rightarrow 0$, $S \rightarrow N$ (limit for 1 rendezvous)

Parallel MC Performance Scaling • Los Alamos Fixed size, serial messages S $S = N/(1 + cN^2)$ <u>2 √c</u> Ν ٧C Scaled size, serial messages 1/c -S = N/(1 + cN)1 / 2c Ν 1/c N = number of slaves

 $c = (s + L/r) / (M_1 t_h)$



VIM Monte Carlo — Example of Estimating Performance

 TREAT reactor, 	con	tinuous-	energy ne	utron transpo	ort
 Per slave: 	M ₁	= 400 k	nistories,		
	L	= .341	MB of tal	ly data	
 Communications: 					
Workstation network, IBM-SP1, IBM-SP1,		P4 + ethernet: s P4 + ethernet: s P4 + EUI-H: s		s ≈ .001 sec s ≈ .001 sec s ≈ .50 µsec	e, r ≈ .8 MB/sec e, r ≈ .8 MB/sec e, r ≈ 8.5 MB/sec
		t _h	T ₁ = 400)t _h T _r	С
Sun Sparc2	.25	sec	100 sec	.43 sec	.004
rs6000/350	.10	sec	40 sec	.43 sec	.011
SP1 - ethernet	.075	5 sec	30 sec	.43 sec	.014
SP1 - EUIH	.075	5 sec	30 sec	.04 sec	.001

Note: T_1 — very repeatable & predictable.

 T_r — difficult to measure on busy machine



VIM Monte Carlo — Measured Performance on SP1 — TREAT reactor



Measured message passing on SP1 is 2-3 times slower than specs (busy machine; experimental software; flaky hardware)

Parallel MC Performance Scaling



$$S = N/(1+cN)$$

 $S_{max} = 1/c$
 $N_{1/2} = 1/c$

N = number of slaves c = (s+L/r) / (M₁t_h)

Examples:

• VIM, TREAT problem

Sun Sparc2 workstation cluster	c=.0043	S _{max} = 233
rs6000/350 workstation cluster	c=.011	$S_{max} = 93$
SP1, using ethernet	c=.014	$S_{max} = 70$
SP1, using EUIH comm.	c=.00134	S _{max} = 748

RACER, "typical" large problem

100 K histories/min, 20 K histories/slave 32 MB tally data, r ~ 1800 MB/sec Cray-C90, using SSD for messages c~.001 S_{max} ~ 1000 (16 processors, max)





Master/slave algorithms work well

- Load-balancing: Self-scheduling
- Fault-tolerance: Periodic rendezvous
- Random numbers: Easy, with LCG & fast skip-ahead algorithm
- Tallies: Use OpenMP "critical sections"
- Scaling: Simple model, more histories/slave + fewer rendezvous
 - Hierarchical: Master/slave MPI, OpenMP threaded slaves
- Portability: MPI/OpenMP, clusters of anything

Remaining difficulties

- Memory size: Entire problem must fit on each slave
 - Domain-decomposition has had limited success
 - Should be OK for reactor problems
 - May not scale well for shielding or time-dependent problems
 - For general 3D geometry, effective domain-decomposition is unsolved problem
 - Random access to memory distributed across nodes gives huge slowdown
 - May need functional parallelism with "data servers"



MCNP5 Parallel Calculations

DOE Advanced Simulation & Computing – ASC





Red – 3 TeraOps



White – 12 TeraOps



Blue Pacific – 3 TeraOps



Lightning Red Storm Blue Gene/L



Blue Mountain – 3 TeraOps (R.I.P.)



Q – 20 TeraOps



- Use **message-passing** to distribute work among slaves ("boxes")
- Use threading to distribute histories among individual cpus on box



• We routinely test MCNP5 on:

- ASCI Bluemountain –
- ASCI White –
- ASCIQ-
- Linux clusters
- Windows PC cluster
- 1,000 processor jobs are "routine"

SGI,	48 boxes x	128 cpus/box
IBM,	512 boxes x	16 cpus/box
HP,	2 x 512 boxes x	4 cpus/box



• Threading

- Individual histories are handled by separate threads
- No thread synchronization is needed during a history
- Implemented by OpenMP compiler directives
- Tallies, RN data, & some temporary variables for history are in thread-private memory

Example:

common **/RN_THREAD**/ RN_SEED, RN_COUNT, RN_NPS !\$OMP THREADPRIVATE (**/RN_THREAD**/) save /RN_THREAD/

- OpenMP critical sections are used for some tallies or variable updates

Example:

!\$OMP CRITICAL (RN_STATS)
RN_COUNT_TOTAL = RN_COUNT_TOTAL + RN_COUNT
\$!OMP END CRITICAL (RN_STATS)

 Message-passing & file I/O are executed only from thread-0 (master thread) for each MPI task



• Message-passing

- In MCNP5, all message-passing is handled by calls to the **dotcomm** package, a communications layer which contains an interface to either MPI or PVM
- Recommend using MPI PVM is obsolete & won't be supported in future



- Either MPI or PVM message-passing is selected in **dotcomm** <u>at compile-time</u>
- Using the dotcomm package & either MPI or PVM, MCNP5 can run in parallel without source code changes on
 - Parallel supercomputers (e.g., ASCI tera-scale computers)
 - COWs (clusters of workstations)
 - Linux clusters
 - PC clusters



N = total number of MPI tasks, master + (N-1) slaves

M = number of OpenMP threads/slave

• Running on parallel systems with MPI only

mpirun -np N mcnp5.mpi i=inp01

• Running with threads only

mcnp5 tasks M i=inp01

• Running on parallel systems with MPI & threads

ASCI Bluemountain (SGI)

mpirun -np N mcnp5.mpi tasks M i=inp01

ASCI Q (HP/Compaq)

prun – n N - c M mcnp5.mpi tasks M i=...

If submitting jobs through a batch system (e.g., LSF), N & M must be consistent with LSF requested resources



• How many threads ?

- Max number of threads = # CPUs per node
 - ASCI Bluemountain: 128 cpus / node
 ASCI Q: 4 cpus /node
 - Laptop PC cluster: 1 cpu / node
- Experience on many systems has shown that a moderate number of threads per slave is efficient; using too many degrades performance

•	ASCI Bluemountain:	4–12 threads/slave usually effective
		>16 threads/slave usually has bad performance
•	ASCI Q:	4 threads/slave is effective

- Rules-of-thumb vary for each system
 - Thread efficiency is strongly affected by operating system design
 - Scheduling algorithm for threads used by operating system is generally designed to be efficient for small number of threads (<16)
 - For large number of threads, context-switching & cache management may take excessive time, giving poor performance
 - Other jobs on system (& their priority) affect thread performance
 - No definite rules need to experiment with different numbers of threads

MCNP5 Parallel Calculations



- Parallel performance is sensitive to number of rendezvous
 - Can't control number of rendezvous directly
 - The following things cause a rendezvous:
 - Printing tallies
 - Dumping to the RUNTPE file
 - Tally Fluctuation Chart (TFC) entries
 - Each cycle of eigenvalue problem

• Use PRDMP card to minimize print/dump/TFC

PRDMP	ndp	ndm	mct	ndmp	dmmp	
ndp= increment for printing tallies ← use large numberndm= increment for dump to RUNTPE ← use large numbermct= flag to suppress time/date info in MCTALndmp= max number of dumps in BUNTPE						
dmmp	= increment	🗲 use large number				
For fixe For eig	ed-source p jenvalue pro	roblems, blems,	increme increme	ents are in <mark>p</mark> ents are in <mark>c</mark>	articles ycles	



- Keff calculations: Use KCODE card for hist/cycle
 - Want to reduce the number of cycles
 - More histories in each cycle
 - Should run hundreds of cycles or more for good results

KCODE	nsrck	rkk	ikz	kct		
nsrck rkk ikz kct	= histories / cyc = initial guess fo = number of init = total number of	cle or Keff ial cyclo of cycles	es to discard s to run	← use	a large nur	nber

Suggested: nsrck \sim (thousands) x (number of processors)

MCNP5 Parallel Scaled Speedup





MCNP5 Parallel Calculations





10 -50

Scaled Parallel Speedup – Eigenvalue Problem







Parallel Processing For Large Monte Carlo Calculations



If a Monte Carlo problem is too large to fit into memory of a single processor



- Need periodic synchronization to interchange particles among nodes
- Use message-passing (MPI) to interchange particles
- Domain decomposition is often used when the entire problem will not fit in the memory of a single SMP node

Parallel Monte Carlo



- Inherent parallelism is on particles
 - Scales well for all problems

Domain decomposition

- Spatial domains on different processors
- Scales OK for Keff or α calculations, where particle distribution among domains is roughly uniform
- Does **not** scale for time-dependent problems due to severe load imbalances among domains

Domain decomposition – scaling with N processors

- Best: performance ~ N (uniform distribution of particles)
- Worst: performance ~ 1 (log

(localized distribution of particles)



- Data is distributed by domain decomposition, but parallelism is on particles
- Solution ?

Parallel on particles + distributed data

- Particle parallelism + Data Decomposition
 - Existing parallel algorithm for particles
 - Distribute data among processor nodes
 - Fetch the data to the particles as needed (dynamic)
 - Essentially same approach as used many years ago for CDC (LCM) or CRAY (SSD) machines
 - Scales well for all problems (but slower)



• Particle parallelism + data decomposition – <u>logical</u> view:



- Mapping of logical processes onto compute nodes is flexible:
 - Could map particle & data processes to different compute nodes
 - Could map particle & data processes to **same** compute nodes
- Can replicate data nodes if contention arises



• Particle parallelism + data decomposition



Local copies of data for particle neighborhood



History modifications for data decomposition

source

- while wgt > cutoff
 - compute distances & keep minimum:
 - dist-to-boundary
 - dist-to-time-cutoff
 - dist-to-collision
 - dist-to-data-domain-boundary
- . move particle
 - pathlength tallies
 - if distance == dist-to-data-domain-boundary fetch new data
- collision physics
- . roulette & split

. . .

Parallel Monte Carlo



Data windows & algorithm tuning

- Defining the "particle neighborhood" is an art
- Anticipating the flight path can guide the pre-fetching of blocks of data
- Tuning parameters:
 - How much data to fetch ?
 - Data extent vs. particle direction ?

Entire physical problem

Los Alamos



• Examples



Parallel Monte Carlo



Point detector tallies are non-local

- Every collision contributes an expected score
- At every collision, "pseudo-particles" are tracked along the path from collision to detector
- Scores from all "pseudo-particles" (including from all split particles) must be tallied together into a single score for the history



Entire physical problem



For Monte Carlo problems which can fit in memory:

- Concurrent scalar jobs ideal for Linux clusters
- Master/slave parallel algorithm (replication) works well
 - Load-balancing: Self-scheduling
 - Fault-tolerance: Periodic rendezvous
 - Random numbers: Easy, with LCG & fast skip-ahead algorithm
 - Tallies: Use OpenMP "critical sections"
 - Scaling: Simple model, more histories/slave + fewer rendezvous
 - Hierarchical: Master/slave MPI, OpenMP threaded slaves
 - Portability: MPI/OpenMP, clusters of anything

For Monte Carlo problems too large to fit in memory:

- Spatial domain decomposition (with some replication) can work for some problems
- Particle parallelism + data decomposition is a promising approach which should scale for all problems

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